## UNDERSTANDING FRENCH CONSUMERS' PREFERENCES FOR FISH

- Eliciting Willingness to Pay by the Use of a Non-Hypothetical Choice Experiment

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Any remaining inaccuracies are ours and ours alone.

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## Practical Information

This booklet consists of four parts; an article and three appendices.

The article is an independent document with a table of contents at the beginning and a list of references at the end.

Appendix I consists of an in-depth procedure of how the variables for the models in the article were chosen. The theory behind the random utility models, prior research and the theory behind the mixed logit models are also presented there. This appendix furthermore contains answers from the survey and regression outputs.

The sections, tables and figures in this appendix begin with an "A". In the article we refer to sections from this appendix as e.g. "(Appendix I, A.6.1.2.)". We refer to tables as e.g. "(Appendix I, Table A.11.2.)" .

This appendix is an independent document with a table of contents at the beginning and a list of references at the end.

Appendix II consists of the survey questions that were relevant for the article. We end a sentence with "(Appendix II)" when we refer to this document.

Appendix III consists of a table with statistics on Norwegian seafood exports. We end a sentence with "(Appendix III)" when we refer to something which is taken from this table.

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#### Abstract

A Real Choice Experiment was used to examine French consumers' valuations of fresh salmon, farmed cod, wild cod, monk and pangasius. The study focuses on the willingness to pay for farmed cod. The participants were recruited by the French National Institute for Agricultural Research. Real economic incentives were introduced in the experiment by letting one randomly drawn choice scenario be binding.

By the use of a mixed logit model we study how consumers value the five fish types and their associated attributes, and how different consumer groups differ in their valuations. We find heterogeneous preferences for all the fish types. The preferences for pangasius are found to be most heterogeneous, while the preferences for salmon are found to be the most homogeneous. On average the participants are willing to pay more for wild cod than for farmed cod.

We further find that the participants value a piece of tail cut lower than a piece loin. Our results suggest that individuals with high income are willing to pay a price premium for both farmed and wild cod. The willingness to pay for salmon decreases with age, and the willingness to pay for wild cod and monk increases with age. Individuals living in single households are willing to pay a price premium for salmon, wild cod and pangasius.

Keywords: fish, aquaculture, consumer groups, real choice, willingness to pay, mixed logit


## Sammendrag

Denne studien undersøker franske konsumenters preferanser for ferske fiskefileter av laks, villtorsk, oppdrettstorsk, breiflabb og pangasius ved hjelp av et valgeksperiment. Studien fokuserer på konsumenters betalingsvillighet for oppdrettstorsk. Deltakerne i eksperimentet ble rekruttert av Frankrikes Nasjonale Institutt for Landbruksforskning.

Økonomiske insentiver er introdusert i eksperimentet ved at ett valg er tilfeldig trukket ut til å være bindende. Ved å ta i bruk en mixed logit-modell, undersøker vi hvordan konsumenter verdsetter de fem fisketypene og deres tilhørende attributter, samt hvordan verdsettingen varierer mellom ulike konsumentgrupper.

Resultatene av analysen viser at det er heterogene preferanser for alle fiskeslagene. Videre viser analysen at preferansene for pangasius er de mest heterogene, mens preferansene for laks er de mest homogene. Deltakerne i eksperimentet er i snitt villige til å betale mer for villtorsk enn for oppdrettstorsk. Videre finner vi at deltakerne foretrekker fileter av loin framfor fileter av halestykker. Våre resultater tyder på at folk med høy inntekt er villige til å betale et prispåslag for både villfanget og oppdrettet torsk. Betalingsvilligheten for laks er avtakende med alder, og betalingsvilligheten for villtorsk og breiflabb er økende med alder. Personer i enslige husholdninger har høyere betalingsvillighet for laks, villtorsk og pangasius enn personer i samboende husholdninger.

Nøkkelord: fisk, akvakultur, konsumentgrupper, reelle valgeksperimenter, betalingsvillighet, mixed logit

## 1. Introduction

Norway has a long tradition as a fishing nation. A long coastline and many fjords provide good conditions for harvesting from the sea. France, on the other hand, is worldwide known for its quality cuisine. France is the second largest importer of Norwegian seafood, only surpassed by Russia (Appendix III). Salmon and cod are two of the most frequently eaten fish types in France (Willemsen 2003, p. 9).

Salmon farming has been very successful in Norway. The farming of cod, however, is still in its initial face and has not yet been able to achieve a similar success as the salmon farming industry. Since France is one of the main importers of Norwegian seafood, an up to date and well informed understanding of the French market is important. The objectives of this article are to study the French consumers' preferences and willingness to pay (WTP) for fish, and to examine how different consumer groups differ in their valuations.

### 1.1. Background

The success of the Norwegian salmon farming industry is visible through the almost hundredfold increase in exported volume over the last thirty years. In 1981 Norway exported 7.452 metric tons of salmon (and rainbow trout), and in 2010 the export volume was 714.484 (Statistics Norway 2012). Salmon was once regarded a high society food in France, but it is now accessible to everyone at an acceptable price (Nilssen \& Monfort 2000). The reduction in price can be ascribed to low production costs resulting from improved technology and breeding techniques. In addition to price reductions there are many other advantages with fish farming. Modern breeding techniques can improve the fish's health, shape, texture, color, and nutritional content. Aquaculture can to some extent comply with the growing global problems of overfishing, since it is possible to control the amount being produced. Aquaculture can also provide jobs in rural areas.

Aquaculture does, however, interfere with the environment and wild populations of fish. This can cause negative externalities if farmed fish escape and spread diseases or genetic material to the wild stocks (Food and Agriculture Organization 2012). Aquaculture also interferes with alternative uses of the coast, which can lead to negative externalities in terms of, e.g. lost tourism.

The use of antibiotics in salmon and cod aquaculture has been highly controversial and has led to criticism from many consumers (see e.g., Gruben (2007), Tveterås (2003) and Food and Agriculture Organization (2012)). However, the preferences concerning controversial products are highly heterogeneous (Alfnes \& Rickertsen 2011). This demonstrates the importance of understanding how consumer groups differ in their attitudes and retail behavior.

### 1.1.1. The Cod Farming Industry Today

Cod farming has not yet experienced a similar success as salmon farming. The cod farming industry still faces challenges in terms of understanding basic biological issues, and in finding production methods that ensure a stable and profitable production. "A boom-like investment period during 2000-2008 and rapid biomass build-up was followed by an almost collapse after the financial crisis in 2008" (Food and Agriculture Organization 2012). Despite the reduced access to capital in the wake of the financial crisis, the interest for cod farming has increased over the last years. The Norwegian production of farmed cod has risen from 300 metric tons in 2003 to about 20.000 metric tons in 2010 (Nereng 2011). Norway produces about $80 \%$ of the world's farmed cod, nonetheless, this quantum was only about one tenth of the annual catch of Norwegian wild cod in 2009 (Olstad 2011).

The marketing channels for farmed cod have so far been the same as those for wild cod. Wild cod has its greatest supply in the springtime, and the farmed cod industry has focused on delivering in the autumn. Farmed cod can, to a greater extent than wild cod, deliver stable supplies to the major retail chains throughout the year. However, high production costs constitute a difficult starting point for competition. Other fish species are competitive in price and partly in quality, e.g. pangasius and tilapia. Moreover, the volume of wild cod to be supplied to the European market is expected to increase over the next years (Toften 2009). According to The Food and Agriculture Organization of the United Nations, it is likely that the growth in cod aquaculture production will be much slower than what was expected a few years ago. They also note that the structure of the industry is currently quite unclear (Food and Agriculture Organization 2012).

Asche (2009) argues that the future success of cod farming will depend on the industry's ability to make use of its competitive advantages rather than to compete with wild cod on
price. Furthermore, Nofima, Europe's largest institute for applied research within the fields of fisheries, aquaculture and food, presents the assessments of the strategic competiveness of the Norwegian cod farming industry in a report Oppdrettstorsk - konkurransegrunnlag, marked og strategiske muligheter (Toften 2009). They argue that cod farmers in reality have two choices: To position themselves at the higher price end of the market and create niche products, or to reduce production costs considerably and increase the production volume (Toften 2009, p. 17-18). It follows that an understanding of how French consumers value farmed cod compared to other fish types is important for the cod farming industry. It is also of interest to examine how consumer groups differ in their valuations of fish products.

### 1.2. The Experiment

To get a better understanding of the French consumers' preferences for fish, and accordingly their WTP, several kinds of experiments were conducted in Dijon in France, May 2008. Dijon is a city of 151000 inhabitants. The fish consumption in Dijon is regarded to be representative for that of non-coastal France (Alfnes \& Rickertsen 2008a). The experiments consisted of a tasting session, a Becker-DeGroot-Marschak (BDM) ${ }^{1}$ bidding session, and a real choice (RC) experiment. 178 participants took part in the experiments and five types of fish were presented; salmon, farmed cod, wild cod, monk and pangasius. In both the BDM session and in the RC experiment the participants evaluated pre-packed fillets of fresh fish.

In the following a fillet of fish is defined as a piece of fresh fish. The experiments aimed to reveal what type of fish and fish attributes the French consumers find attractive, and to what extent they are willing to pay for these. The experiments were further meant to examine how different consumer groups differ in their WTP for the different fish types.

An understanding of French consumer preferences is important for several reasons. If farmed cod is regarded as undesirable to a great share of the French consumers, a cod farmer might wish to focus on other markets, like e.g. the Russian market. Furthermore, a fish fillet can have many attributes. One important attribute is price, a highly ranked decision criterion in most economic transactions. Knowledge about the effects of this key decision factor is important, both from a marketing point of view and from a production cost perspective. Another attribute of a fish fillet is the type of cut. A fillet can have different types of cut, e.g.

[^0]loin (front cut), tail or round cut. A significant difference in preferences between the different types of cut is valuable information for a fish farmer.

It is also of importance to know how consumer groups differ in their WTP. For instance, if one is able to detect what types of fish people with high income value the most, one can address these products towards this group. Some types of fish, like e.g. wild cod and monk, have a long tradition in the French cuisine, while salmon became accessible at an acceptable price in the 90 s (Nilssen \& Monfort 2000). Hence, it is expected that older consumers differ in their preferences for the different types of fish compared to younger consumers.

Single people make up a non-negligible fraction of the French consumers. In 2009, 32\% of the French lived in one-person households (Statistics Canada 2009). Their eating habits, and hence fish consumption habits, may differ from those of the cohabiting households. Fish consumption habits may also differ between genders, households with or without children, and consumers with different levels of education.

In addition to taking part in the tasting session, the BDM and the RC experiment, the participants answered a survey regarding fish likings, fish buying and eating habits, attitudes towards fish farming, demographics etc. This article focuses on the results from the RC experiment. The results will be compared with answers from the survey.

An RC experiment is meant to mimic a normal grocery store situation, which is a situation most people are familiar with. The participants made real, i.e., non-hypothetical, choices over 16 choice scenarios. Each choice scenario had three fillets of different type. Prices were posted in advance. For every choice scenario the participants chose which fillet they wanted to buy, or a non-of-these (NOT) alternative. Real economic incentives were introduced by letting one randomly drawn choice scenario be binding. This was done to remove the hypothetical bias that may arise from non-consequential choice experiments. RC experiments are a relatively new method for studying consumer preferences and eliciting WTP. This methodology is, however, increasingly utilized by researchers (Gracia et al. 2011).

### 1.2.1. Prior Research

Alfnes et al. (2006) study consumers' WTP for the color of salmon by the use of an RC experiment. Olesen et al. (2010) use an RC experiment to elicit consumers' WTP for organic and welfare-labeled salmon.

Some studies that have employed RC experiments have focused on the differences in WTP values between $R C$ experiments and experimental auctions (EA). In EAs participants take part in a bidding session. There are many versions of EAs, for example a second-price sealedbid auction, also known as a Vickrey auction. In a Vickrey auction participants submit sealed bids for a product, and the highest bidder buys the product for the price of the second highest bid (Alfnes \& Rickertsen 2011). Gracia et al. (2011) find that valuations elicited from EAs can differ from those of RC experiments. Lusk and Schroeder (2006) find that EA bids were significantly lower than the estimated WTP from RC experiments.

Other literature including RC experiments focuses on the hypothetical bias, i.e., the differences in estimated WTP from hypothetical and non-hypothetical choice experiments. An example of a hypothetical choice experiment is the stated choice experiment. In stated choice experiments participants make hypothetical choices over a set of one or more choice scenarios. Participants are asked to pick the product they would have bought, given that it was a real life situation. Lusk and Schroeder (2004) found that WTP values obtained from a stated choice experiment exceeded the WTP values obtained from an RC experiment.

### 1.3. The Random Utility Model and Mixed Logit

Unlike Lusk and Schroeder (2004) who used beefsteak products that varied only in prices over choice scenarios, the experiment used for this study utilized fresh fillets of fish that varied in both prices and products over choice scenarios. By the use of a mixed logit model, we have estimated a random utility model (RUM) for each fish type used in the experiment. A RUM assumes that an individual's utility from choosing a product is a function of observable and, to the researcher, unobservable attributes of both the product and the individual.

Contrary to a standard logit model, a mixed logit model allows for heterogeneous preferences in the population (Train 2009, p. 134-137). In addition to estimating the utility an individual
obtains from choosing a fillet of fish, we estimated a distribution that describes the preference heterogeneity in the population. Chang et al. (2009) find that mixed logit models can have superior performance over other discrete choice models in predicting actual retail shopping behavior. From the utility functions we estimated WTP.

To the best of our knowledge, no earlier research has used an RC experiment to study the preferences for salmon, farmed cod, wild cod, monk and pangasius by including both product attributes and consumer characteristics into the analysis.

The remainder of this article is organized as follows: First we present the sample data and describe the experimental design, before we introduce the econometric model used to analyze the data. The results and estimated WTP are then presented, followed by a discussion of how product attributes and consumer characteristics affect the WTP for fish. We also discuss possible factors that could bias WTP both upwards and downwards. We conclude with some thoughts about the future market potential for farmed cod.

## 2. Sample Data

The French National Institute for Agricultural Research (INRA) randomly drew 178 participants from their consumer panel to the take part in the experiments. A requirement to participate was that they ate fish at least once a month (Appendix II, p. 2).

Table 2.1 gives a descriptive summary of the participants. The participants' age ranged from 21 to 70 years. There were a few more women than men. About one third had children. A quarter of the participants' households' gross monthly income was more than 3000 EUR. The remaining had either less income or did not want to report it. 71\% were married/cohabiting, the others were single. Approximately one third of the participants had higher education ${ }^{2}$.

[^1]Table 2.1. Consumer Characteristics of the 178 Participants

| Variable Definition |  |
| :--- | ---: |
| Gender | $42 \%$ |
| Male | $58 \%$ |
| Female | 46.86 |
| Age (sample average) |  |
| Children | $66 \%$ |
| No children | $34 \%$ |
| One or more children | $65 \%$ |
| Income | $24 \%$ |
| Up to 3000 EUR per month | $11 \%$ |
| More than 3000 EUR per month |  |
| Do not know/Do not want to answer | $71 \%$ |
| Marital Status | $29 \%$ |
| Married/Cohabiting |  |
| Single | $63 \%$ |
| Education | $37 \%$ |
| No higher education |  |
| Higher education |  |

Higher education $37 \%$

## 3. Experimental Procedure

The experiment went over a period of seven days. There were two sessions each day, one at lunch time and one at dinner time. The experiment had nine steps. Step 1: The participants were explained the procedure. Step 2: They took part in a tasting session. The fish was heated to 70 degrees Celsius by a professional chef, and the participants were served a portion of 50 grams of each fish type. The order of the servings was randomized to avoid relative taste bias. As an example, an individual may perceive pangasius differently if served immediately after salmon than if he or she was served pangasius first. Step 3: The participants took part in the BDM bidding session and placed bids on a computer. Step 4: They marked on a questionnaire which alternative they wanted to buy in each choice scenario in the RC experiment. There were three alternatives in each scenario as well as the NOT option. See Figure 3.1 for an illustration of a choice scenario questionnaire and Figure 3.2 for an illustration of a choice set. Step 5: They drew a card to determine their binding scenario. The draw was done without
replacement, so only one participant could be assigned to one scenario. Hence, there were maximum 16 participants in each session. The binding scenario was imposed to reveal true WTP and to avoid the hypothetical bias that may inflate WTP. Step 6: The participants answered the survey while being served dessert. Step 7: Each participant received the fish fillet he or she had chosen in the binding scenario. Step 8: They went to the cashier and got paid 25 EUR less the price of the fillet from the binding scenario.

|  | Alternative 1 | Alternative 2 | Alternative 3 |
| :--- | :---: | :---: | :---: |
|  | $€$ | $€$ | $€$ |
| Box 1 | Salmon | Farmed Cod | Monk |
| I would choose <br> (Check x one) | $\bigcirc$ | $\bigcirc$ | $\bigcirc$ |
|  |  |  |  |
|  | None of these three alternatives | $\bigcirc$ |  |

Figure 3.1. Example of Choice Scenario Questionnaire


Figure 3.2. An Example of a Choice Set

## 4. Products and Experimental Design

The experiment was intended to mimic a normal grocery store situation. The salmon and farmed cod was transported from Norway. The wild cod and the monk were caught in the Northern Atlantic and the pangasius was imported from Vietnam. Each fillet was cut into a 300 gram piece. Salmon, farmed cod and wild cod were either of tail cut or loin. Pangasius and monk did not have tail cuts, as this is not common for these fish types in French supermarkets.

The prices differed between 1.45 and 11.95 EUR for a 300 gram fillet, which corresponds to 4.83 and 39.83 EUR per kilogram. For a product attribute description, see Table 4.1. Pangasius and monk were included as cheap and expensive alternatives to cod, respectively. The participants were not informed about this to avoid framing effects. Framing effects implies giving clues to the participants about how they are supposed to value or perceive the products under scrutiny. If there was a negative focus on e.g. pangasius during the explanation of the procedure, this could have affected the participants' choices in the RC experiment and potentially bias WTP.

Each fillet was in a box laminated with plastic. The information provided on the packages was similar to what one usually finds in supermarkets: fish type, weight, production method (farmed or wild), price and region of origin. Since the participants had taken part in a tasting session before the experiment, they were familiar with the taste of the different fish. We did not have access to the taste scores while analyzing the data.

Table 4.1. Product Attributes of the Fish Fillets

|  | Average Price <br> EUR/KG | Min Price <br> EUR/KG | Max Price <br> EUR/KG | Cut |
| :--- | :---: | :---: | :---: | :--- |
| Salmon | 14.88 | 6.50 | 26.50 | Loin and Tail |
| Farmed Cod | 20.78 | 9.83 | 36.50 | Loin and Tail |
| Wild Cod | 20.79 | 9.83 | 36.50 | Loin and Tail |
| Monk | 27.88 | 18.17 | 39.83 | Round Cut |
| Pangasius | 10.82 | 4.83 | 16.50 | Fillet |

The table shows the average, the minimum and the maximum prices in $E U R / K G$ of fish fillets over all choice scenarios, as well as the type of cut.

The same type of fillet did not occur more than once in each choice scenario. Note that there were five different types of fish and eight different types of fillets. Both prices and products varied among the scenarios. This differs from the design used by Lusk and Schroeder (2004), who used one set of products (five beefsteaks) that varied only in prices over choice scenarios. By letting each choice scenario contain a unique set of fillets, a coincidental, unattractive fillet would not have the same negative effect on WTP than would have been the case if the choice scenarios only varied in prices and not in products.

The variation in prices and products, as well as the positioning of the products in each scenario, was derived from a fractional factorial design. The fractional factorial design will, asymptotically, remove left - or right hand bias, i.e., the tendency to systematically choose a product that is positioned to the right or to the left of another product. The participants could start at any one of the 16 choice scenarios. This removed equal anchoring effects for all the participants and made the sessions take less time. Anchoring effects refers to affixing the prices one observes in the first choice scenario to the prices in the following choice scenarios. For instance, if the prices in the first scenario were very low, one is likely to compare the prices in the following scenarios with the first one. Hence one can obtain an unrealistic picture of the prices, and this can in turn affect WTP. Anchoring effects are commonly found in recent studies, see e.g. Ariely (2010, p. 25-53). Including only three alternatives in each choice scenario lessened the cognitive burden on the participants.

## 5. The Econometric Model

The theory underlying the utility functions in this study is based on Lancastrian consumer theory. Lancastrian consumer theory proposes that the utility associated with a good can be decomposed into separate utilities for the components of the attributes (Loureiro \& Umberger 2004).

The utility an individual obtains from the different choices can be decomposed into observable and unobservable parts. The observable parts are known to both the individual and to the researcher. The observable parts are the known attributes of the fish as well as the known consumer characteristics. We assume the parameter estimates for the observable variables to be linear in parameters. The unobservable parts are known to the individual but
not to the researcher. The unobservable part is represented by a stochastic error term, and hence utility is random. In other words we estimate random utility functions. The stochastic error term is assumed to be independent and identically distributed (iid) extreme value. The utility functions are estimated by mixed logit. Train (2009, p. 134-147) shows that a mixed logit model can approximate any random utility model.

### 5.1. How the Mixed Logit Model Obviates the Limitations of the Standard Logit Model

The mixed logit model obviates three restrictions from the standard logit model:

1. It allows for random taste variation. A mixed logit model can have both fixed and random parameters. The fixed parameters are to be interpreted as if they were standard logit. The random parameters have a distribution with a mean and a standard deviation. This is to capture preference heterogeneity for a product or a product attribute. It is up to the researcher to choose an appropriate distribution for the random parameters. The parameters for the variables representing each fish type are chosen to be the random in this article ${ }^{3}$. Hence, they take people's heterogeneous preferences into account and provide more information about consumer preferences than would have been the case with a standard logit model. We assume a normal distribution for the random parameters ${ }^{4}$.
2. Unrestricted substitution patterns. In a standard logit model the relative probabilities of choosing one alternative over the other is the same, no matter what the other choice alternatives are (Train 2009, p. 34-75). This is known as the property of independence of irrelevant alternatives (IIA). Assuming that the relative probabilities between two alternatives are independent of other alternatives can, in many situations, be a strong assumption. The mixed logit model relaxes this assumption. In the mixed logit model the relative probabilities depend on all the data (Train 2009, p. 134-147). For example, the ratio of the probabilities of choosing farmed $\operatorname{cod}(f c)$ to salmon ( $s a$ ), $P_{f c} / P_{s a}$, also depends on alternatives other than farmed cod and salmon, for example monk.

[^2]3. Correlations in unobserved factors over time or choice scenarios. The mixed logit model can, in addition to estimating a distribution for each random parameter, also estimate the covariance between the random parameters. The dataset from the RC experiment is a panel dataset over sixteen choice scenarios. The choices an individual makes over the sixteen scenarios are likely to be correlated. Assume an individual has a strong preference for cod. Perhaps this person is likely to choose cod in every choice scenario where cod is present, no matter if it is farmed or wild. This implies a positive correlation between choosing farmed and wild cod. This correlation is captured by the covariance matrix. A positive and significant correlation between the preferences for the two types of fish indicates that these products might be substitutes. Allowing for correlations in unobserved factors over time or choice scenarios is an optional feature of the mixed logit model.

### 5.2. Utility as a Function of Observable and Unobservable Variables

Generally the utility an individual $n$ obtains from choosing alternative $i$ in a choice scenario $s$ can be specified as;

$$
\begin{equation*}
U_{n i s}=\boldsymbol{\beta} \boldsymbol{x}_{n i s}+\boldsymbol{\eta}_{n} \mathbf{z}_{n i s}+\varepsilon_{n i s}, \tag{1}
\end{equation*}
$$

where $\boldsymbol{x}_{n i s}$ and $\boldsymbol{z}_{n i s}$ are vectors of observed variables relating to individual $n$, alternative $i$ and choice scenario $s . \boldsymbol{\beta}$ is a vector of fixed coefficients. $\boldsymbol{\eta}_{n}$ is a vector of random parameters with an estimated mean $\boldsymbol{\mu}_{\eta}$ and standard deviation $\boldsymbol{\sigma} . \varepsilon_{n i s}$ is a stochastic iid extreme value error term and varies over individuals and choices, with an expected value of zero.

The subscripts in the models we will present are defined as:
$n=1,2, \ldots, 178 \quad$ indices the participants in the experiment
$i=1,2, \ldots, 5 \quad$ are the fish types to choose from
$i=1$ : Salmon (sa)
$i=2:$ Farmed $\operatorname{Cod}(f c)$
$i=3:$ Wild Cod (wc)
$i=4:$ Monk (mo)
$i=5$ : Pangasius (pa)
$s=1,2, \ldots, 16 \quad$ are the choice scenarios.

An individual $n$ chooses alternative $i$ if and only if $U_{n i s}>U_{n j s}$ for all $i \neq j$. Assume an individual faces a choice scenario. Assume this is a married female, 47 years of age, with higher education and two children. There are three fillets of fish in front of her, e.g. salmon, farmed cod and monk. Each fillet has a price and a cut. She will only choose farmed cod if the utility she obtains from choosing that specific fillet of farmed cod is higher than the utility she would obtain from choosing any other alternatives. Hence, utility is a function of the attributes of the fish fillets, as well as characteristics of the individual. In addition there might be attributes of the fish and characteristics of the individual that we do not observe that might affect her choice. This is captured by the error term $\varepsilon_{n i s}$.

### 5.2.1. The Distribution of the Random Parameters

To better understand how the distributions of the random parameters are derived, we rewrite equation (1) as:

$$
\begin{equation*}
U_{n i s}=\boldsymbol{\beta} \boldsymbol{x}_{n i s}+\boldsymbol{\mu}_{\eta} \mathbf{z}_{n i s}+\left(\boldsymbol{\eta}_{n}-\boldsymbol{\mu}_{\eta}\right) \mathbf{z}_{n i s}+\varepsilon_{n i s} \tag{2}
\end{equation*}
$$

The $\boldsymbol{\mu}_{\eta}$ is a vector of coefficients representing the expected (average) value of the random parameters. Given fish type $i$, every individual $n$ has his or her estimated preference, $\eta_{n}$. This may differ from the estimated mean preference $\mu_{\eta}$. However, the expected difference from the mean is zero; hence the term inside the parentheses in equation (2) has an expected value of zero. Since there are 178 participants in the experiment, there are 178 such differences from the mean. These differences make up a distribution, assumed to be normally distributed. Hence:

$$
\left(\eta_{n}-\mu_{\eta}\right) \sim N(0, W)
$$

from which it follows that

$$
\eta_{n} \sim N\left(\mu_{\eta}, W\right)
$$

where $W$ denotes the covariance matrix for the random parameters. The square roots of the diagonal elements of the covariance matrix are the standard deviations of the random parameters (see Table 5.2.2.1). Large standard deviations imply a great extent of heterogeneity. Small standard deviations imply relatively homogenous preferences in the population. Since we have assumed a normal distribution for the random parameters, they can be illustrated as in Figure 5.2.1.1.


Figure 5.2.1.1. An Example of a Random Parameter Distribution

Assume Figure 5.2.1.1 displays the estimated utility for farmed cod. The parameter estimate, $\mu_{\eta}$, is the expected value. The area to the left of the vertical line represents the share of the population who obtains a below average utility from choosing farmed cod, while the area to the right of the vertical line represents the share of the population who obtains an above average utility from choosing farmed cod.

### 5.2.2. Correlations between the Random Parameters

The off-diagonal elements of the covariance matrix $W$ provide information about the correlations in preferences between the various fish types. A positive and significant covariance between e.g. salmon and monk implies that an individual who is likely to choose salmon is also likely to choose monk. Table 5.2.2.1 shows an example of a covariance matrix.

Table 5.2.2.1. An Example of a Covariance Matrix of the Random Parameters

|  | Salmon | Farmed Cod | Wild Cod | Monk | Pangasius |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Salmon | $\sigma_{11}^{2}$ |  |  |  |  |
| Farmed Cod | $\sigma_{21}^{2}$ | $\sigma_{22}^{2}$ |  |  |  |
| Wild Cod | $\sigma_{31}^{2}$ | $\sigma_{32}^{2}$ | $\sigma_{33}^{2}$ |  |  |
| Monk | $\sigma_{41}^{2}$ | $\sigma_{42}^{2}$ | $\sigma_{43}^{2}$ | $\sigma_{44}^{2}$ |  |
| Pangasius | $\sigma_{51}^{2}$ | $\sigma_{52}^{2}$ | $\sigma_{53}^{2}$ | $\sigma_{54}^{2}$ | $\sigma_{55}^{2}$ |
| $l=$ sa $=$ Salmon, $2=$ fc $=$ Farmed Cod, $3=w c=$ Wild Cod, $4=m o=$ Monk, $5=p a=$ Pangasius |  |  |  |  |  |

Since the covariance matrix is symmetric about its diagonal only the lower triangular matrix is displayed in Table 5.2.2.1. If all the $\sigma^{2}{ }_{i i}$ are zero, the mixed logit collapses to a standard logit.

### 5.3. Model Specification

Two models are presented in this article. Model 1 includes product attributes only. The two product attributes under scrutiny are the price and the type of cut, defined by the variables Price and Tail. Price captures the price sensitivity and its expected sign is negative. This parameter is set to be fixed, assuming homogenous price sensitivity in the population. Tail is a dummy variable taking the value of 1 if the fish is a tail cut, and 0 otherwise. Salmon, farmed cod and wild cod have tail cuts as well as loins. Pangasius and monk do not have tail cuts.

The second model, Model 2, also incorporates how the consumer characteristics income, age and marital status affect an individual's choice. These consumer characteristics are defined by the variables Income, Age and Single respectively.

Income is a dummy variable taking the value of 1 if the participant's household's gross monthly income is more than or equal to 3000 EUR, and 0 otherwise ${ }^{5} .43$ out of the 178 participants (24\%) belong to the high income group. Age is a continuous variable measured in years. The participants' age ranged from 21 to 70 years. Single is a dummy variable taking the value of 1 if the household consists of a single person with or without children, and 0 otherwise. Of the 178 participants in the experiment, $51(29 \%)$ were single. The five types of fish are represented by the vectors $\boldsymbol{z}_{i}$, where $i=\{1=$ Salmon, $2=$ Farmed Cod, $3=$ Wild Cod, $4=$ Monk, $5=$ Pangasius $\}$. Hence each vector has the form:

[^3]\[

z_{1}=z_{s a}=\left[$$
\begin{array}{l}
1 \\
0 \\
0 \\
0 \\
0
\end{array}
$$\right], z_{2}=z_{f c}=\left[$$
\begin{array}{l}
0 \\
1 \\
0 \\
0 \\
0
\end{array}
$$\right], z_{3}=z_{w c}=\left[$$
\begin{array}{l}
0 \\
0 \\
1 \\
0 \\
0
\end{array}
$$\right], z_{4}=z_{m o}=\left[$$
\begin{array}{l}
0 \\
0 \\
0 \\
1 \\
0
\end{array}
$$\right], z_{5}=z_{p a}=\left[$$
\begin{array}{l}
0 \\
0 \\
0 \\
0 \\
1
\end{array}
$$\right]
\]

To estimate the effect of the consumer characteristics on the preferences for fish, interaction terms with the variables representing the fish types, $\boldsymbol{z}_{i}$, are necessary. For example, to estimate the age effect on farmed cod, one must multiply the variable Age with the vector representing farmed cod, i.e., $\boldsymbol{z}_{2}=\boldsymbol{z}_{f c}=[0,1,0,0,0]$. This is the structure of the mixed logit model. The fact that the consumer characteristics do not vary over choice scenarios makes the interaction terms necessary.

Since only Salmon, Farmed Cod and Wild Cod had two types of cut, the Tail variable could not be defined as an explanatory variable on its own. To capture the effect of the Tail variable, interaction terms with Salmon, Farmed Cod and Wild Cod were necessary. This is done by defining the $\boldsymbol{y}_{i}$-vectors, where $i=\{1=$ Salmon, $2=$ Farmed Cod, $3=$ Wild Cod $\}$. For example, to estimate the tail effect of salmon, one must multiply the variable Tail with the vector representing salmon, i.e., $\boldsymbol{y}_{1}=\boldsymbol{y}_{s a}=[1,0,0]$. For Monk and Pangasius the $\boldsymbol{y}$-vectors are zero by default. For a detailed outline of how the utility models are set up, see Appendix I, section A.4.

Both models presented in this article allow for correlations in unobserved factors over choice scenarios. Hence we estimate a covariance matrix for the random parameters instead of only the standard deviations.

In addition to Model 1 and Model 2 we estimated a third model, Model 3, that included the product attributes and consumer characteristics in Model 2 as well as gender, presence of children in the household, and education. The additional consumer characteristics included in Model 3 all turned out to be insignificant. The output and analysis for this model is left to the appendix (Appendix I, section A.9).

### 5.3.1. Expected Utility

Since we are interested in estimating the individual specific average utility over all choice scenarios, we leave out the subscript $s$ from now on. Having defined the variables and the interaction terms, the expected utility from choosing a product is defined as:

## Model 1:

(3) $E\left(U_{i}\right)=\alpha$ Price $+\boldsymbol{\mu}_{\eta} \mathbf{z}_{i}+\boldsymbol{\delta} \boldsymbol{y}_{i}$ Tail

Note that since no consumer characteristic interactions are included in Model 1, the subscript $n$ is omitted. $\alpha$ is the price coefficient, $\boldsymbol{\mu}_{\eta}$ is a vector of the expected values of the random parameters, and $\boldsymbol{\delta}$ is a vector of coefficients for the tail interaction terms.

## Model 2:

(4) $E\left(U_{n i}\right)=\alpha$ Price $+\boldsymbol{\mu}_{\eta} \mathbf{z}_{i}+\boldsymbol{\delta} \boldsymbol{y}_{i}$ Tail $+\boldsymbol{z}_{i}\left[\boldsymbol{\gamma}_{1}\right.$ Income $_{n}+\boldsymbol{\gamma}_{2}$ Age $_{n}+\boldsymbol{\gamma}_{3}$ Single $\left._{n}\right]$

The $\gamma_{\mathrm{s}}$ are coefficient vectors for the interaction terms between the variables Income, Age and Single and the $\boldsymbol{z}_{i}$-vectors representing the fish types. Note that in equation (1) the $\boldsymbol{\beta}$ vector incorporates the coefficients $\alpha$ and $\boldsymbol{\delta}$, and in equation (2) the $\boldsymbol{\beta}$-vector incorporates the coefficients $\alpha, \boldsymbol{\delta}$ and the $\boldsymbol{\gamma}$ s.

### 5.3.2. Estimating WTP

The utility an individual $n$ obtains from the NOT alternative is normalized to zero. Hence a positive utility indicates a willingness to pay for a product. Theoretically, a negative utility implies that an individual should be compensated from choosing a product. In practice it means that he or she does not want to buy it, i.e., he or she prefers NOT to choosing it.

From the estimated utility functions it is possible to estimate the $W T P_{n i}$ for the various fish types $i$, given consumer characteristics of individual $n$. The maximum amount an individual is willing to pay for a product is the price at which he or she is indifferent about buying the product and not buying it. To estimate this, we can set the utility to equal zero in equations (3) and (4), i.e., $E\left(U_{i}\right)=0$ and $E\left(U_{n i}\right)=0$, and solve with respect to Price.

WTP Model 1:
(5) $W T P_{i}=-\frac{1}{\alpha}\left(\boldsymbol{\mu}_{\eta} \mathbf{z}_{i}+\boldsymbol{\delta} \boldsymbol{y}_{i}\right.$ Tail $)$

WTP Model 2:
(6) $W T P_{n i}=-\frac{1}{\alpha}\left(\boldsymbol{\mu}_{\eta} \boldsymbol{z}_{i}+\boldsymbol{\delta} \boldsymbol{y}_{i}\right.$ Tail $+\boldsymbol{z}_{i}\left[\boldsymbol{\gamma}_{1}\right.$ Income $_{n}+\boldsymbol{\gamma}_{2}$ Age $_{n}+\boldsymbol{\gamma}_{3}$ Single $\left.\left.{ }_{n}\right]\right)$

The estimated WTP individual $n$ obtains from choosing product $i$ is a function of the observable variables scaled down by the negative inverse of the price sensitivity parameter $\alpha$.

## 6. Results

In the following "significant" refers to significance at the $5 \%$ level, unless otherwise stated. The terms "utility from choosing" and "preferences" will be used interchangeably.

The results from Model 1 and Model 2 are presented in Table 6.1. The information of interest from the estimated parameters is the sign (positive or negative), the level of significance ( $p$ value) and the relative magnitude between the parameter estimates.

First we focus on the utility obtained from the different fish types and their associated attributes. We begin by presenting the results that are common for both models, before we focus on the results from Model 1 and Model 2 separately. Thereafter we present the estimated WTP from both models.

Table 6.1. Empirical Estimates for Model 1 and Model 2

|  | Model 1 Product Attribute Model |  |  | Model 2 <br> Product Attribute and Consumer Characteristics Interaction Model |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Variables | Estimate |  | Std.Err | Estimate |  | Std.Err |
| Price | -0.219 | *** | (0.010) | -0.220 | *** | (0.010) |
| Salmon | 2.995 | *** | (0.237) | 3.726 | *** | (0.586) |
| Farmed_Cod | 2.359 | *** | (0.328) | 0.928 |  | (0.830) |
| Wild_Cod | 2.922 | *** | (0.286) | 1.121 |  | (0.687) |
| Monk | 2.793 | *** | (0.357) | 1.050 |  | (1.027) |
| Pangasius | -1.518 | *** | (0.520) | -1.525 |  | (1.269) |
| Tail * Salmon | -0.430 | *** | (0.117) | -0.431 | *** | (0.118) |
| Tail * Farmed_Cod | -1.434 | *** | (0.193) | -1.449 | *** | (0.192) |
| Tail * Wild_Cod | -0.667 | *** | (0.152) | -0.672 | *** | (0.152) |
| Income * Salmon |  |  |  | 0.529 |  | (0.354) |
| Income * Farmed_Cod |  |  |  | 0.862 | * | (0.475) |
| Income * Wild_Cod |  |  |  | 1.237 | *** | (0.393) |
| Income * Monk |  |  |  | 0.302 |  | (0.585) |
| Income * Pangasius |  |  |  | 0.924 |  | (0.708) |
| Age * Salmon |  |  |  | -0.022 | ** | (0.011) |
| Age * Farmed_Cod |  |  |  | 0.023 |  | (0.015) |
| Age * Wild_Cod |  |  |  | 0.029 | ** | (0.013) |
| Age * Monk |  |  |  | 0.038 | ** | (0.019) |
| Age * Pangasius |  |  |  | -0.010 |  | (0.024) |
| Single * Salmon |  |  |  | 0.974 | *** | (0.320) |
| Single * Farmed_Cod |  |  |  | 0.639 |  | (0.421) |
| Single * Wild_Cod |  |  |  | 0.676 | * | (0.363) |
| Single * Monk |  |  |  | 0.238 |  | (0.527) |
| Single * Pangasius |  |  |  | 1.212 | * | (0.663) |
| Standard Deviations |  |  |  |  |  |  |
| Salmon | 1.759 | *** | (0.146) | 1.680 | ** | (0.134) |
| Farmed_Cod | 2.229 | *** | (0.222) | 2.138 | *** | (0.199) |
| Wild_Cod | 2.093 | *** | (0.173) | 1.888 | *** | (0.160) |
| Monk | 2.652 | *** | (0.362) | 2.499 | *** | (0.287) |
| Pangasius | 3.260 | *** | (0.466) | 3.186 | *** | (0.479) |
| Number of observations | 11380 |  |  | 11380 |  |  |
| Number of participants | 178 |  |  | 178 |  |  |
| LR Chi-Squared | 961.61 |  |  | 854.21 |  |  |
| Log-likelihood | -2534.4973 |  |  | -2508.7689 |  |  |
| AIC | 5126.9946 |  |  | 5105.5378 |  |  |

Significance codes: $\quad \alpha=0.01^{* * *} \quad \alpha=0.05^{* *} \quad \alpha=0.1^{*}$

### 6.1. Preference Heterogeneity

The parameter estimates for Salmon are positive and significant in both models. The standard deviations for all the fish types are significantly different from zero at the $1 \%$ level. This suggests heterogeneous preferences for all the fish types, despite controlling for correlations in preferences. Salmon has, in addition to the highest expected utility, the narrowest distribution. This can be seen from the relatively low standard deviation in Table 6.1, indicating more homogeneous preferences for salmon relative to the other fish types. Pangasius has the lowest expected utility and the widest distribution. This can be seen from the relatively large standard deviation in Table 6.1. This indicates more heterogeneous preferences for pangasius relative to the other fish types.

### 6.2. Price Effect

The price parameter is negative and significant at the $1 \%$ level in both models. This implies that when the price of the product increases, the utility an individual obtains from choosing it decreases. This is expected, and in accordance with classical microeconomic theory of demand. This also supports Nilssen and Monfort (2000) findings of French consumers being price conservative.

### 6.3. Tail Effect

In both models there are significant negative signs on the parameters for the tail interaction terms. The participants are most negative to tail cuts of farmed cod and least negative to tail cuts of salmon. The parameter estimates for the tail interaction terms are quite similar in the two models. This implies that the interaction terms including Tail are not correlated with the interaction terms including the consumer characteristics. The unambiguous negative parameter estimates for fillets of tail cut could raise questions to producers on how to most profitably utilize the tail fillets.

### 6.4. Model 1 - Product Attribute Model

In Model 1, where no consumer characteristics interaction terms are included, we find positive and significant utility from choosing salmon, farmed cod, wild cod and monk. The coefficient for Pangasius is negative. Hence, on average the participants prefer to choose the NOT alternative over pangasius. The parameter estimates for Farmed Cod, Wild Cod and Monk are positive and significant. The parameter estimate for Monk is not significantly different from those of neither Farmed Cod (Wald p-value 0.2081) nor Wild Cod (Wald pvalue 0.6657 ). This indicates that the utility from choosing monk is not significantly different from the utility from choosing cod. The parameter estimate for Wild Cod is significantly higher than that of Farmed $\operatorname{Cod}$ (Wald $p$-value 0.0329 ). Hence, they value wild cod higher than farmed cod. Also, the coefficient for a tail fillet of wild cod is significantly higher than the coefficient for a tail fillet of farmed cod (Wald $p$-value 0.0007 ).

### 6.5. Model 2 - Product Attribute and Consumer Characteristics Interaction Model

Contrary to the results from Model 1, the parameter estimates for Farmed Cod, Wild Cod and Monk are positive but not significantly different from zero, when Income, Age and Single are controlled for.

### 6.5.1. Income Effect

Higher income positively affects the preferences for farmed cod at the $10 \%$ level and wild cod at the $1 \%$ level. This indicates that higher income groups have a higher preference for cod relative to those with lower income. Apart from farmed cod and wild cod, there are no significant differences in the preferences for fish between the low and high income groups. Since $92 \%$ of the respondents eat fish for lunch or dinner at home weekly (Appendix I, Table A.11.1), fish may be regarded as a basis food, and this may explain why the income effect for fish is relatively low.

### 6.5.2. Age Effect

The preferences for salmon are decreasing with age. Farmed salmon was introduced to the French market in the late 1970s, but did not become common until the 1990s (Nilssen \& Monfort 2000). Hence, it is likely that a greater share of young people relative to old people has adopted salmon as part of their eating habits. Older people tend to have higher preferences for wild cod and monk relative to younger people. Farmed cod appears to appeal equally to young and old people.

These results accord well with the findings of Alfnes and Rickertsen (2008b) from an experiment in Dijon in December 2007. They find that people below the age of 60 years gave higher taste scores to salmon relative to people over 60 . They also find that both wild cod and monk is higher ranked among older consumers compared to younger consumers.

### 6.5.3. Single versus Married/Cohabiting

Single people obtain a higher utility from choosing salmon compared to married/cohabiting people. The survey results show that the majority thinks salmon and cod are easier to prepare than monk and pangasius ( $93 \%, 85 \%, 43 \%$ and $33 \%$ respectively (Appendix I, Table A.11.2)). It is likely to believe that single households prefer to cook food that is easy and fast to prepare. Hence, it comes as no surprise that single people have higher preferences for salmon relative to married/cohabiting people. However, time spent on preparing meals is decreasing in the whole French population (Nilssen \& Monfort 2000). Seen from another perspective, salmon is regarded to be more expensive than both cod and pangasius (Table 6.7.1, column 7), and it is likely that single households have less disposable income than cohabiting households. In fact, the majority of the singles (86\%) in the experiment belong to the low income group. Single people are also more positive to pangasius relative to married/cohabiting people. This is expected since pangasius is cheaper than the other fish types.

Of particular interest is that single people obtained a higher utility from choosing wild cod relative to married/cohabiting people. This is odd, since a greater fraction of the single people perceives farmed fish as healthier ( $45 \%$ ) than do the married/cohabiting people ( $25 \%$ ) (Appendix I, Table A.11.3). Additionally, more married/cohabiting people agree to the
statement that wild fish is healthy food (87\%) than do single people (73\%) (Appendix I, Table A.11.3). From the survey responses, single people appear to have higher confidence in farmed fish than married/cohabiting people do. Farmed cod, however, appears to appeal equally to both single and married/cohabiting people.

### 6.6. Correlations in Preferences - The Covariance Matrices

The correlations in preferences for the fish types are displayed in Table 6.6.1. In both models many significant and positive correlations between the preferences for the fish types occur. This implies that those who get an above average utility from choosing e.g. salmon also get an above average utility from choosing e.g. farmed cod, wild cod and pangasius. Particularly high is the correlation between wild cod and monk, suggesting that these are good substitutes for each other. The correlation between farmed cod and wild cod is also relatively high.

Table 6.6.1. Covariance Matrices for Model 1 and Model 2

| Model 1 | Salmon | Farmed Cod | Wild Cod | Monk | Pangasius |
| :--- | :---: | :--- | :--- | :--- | :---: |
| Salmon | $3.096^{* * *}$ |  |  |  |  |
| Farmed Cod | $2.006^{* * *}$ | $4.971^{* * *}$ |  |  |  |
| Wild Cod | $1.519^{* * *}$ | $3.579^{* * *}$ | $4.380^{* * *}$ |  |  |
| Monk | $0.995^{* *}$ | $3.757^{* * *}$ | $4.358^{* * *}$ | $7.034^{* * *}$ |  |
| Pangasius | $2.701^{* * *}$ | 1.253 | 0.0661 | 1.277 | $10.625^{* * *}$ |


| Model 2 | Salmon | Farmed Cod | Wild Cod | Monk | Pangasius |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Salmon | $2.823^{* * *}$ |  |  |  |  |
| Farmed Cod | $1.784^{* * *}$ | $4.572^{* * *}$ |  |  |  |
| Wild Cod | $1.276^{* * *}$ | $3.153^{* * *}$ | $3.564^{* * *}$ |  |  |
| Monk | 0.546 | $3.113^{* * *}$ | $3.802^{* * *}$ | $6.244^{* * *}$ |  |
| Pangasius | $2.423^{* * *}$ | $1.199^{*}$ | 0.443 | 0.819 | $10.152^{* * *}$ |
| Significance codes: |  | $\alpha=0.01^{* * *}$ | $\alpha=0.05^{* *}$ |  | $\alpha=0.1^{*}$ |

The tables show the correlations in unobserved factors over choice scenarios between the fish types, expressed by covariances.

The significant positive correlations may reflect that the participants, who chose fish instead of NOT in one choice scenario, were likely to do likewise in other choice scenarios. And the participants, who were likely to choose NOT in one choice scenario, were likely to do so in the other choice scenarios. Hence, these results must be interpreted with caution. For a distribution of how frequently the participants chose NOT, see Figure 6.6 .1 below.


Figure 6.6.1. Distribution of How Frequently the Participants Chose the NOT Alternative

### 6.7. WTP Estimates for Model 1 - Product Attribute Model

Of more economic interest than the somewhat vague concepts of "utility" and "preferences," is the willingness to pay. The second column of Table 6.7 .1 shows the estimated WTP values from Model 1, where no consumer characteristics are controlled for. The third and fourth columns show the lower and upper limits of the $95 \%$ confidence intervals for the WTP values. The fifth column shows prices per kilogram of the five fish types found in grocery stores in Dijon in May 2008 ${ }^{6}$. The sixth column shows the weighted average price of the fish over all choice scenarios in the experiment. The last column shows the average price per kilogram guessed by the participants when asked about this in the survey.

Table 6.7.1. WTP Values for Model 1 Compared with Prices Found in Grocery Stores, Average Price Over Scenarios, and Average Guess by the Participants

| Variables | $\begin{gathered} \text { Mean } \\ \text { WTP/KG } \end{gathered}$ | 95\% Confidence Interval |  | Average <br> Price/KG <br> Found in Grocery Stores | Weighted Average Price/KG Over all Choice Scenarios | Guessed Price/KG by the Participants |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Lower <br> Limit | Upper <br> Limit |  |  |  |
| Salmon | 13.69 | 12.19 | 15.18 | 16.00 | 14.88 | 14.72 |
| Farmed cod | 10.78 | 8.36 | 13.20 | 24.25* | 20.78 | 13.67 |
| Wild cod | 13.35 | 11.45 | 15.25 | 24.25* | 20.79 | 13.67 |
| Monk | 12.76 | 10.03 | 15.49 | 34.00 | 27.88 | 21.17 |
| Pangasius | -6.94 | -11.72 | -2.15 | 10.00** | 10.82 | 9.34 |
| Tail * Salmon | -1.97 | -2.96 | -0.97 |  |  |  |
| Tail * Farmed Cod | -6.55 | -8.15 | -4.96 |  |  |  |
| Tail * Wild Cod | -3.05 | -4.33 | -1.77 |  |  |  |

The confidence interval is of $95 \%$ confidence level.

* It is unclear whether the prices found for cod were farmed or wild.
** The price for pangasius were the price for frozen pangasius.

The estimated WTP for salmon is 13.69 EUR per kilogram. This price is fairly close to the average price of 16.00 EUR per kilogram of an equivalent fillet of salmon found in grocery stores in Dijon at the time of the experiment. The estimated WTP for salmon is also close to the weighted average price for salmon over all choice scenarios (14.88 EUR/KG) and the

[^4]average price guessed by the participants (14.72 EUR/KG). For all the other fish types the WTP values are below the prices found in grocery stores.

Estimated WTP for pangasius is -6.94 EUR per kilogram. Theoretically this means that an individual, on average, should be compensated 6.94 EUR to accept one kilogram of pangasius. In practice it means that, on average, an individual prefers the NOT alternative to buying pangasius at the given prices. A possible reason for the low WTP estimates is that quite many participants chose the NOT alternative in many choice scenarios (Figure 6.6.1). This will pull the price parameter downwards. None of the participants, however, chose the NOT alternative in all choice scenarios.

By converting the estimated utility distributions, i.e., the estimated average utilities and standard deviations, into WTP values, we can graphically see the preference heterogeneity for each fish type. In Figure 6.7.1 the distributions for salmon, farmed cod and pangasius are found in the graph to the left. The distributions for farmed cod, wild cod and monk are found in the graph to the right.


-ー- Salmon ——Farmed_Cod $\cdot$...... Pangasius


Figure 6.7.1. Distributions for the Random Parameters from Model 1 Converted to WTP Values

As the graph to the left in Figure 6.7.1 illustrates, the WTP for pangasius is lower than that of farmed cod and salmon. The large heterogeneity in WTP values for pangasius is reflected by the wide distribution. The relatively homogenous WTP values for salmon are illustrated by
the narrower curve. As shown in the graph to the right in Figure 6.7.1, the estimated WTP for farmed cod, wild cod and monk are quite similar. Their distributions are also relatively similar, indicating that the share of the population that is willing to pay more than average and the share of the population that is willing to pay less than average for these fish types, are relatively equally distributed in the population.

### 6.8. WTP Estimates for Model 2 - Product Attribute and Consumer Characteristics

## Interaction Model

The WTP values obtained from Model 2, which includes the consumer characteristics Income, Age and Single, are presented in Table 6.8.1.

Table 6.8.1. WTP Results for Model 2 - Product Attribute and Consumer Characteristics Interaction Model

|  |  | $95 \%$ Confidence |  |
| :--- | :---: | :---: | :---: |
|  |  | Interval |  |
|  | Mean | Lower | Upper |
| Variables | 16.94 | 11.96 | 21.92 |
| Salmon | 4.22 | -3.10 | 11.54 |
| Farmed_Cod | 5.09 | -0.92 | 11.11 |
| Wild_Cod | 4.77 | -4.31 | 13.86 |
| Monk | -6.93 | -18.30 | 4.44 |
| Pangasius | -1.96 | -2.95 | -0.97 |
| Tail * Salmon | -6.59 | -8.17 | -5.00 |
| Tail * Farmed_Cod | -3.06 | -4.33 | -1.78 |
| Tail * Wild_Cod | 2.40 | -0.76 | 5.56 |
| Income *Salmon | 3.92 | -0.32 | 8.15 |
| Income *Farmed_Cod | 5.62 | 2.11 | 9.14 |
| Income * Wild_Cod | 1.37 | -3.84 | 6.58 |
| Income *Monk | 4.20 | -2.11 | 10.51 |
| Income *Pangasius | -0.10 | -0.20 | 0.00 |
| Age * Salmon | 0.10 | -0.03 | 0.24 |
| Age * Farmed_Cod | 0.13 | 0.02 | 0.25 |
| Age * Wild_Cod | 0.17 | 0.00 | 0.34 |
| Age * Monk | -0.05 | -0.26 | 0.16 |
| Age * Pangasius | 4.43 | 1.56 | 7.29 |
| Single * Salmon | 2.91 | -0.85 | 6.66 |
| Single * Farmed_Cod | 3.07 | -0.17 | 6.32 |
| Single * Wild_Cod | 1.08 | -3.62 | 5.78 |
| Single * Monk | -0.40 | 11.42 |  |
| Single * Pangasius |  |  |  |

As an example, the estimated WTP for a kilogram of loin of farmed cod for a 45 year old, single person, who belongs to the low income group, is:
$4.22-6.59 * 0+3.92 * 0+0.10 * 45+2.91 * 1=11.63$ EUR/KG.

Table 6.8.2 shows WTP values for different consumer groups. Age is set to 47, which is the sample average. The most striking features of Table 6.8.2 are the large differences in WTP values between fillets of loin and fillets of tail cut of farmed cod. All consumer groups, given that Age is 47, are willing to pay about twice as much for a loin of farmed cod than a tail fillet of farmed cod.

Another interesting finding is that people in the high income group, regardless of marital status, and given that Age is 47, are willing to pay more for a loin of wild cod than a loin of salmon. This again demonstrates the positive income effect on wild cod found in Table 6.1.

All WTP values for farmed cod are below the prices found in grocery stores in Dijon in May 2008, given that Age is 47 .

Table 6.8.2. Examples of WTP Values in EUR/KG per Consumer Group (Age $=47$ )

|  | Low Income <br> Single | High Income <br> Single | Low Income <br> Married | High Income <br> Married |
| :--- | :---: | :---: | :---: | :---: |
| Salmon Loin | 16.57 | 18.97 | 12.14 | 14.55 |
| Salmon Tail | 14.61 | 17.01 | 10.18 | 12.59 |
| Farmed Cod Loin | 12.00 | 15.92 | 9.10 | 13.01 |
| Farmed Cod Tail | 5.42 | 9.33 | 2.51 | 6.43 |
| Wild Cod Loin | 14.41 | 20.03 | 11.34 | 16.96 |
| Wild Cod Tail | 11.35 | 16.98 | 8.28 | 13.90 |
| Monk | 14.02 | 15.39 | 12.94 | 14.31 |
| Pangasius | -3.64 | 0.56 | -9.15 | -4.95 |

By dividing consumers into age groups, we can see differences in WTP values between younger and older consumers. In Table 6.8.3 the consumers are divided into one group where Age is set to 30, and another group where Age is set to 60. All consumers are assumed to be married/cohabiting, that is, Single $=0$.

Older consumers have higher WTP values for both farmed and wild cod relative to younger consumers. However, from Table 6.1 we find no significant age effect on farmed cod, but the $p$-value is 0.12 , hence close to significance at the $10 \%$ level (Appendix I, Table A.12.2.1). The married/cohabiting participants aged 60 have, regardless of income, higher WTP values for loins of wild cod than loins of salmon.

All WTP values for farmed cod are below the prices found in grocery stores in Dijon in May 2008.

Table 6.8.3. Examples of WTP Values in EUR/KG per Consumer Group (Single $=0$ )

|  | Age 30 <br> Low Income | Age 60 <br> Low Income | Age 30 <br> High Income | Age 60 <br> High Income |
| :--- | :---: | :---: | :---: | :---: |
| Salmon Loin | 13.88 | 10.82 | 16.28 | 13.22 |
| Salmon Tail | 11.92 | 8.86 | 14.32 | 11.26 |
| Farmed Cod Loin | 7.33 | 10.45 | 11.25 | 14.36 |
| Farmed Cod Tail | 0.75 | 3.86 | 4.66 | 7.78 |
| Wild Cod Loin | 9.08 | 13.06 | 14.70 | 18.69 |
| Wild Cod Tail | 6.02 | 10.01 | 11.64 | 15.63 |
| Monk | 9.98 | 15.19 | 11.36 | 16.57 |
| Pangasius | -8.35 | -9.77 | -4.15 | -5.57 |

It is worth noting that the $95 \%$ confidence intervals for the WTP estimates in Model 2 (Table 6.8.1) are wide. The WTP values for all fish types but salmon range from a negative to a positive value. This indicates that there is uncertainty related to the estimates, and the results must be interpreted with caution.

## 7. Discussion

In the next sections we discuss the results and relate them to the answers from the survey. The terms "survey respondents" and "participants" are used interchangeably.

Generally, the preferences and the WTP values for farmed cod are lower than those of wild cod. Could attitudes towards production method (fish farming vs. wild catching) and environmental concerns be underlying factors determining these differences? In the survey, $83 \%$ agrees to the statement that wild fish is healthy food, but only $31 \%$ agrees to the equivalent statement for farmed fish (Appendix I, Table A.11.3). This implies that the participants regard wild cod to be healthier than farmed cod, and perhaps they are willing to pay a price premium for the fish they consider the healthier. This suggests that producers of farmed cod potentially could improve the image of farmed fish by focusing on healthiness.
$60 \%$ of the survey respondents agrees to the statement that wild fish is safe to eat, but only one third (33\%) agrees to the equivalent statement for farmed fish (Appendix I, Table A.11.3). However, this contradicts the unambiguous positive attitudes towards salmon, since almost all salmon sold on the French market is farmed. To which extent the participants (and French consumers in general) are aware of the salmon being farmed is unclear.

### 7.1. Environment and Animal Welfare - Attitude-Behavior Gap?

More people have reported that they are concerned about the environmental impact of the production of wild fish (77\%) than of farmed fish (60\%) (Appendix I, Table A.11.3). This suggests that the participants do not refrain from eating farmed fish on environmental grounds. When it comes to environmental sustainability, $76 \%$ reported that they are concerned about the environmental sustainability of fisheries, and $70 \%$ reported the same for fish farming (Appendix I, Table A.11.3).

Regarding animal welfare, the survey responses reveal no significant difference in the attitudes towards the welfare of farmed fish and the welfare of wild fish (Welch $p$-value $=$ 0.3779 ) (Appendix I, Table A.11.1.1). This indicates that the attitudes towards fish farming are not significantly stronger than the attitudes towards wild fish catching.

However, more than half of the participants report that they are concerned about the welfare of farmed as well as wild fish. This comes as no surprise, since it is easy to agree to such a statement when answering a survey (Appendix I, Table A.11.3). Verbeke et al. (2007) found in a survey conducted on Flemish women that although consumers attach high perceived importance to sustainability and ethics related to fish, this perceived importance is not correlated with fish consumption or attitudes towards fish eating. Attitudes alone are often a poor predictor of marketplace behavior. The survey responses from this experiment indicate that beliefs about food safety and health perceptions are the main determinants for favoring wild fish to farmed fish. Hence, we might find an attitude-behavior gap among the French consumers as well.

### 7.2. Region of Origin

In general there is a positive view of fresh farmed fish from France (76\%) and Northern Europe (72\%) and wild fish from the Atlantic North (86\%) (Appendix I, Table A.11.1). There is attached high skepticism towards fresh farmed fish from third world countries. $65 \%$ does not have a positive view of fresh farmed fish from third world countries (Appendix I, Table A.11.1). According to a marketing survey conducted by Marint Verdiskapingsprogram, it is unheard of to write on a restaurant menu in France that the fish is farmed. It is, however, common to write the country of origin on the menu (Solheim 2010). This supports the notion that knowledge about the origin of the fish is a highly valued criterion in France.

### 7.3. Competition from Pangasius

From the RC results and the survey answers it is evident that the preferences and WTP for cod are greater than those of pangasius. Two thirds of the respondents agree to the statement that cod tastes good, while only $19 \%$ agrees to the equivalent statement for pangasius. In fact, $61 \%$ disagrees to the statement that pangasius tastes good (Appendix I, Table A.11.2). About half of the respondents regard cod as safe to eat, but only $10 \%$ regards pangasius as safe to eat (Appendix I, Table A.11.2). This may be because pangasius is from Vietnam, as the majority does not have a positive view of fresh farmed fish from third world countries.

The correlations in preferences between salmon, wild cod, farmed cod and monk were generally positive and significant (Table 6.6.1). However, the results from Model 1 suggest no correlations in the preferences between pangasius and farmed cod and wild cod. This suggests that those who obtain an above average utility from choosing farmed or wild cod do not necessarily obtain an above average utility from choosing pangasius.

Despite the somewhat unenthusiastic attitudes towards pangasius among the sample in this experiment, it is worth noting that pangasius is the ninth most consumed fish in the US today. A great advantage the pangasius has relative to other fish species is that it can breathe air; hence it can be produced in great volumes with little space (Greenberg 2011). Asian labor costs are in general lower, and the environmental restrictions are often more lax than in Europe. Hence, European fish farmers may face real challenges from the Asian fish farming industry.

### 7.4. Competition from Salmon

Knowledge about substitution effects between salmon and cod are important for both salmon and cod farmers. Our results suggest that, on average, the preferences for salmon outweigh those of both farmed and wild cod, despite the perceived relative higher price of salmon (Table 6.7.1, column 7). The significant and positive correlations between salmon, farmed cod and wild cod do, however, indicate that those who like salmon also like cod (Table 6.6.1). This may imply that salmon and cod could be substitutes. On the other hand, the correlations between farmed cod, wild cod and monk are even higher, indicating that cod (both farmed and wild) faces stronger competition from monk than from salmon.

Asche and Hannesson (1997) find that salmon does not compete on the whitefish market in France. They argue that salmon is often consumed as luncheons and as starters, while whitefish are more traditionally consumed as main dishes. Since salmon and whitefish are not consumed in similar product forms, they do not compete with each other. These findings are from 1997, and are based on data from 1983 to 1995. The consumption patterns in France may have changed considerably since then.

### 7.5. Farmed Cod versus Wild Cod - External Validity?

Through the tasting session, the BDM, and the RC experiment the participants were exposed to a distinction between farmed and wild cod. However, neither wild salmon nor wild pangasius were alternatives. Had wild salmon been an option, it may be that the WTP values for wild salmon would have been even higher than the WTP values already obtained for (farmed) salmon.

Despite the fractional factorial design, the visible distinction between wild and farmed cod may have biased the choices towards preferring wild to farmed cod. In an ordinary supermarket situation though, consumers make choices based on habits, and may pay less attention to the production method. It is likely that the distinction between farmed and wild cod is not that clear in the field, i.e., outside the laboratory. Hence this experiment may suffer from some lack of external validity. However, Chang et al. (2009) found a high level of external validity in non-hypothetical methods when comparing different preference elicitation methods for the product categories ground beef, wheat flour and dishwashing liquid. It is also
worth noting that in a tasting session of a similar experiment in Dijon in December 2007, Alfnes and Rickertsen (2008b) found that $55 \%$ of the participants gave farmed cod a higher taste score than wild cod.

### 7.6. Factors That May Bias WTP

In this RC experiment the only choice options were fish or NOT. This enabled us to examine the relative WTP for one fish type over the other. In an ordinary supermarket situation there are also other options like meat, egg, chicken, vegetables etc. How will this absence of outside options influence the results? Alfnes et al. (2006) argue that this may cause the participants to choose the NOT alternative too seldom. This may affect the price parameter and bias the WTP upwards. Even though real economic incentives were introduced by letting one choice scenario be binding, the price of the binding scenario makes up a very small fraction of a total household budget, especially when the participants knew they would get paid to participate. It follows that the binding scenario might not outweigh the upward bias from the absence of outside options.

Another aspect of this experiment is that many participants may find it exiting to take part in an experiment like this. Taking part in such an experiment may temporarily increase their desire for fish. This "new desire" may increase their motivation to buy fish that specific day, and bias WTP upwards. But when they go to the grocery store the next day, the "new desire" might be gone, and WTP might be at a lower level than during the experiment. The absence of outside options, and a possible new earned "desire" for fish, could be possible reasons for an upward bias of WTP.

On the other hand, it is probable that a great share of the participants did not intend to purchase fish on the exact day of the experiment. Corrigan and Rousu (2008) found that consumers who intended to buy bananas on the same day they took part in an experiment, had WTP values closer to the market price than the consumers who did not intend to buy bananas the same day. Those who intended to buy bananas on the day of the experiment had WTP values above those who did not intend to buy bananas. Hence, it is likely that the majority of the 178 participants behaved as non-buyers at the time of the experiment. A fresh fillet of fish is liable to rot if not refrigerated shortly after acquisition. The participants might find bringing with them a fresh fillet of fish as unpractical, and only choose a fillet in a
choice scenario if they felt it were a remarkably good deal. This may explain why many participants chose the NOT alternative in many choice scenarios, which biases WTP downwards. Had the experiment involved a nonperishable good, e.g. chocolate bars, which could easily be stored until its consumption value is higher, we might obtain higher WTP values.

Alfnes and Rickertsen (2011) recommend not including participants with a nonresponse to all alternatives, because those participants do not reveal anything about their relative valuations of the products included in the experiment. According to microeconomic theory, only relative prices matter. In our experiment there were no non-responses, but ten participants chose a fillet only once and chose NOT in all other choice scenarios (Figure 6.6.1). This may have biased the WTP results somewhat downwards.

### 7.7. Comparing RC with Other WTP Elicitation Methods

As mentioned in the introduction, other WTP elicitation methods are available, such as EA and stated choice experiments. In an EA participants are asked to be price makers, while in an RC experiment they are asked to be price takers. Being a price maker may deflate WTP, whereas being a price taker may inflate WTP. As Alfnes and Rickertsen (2011) point out, a weakness with the RC method compared with the EA is that the WTP is not directly observable. The WTP must be estimated based on the choices all the participants make. "Hence, the estimated WTP for each participant is affected by the responses of other participants and sensitive to the model specification" (Alfnes \& Rickertsen 2011). For instance, if one participant obtains an infinitely high utility from choosing salmon and does not consider price in his or her choices, the WTP values of the other participants will be affected. However, the WTP values from this experiment were lower than the prices found in grocery stores. This indicates that the results are not inflated above market price.

In stated choice experiments the respondents are, like in an RC experiment, price takers. Stated choice experiments differ from RC experiments in that each choice is hypothetical or non-consequential, hence real economic incentives are absent. Chang et al. (2009) find that non-hypothetical elicitation methods outperform hypothetical experiments. WTP values from non-consequential experiments are found to almost always exceed WTP values from consequential elicitation methods (Gracia et al. 2011). This is known as the hypothetical bias.

When respondents are aware of the fact that the choices they make have no economic consequences, they tend to accept higher prices than when they make inconsequential choices. Hence, the results from this experiment have empirical support of being closer to the true WTP of the consumers than would have been the case in a stated choice experiment.

### 7.8. Alternative Model Specifications

This experiment's main objectives were to study consumer preferences and WTP for the five fish types, and examine how different consumer groups value these. To study this, several approaches are available. For an even better understanding of how individuals value one fish type relative to another fish type, participants that answered NOT in more than a certain number of choice scenarios could be omitted from the analysis (see e.g. Figure 6.6.1). The downside of this is that the dataset would be reduced.

In our analysis the Tail variable was not interacted with the different consumer groups. A further investigation of how the different consumer groups value a fillet of tail cut could be a topic for further research.

Of the five fish types presented, three were farmed and two were wild. To better understand French consumers' attitudes towards farmed fish, one could include a variable indicating if the fish is farmed or not.

The survey answers revealed that the participants did care about the region of origin of the fish. Hence including variables relating to region of origin could be of interest.

Before the RC experiment took place, the participants had been through a tasting session. Hence they were familiar with the taste of the different types of fish. Including taste scores as variables could also possibly lead to some interesting findings.

## 8. Conclusion

In this article we have analyzed data from an RC experiment. We investigated the French consumer preferences and WTP for fish. We estimated random utility functions by mixed logit to capture preference heterogeneity in the population. We estimated one product attribute model (Model 1) and one model which included both products attributes and consumer characteristics (Model 2).

WTP values varied both with the attributes of the fish and with the consumer characteristics. The participants were willing to pay less for a fillet of tail than for a loin. Particularly low was the WTP for a tail fillet of farmed cod.

The WTP for wild cod appears to be slightly higher than that of farmed cod. Salmon was found to be the most desirable choice among the participants. On average, we found positive WTP values for all the fish types except for pangasius. The preferences for all the fish types were heterogeneous, which implies that certain segments of the population have WTP values above average, whereas other segments have WTP values below average. People with higher income are willing to pay a price premium for both farmed and wild cod. Higher age is associated with higher WTP for wild cod and monk, and lower WTP for salmon. Single households are willing to pay a price premium for salmon, wild cod and pangasius.

Predominantly, all WTP values were below the average price of the five fish types found in grocery stores in Dijon at the time of the experiment. Gender, education and presence of children in the household did not significantly affect the participants' choices, and accordingly, had no effect on WTP.

The participants in the experiment also answered a survey on fish likings, fish buying habits, attitudes toward production methods etc. The answers from the survey corresponded well to the results obtained from the RC experiment.

In the introduction we referred to two possible strategies for the future of the cod farming industry. The first strategy is to position itself at the higher price end of the market and create niche products. A second approach is to reduce production costs considerably, and increase the production volume. Our results indicate that the high income group is willing to pay a price premium for both farmed and wild cod. This speaks in favor of choosing the first strategy. However, the survey results suggest that the participants regard cod as both safer to
eat and tastier than pangasius. Hence, if cod farmers were able to reduce the price considerably, cod has a competitive advantage over pangasius. This speaks in favor of choosing the second strategy. Since it is unrealistic to assume that the cod farming industry can compete with the Asian whitefish farming industry on price, we believe the first strategy is more feasible.

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Appendix I

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## A.1. A Short Introduction to Lancastrian Consumer Theory

The theory behind the utility models from the article is based on Lancastrian consumer theory. Traditional consumer theory postulates that the good itself generates utility for a consumer. Lancaster, on the other hand, proposes that the goods are components of different attributes, and that the summation of the utilities associated with the attributes determines a person's utility for the good. Following Lancaster (1966), "the chief technical novelty lies in breaking away from the traditional approach that goods are the direct objects of utility and, instead, supposing that it is the properties or characteristics of the goods from which utility is derived." The attributes of fish in general are the fish types, and the five fish types used in our experiment had different cuts and prices. In the empirical specification of the utility functions in the article, we included product attributes and consumer characteristics. Although a consumer characteristic, like age, is not in itself an attribute of the product, a person's age may, however, affect how he or she perceives the product.

## A.2. Other WTP Elicitation Methods and Prior Research

To study consumer preferences and elicit WTP, other methods than real choice (RC) experiments are available, for example experimental auctions (EA), Becker-DeGrootMarschak (BDM) mechanisms or stated choice (SC) experiments. In EAs the participants take part in a bidding session. There are many versions of EAs, for example a second-price sealedbid auction, also known as a Vickrey auction. In a Vickrey auction participants submit sealed bids for a product. The highest bidder buys the product for the price of the second highest bid (Alfnes \& Rickertsen 2011).

The BDM mechanism is similar to an EA. A participant submits a sealed bid for a product. The sales price is determined by a draw from a distribution of numbers from zero to the highest anticipated bid. If the bid is higher than the drawn price, the participant buys the product for the price picked from the draw. Strategically a BDM is equal to an EA (Alfnes \& Rickertsen 2011). A special feature of the BDM is that it is possible for only one person to participate. EAs and BDMs might seem like unfamiliar situations for consumers. No prices are posted in advance.

In SC experiments, participants make hypothetical choices over a set of one or more choice scenarios. Participants are asked to pick the product they would have bought; given that it was a real life situation. Prices are posted in advance.

Corrigan and Rousu (2008) study the differences in consumers' WTP for perishable and nonperishable goods by the use of EAs. Loureiro and Umberger (2004) use an SC experiment to study which beef attributes that affect consumer preferences, and the corresponding effect on WTP. Wolf et al. (2011) use an SC experiment to estimate consumers' WTP for half and whole gallons of milk. They study consumers' responses on attributes such as labeling with information on rbST-content. Another SC experiment that studies consumers' response on product labeling is James et al. (2009). They assess WTP values of organic, local and nutrition attributes on applesauce. A study of consumers' responses to animal welfare was conducted by Tonsor et al. (2009). They use an SC experiment with labeling on the use of gestation crates in the production as an attribute.

## A.3. The Mixed Logit Model

The standard logit model estimates the logarithm of the odds of an outcome as:
(1) $\log \left(\frac{\pi_{j}}{\pi_{j}}\right)=\beta_{j 0}+\beta_{j 1} x_{1}+\beta_{j 2} x_{2}+, \ldots,+\beta_{j k} x_{k}, \quad j=1, \ldots, J-1$,
where $J$ is the baseline category, and $\pi_{j}$ are the outcome probabilities given by:
(2) $\pi_{j}=\frac{e^{\beta_{j} x}}{\sum_{j=1}^{J} e^{\beta_{j} x}}$.
(Agresti 2007, p. 174-176)
When estimating a discrete choice random utility model, and assuming it is linear in parameters, one can interpret the logarithm of the odds as the utility an individual obtains
from making a choice. The linear and ordinal nature of the logarithm of the odds makes a utility model straight forward to interpret.

The mixed logit model is more flexible than the standard logit. It obviates the limitations of the standard logit model in three ways: It allows for random taste variation, it takes unrestricted substitution patterns into account, and it allows for correlations in unobserved factors over time or choice scenarios (Train 2009, p. 134).

Any random utility model (RUM) can, to any degree of accuracy, be approximated by a mixed logit model with the right choice of variables and distribution for the random parameters (Train 2009, p. 142). The random parameters have, in addition to their expected values, a distribution chosen by the researcher (chosen to be normal in the article). Mixed logit probabilities are the integrals of logit probabilities over the density of the random parameters (Train 2009, p. 135). The probability that individual $n$ chooses alternative $i$ in one given choice scenario is given by:
(3) $P_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta})=\int L_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta}) f(\boldsymbol{\eta}) d \boldsymbol{\eta}$,
where $\boldsymbol{\beta}$ is a vector of the fixed parameters and $\boldsymbol{\eta}$ is a vector of the random parameters. The $\beta \mathrm{s}$ are to be interpreted as if they were standard logit. $f(\boldsymbol{\eta})$ is the density function of the random parameters and $L_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta})$ is the standard logit probability evaluated at parameters $\boldsymbol{\beta}$ and $\boldsymbol{\eta}$ :
(4) $L_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta})=\frac{e^{\beta x_{n i}+\eta_{n} z_{n i}}}{\sum_{j=1}^{J} e^{\beta x_{n j}+\eta_{n} z_{n j}}}$.

Note that $i$ is an element in the array $j=1,2, \ldots, J$. In the models presented in the article $i$ represents the five fish types. Equation (3) can be extended to allow for repeated choices over time or choice scenarios $s$ (Train 2009, p. 145). Consider a sequence of choices $\boldsymbol{i}=$ $\left\{i_{1}, i_{2}, \ldots, i_{S}\right\}$. The probability that an individual makes this exact sequence of choices over a set of choice scenarios is:
(5) $P_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta})=\int L_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta}) f(\boldsymbol{\eta}) d \boldsymbol{\eta}$,
where $L_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta})$ is the product of the logit probabilities evaluated at parameters $\boldsymbol{\beta}$ and $\boldsymbol{\eta}$ :
(6) $L_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta})=\prod_{s=1}^{S}\left[\frac{e^{\beta x_{n i s}+\eta_{n} z_{n i s}}}{\sum_{j=1}^{J} e^{\beta x_{n j s}+\eta_{n} z_{n j s}}}\right]$.

There were sixteen choice scenarios in the RC experiment presented in the article. The data is therefore treated as panel data. The estimated models contain both fixed and random parameters. Since we have assumed the random parameters to be normally distributed with mean $\boldsymbol{\mu}_{\eta}$, and covariance matrix $W$, the density of the random parameters are given by $\phi\left(\boldsymbol{\eta} \mid \boldsymbol{\mu}_{\eta}, W\right)$. Allowing for correlations in unobserved factors over choice scenarios makes $W$ a covariance matrix, rather than just the standard deviations of the random parameters. The standard deviations can be obtained by taking the square root of the diagonal elements of the covariance matrix. The probability of a given sequence over the sixteen choice scenarios $\boldsymbol{i}=\left\{i_{1}, i_{2}, \ldots, i_{16}\right\}$ is given by:
(7) $P_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta})=\int L_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta}) \phi\left(\boldsymbol{\eta} \mid \boldsymbol{\mu}_{\eta}, W\right) d \boldsymbol{\eta}$,
where $L_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta})$ is the standard logit probability evaluated at parameters $\boldsymbol{\beta}$ and $\boldsymbol{\eta}$ :

$$
\text { (8) } L_{n i}(\boldsymbol{\beta}, \boldsymbol{\eta})=\prod_{s=1}^{16}\left[\frac{e^{\beta x_{n i s}+\eta_{n} z_{n i s}}}{\sum_{j=1}^{5} e^{\beta x_{n j s}+\eta_{n} z_{n j s}}}\right] \text {. }
$$

Note that the sum in the denominator in equation (8) ranges from one to five since there are five different fish types in the experiment. The vector for the random parameters has an estimated mean $\boldsymbol{\mu}_{\eta}$, and covariance matrix $W$. Since the chosen distribution for the random parameters is normal, equation (7) must be integrated over all values of $\boldsymbol{\eta}$, that is, from minus infinity to plus infinity. The density $\phi\left(\boldsymbol{\eta} \mid \boldsymbol{\mu}_{\eta}, W\right)$ is given by the normal distribution. Mixed
logit probabilities cannot be calculated analytically. A numerical approach is necessary (Train 2009, p. 144).

## A.4. The Econometric Model and Model Setup

The utility model in the article can be thought of as five different utility functions; one for each fish type $i$. Each fish type is assigned a number from one to five, that is, $i=\{1=$ Salmon, $2=$ Farmed Cod, $3=$ Wild Cod, $4=$ Monk, $5=$ Pangasius $\}$. The utility of choosing NOT is normalized to zero. Hence, the utility one obtains from choosing the different fish types is the utility one obtains compared to choosing NOT.

The socio-demographic variables (the consumer characteristics) do not vary over choice scenarios. To capture the effect of the consumer characteristics variables, one must multiply them with the variables representing the fish types. This accords well with thinking of the utility functions as one for each fish type.

The variables representing each fish type are defined by the $\boldsymbol{z}_{i}$-vectors:

$$
z_{1}=\boldsymbol{z}_{s a}=\left[\begin{array}{l}
1 \\
0 \\
0 \\
0 \\
0
\end{array}\right], z_{2}=\boldsymbol{z}_{f c}=\left[\begin{array}{l}
0 \\
1 \\
0 \\
0 \\
0
\end{array}\right], z_{3}=\boldsymbol{z}_{w c}=\left[\begin{array}{l}
0 \\
0 \\
1 \\
0 \\
0
\end{array}\right], \boldsymbol{z}_{4}=\boldsymbol{z}_{m o}=\left[\begin{array}{l}
0 \\
0 \\
0 \\
1 \\
0
\end{array}\right], z_{5}=\boldsymbol{z}_{p a}=\left[\begin{array}{l}
0 \\
0 \\
0 \\
0 \\
1
\end{array}\right]
$$

To estimate the consumer characteristics effect on the preferences for fish, interaction terms are necessary. For example, to estimate the age effect, one must multiply the variable Age with the $\boldsymbol{z}_{i}$-vectors. That is:

$$
\begin{aligned}
& \text { Age } * \mathbf{z}_{1}=\text { Age } * \boldsymbol{z}_{s a}=\text { Age }\left[\begin{array}{l}
1 \\
0 \\
0 \\
0 \\
0
\end{array}\right], \\
& \text { Age } * \mathbf{z}_{2}=\text { Age } * \mathbf{z}_{f c}=\text { Age }\left[\begin{array}{l}
0 \\
1 \\
0 \\
0 \\
0
\end{array}\right], \\
& \text { Age } * \mathbf{z}_{3}=\text { Age } * \mathbf{z}_{w c}=\text { Age }\left[\begin{array}{l}
0 \\
0 \\
1 \\
0 \\
0
\end{array}\right], \\
& \text { Age } * \mathbf{z}_{4}=\text { Age } * \mathbf{z}_{m o}=\text { Age }\left[\begin{array}{l}
0 \\
0 \\
0 \\
1 \\
0
\end{array}\right], \\
& \text { Age } * \mathbf{z}_{5}=\text { Age } * \mathbf{z}_{p a}=\text { Age }\left[\begin{array}{l}
0 \\
0 \\
0 \\
0 \\
1
\end{array}\right] .
\end{aligned}
$$

For salmon, farmed cod and wild cod the cut was either a tail cut or a loin (front cut). Thus we get three vectors $\boldsymbol{y}_{i}$ representing the three fish types that had either tail cut or loin:

$$
\boldsymbol{y}_{1}=\boldsymbol{y}_{s a}=\left[\begin{array}{l}
1 \\
0 \\
0
\end{array}\right], \boldsymbol{y}_{2}=\boldsymbol{y}_{f c}=\left[\begin{array}{l}
0 \\
1 \\
0
\end{array}\right], \boldsymbol{y}_{3}=\boldsymbol{y}_{w c}=\left[\begin{array}{l}
0 \\
0 \\
1
\end{array}\right],
$$

To estimate the tail effect on preferences for fish, one must multiply the $\boldsymbol{y}_{i}$-vectors with the Tail variable:

$$
\begin{aligned}
& \text { Tail } * \boldsymbol{y}_{1}=\text { Tail } * \boldsymbol{y}_{s a}=\text { Tail }\left[\begin{array}{l}
1 \\
0 \\
0
\end{array}\right], \\
& \text { Tail } * \boldsymbol{y}_{2}=\text { Tail } * \boldsymbol{y}_{f c}=\text { Tail }\left[\begin{array}{l}
0 \\
1 \\
0
\end{array}\right], \\
& \text { Tail } * \boldsymbol{y}_{3}=\text { Tail } * \boldsymbol{y}_{w c}=\text { Tail }\left[\begin{array}{l}
0 \\
0 \\
1
\end{array}\right],
\end{aligned}
$$

Model 1 in the article, with no consumer characteristics included, is specified as:
(9) $E\left(U_{i}\right)=\alpha$ Price $+\boldsymbol{\mu}_{\eta} \boldsymbol{z}_{i}+\boldsymbol{\delta} \boldsymbol{y}_{i}$ Tail,
where $\alpha$ is the price coefficient, $\boldsymbol{\mu}_{\eta}$ is a vector of the expected values of the random parameters and $\boldsymbol{\delta}$ is a vector of coefficients for the interaction terms including Tail. Note that the $n$ subscript is omitted in equation (1) since no consumer characteristics are included in Model 1.

Model 2 in the article, which includes the consumer characteristics Income, Age and Single, is specified as:

$$
\begin{equation*}
E\left(U_{n i}\right)=\alpha \text { Price }+\boldsymbol{\mu}_{\eta} \boldsymbol{z}_{i}+\boldsymbol{\delta} \boldsymbol{y}_{i} \text { Tail }+\mathbf{z}_{i}\left[\boldsymbol{\gamma}_{1} \text { Income }_{n}+\boldsymbol{\gamma}_{2} \text { Age }_{n}+\right. \tag{10}
\end{equation*}
$$

$$
\left.\boldsymbol{\gamma}_{3} \text { Single }_{n}+\boldsymbol{\gamma}_{4} D N W A_{n}\right],
$$

where the $\boldsymbol{\gamma}$ are coefficient vectors for the socio-demographic interactions with each fish type. Since the consumer characteristics Income, DNWA, Age and Single are interacted with variables representing each fish type, they are multiplied by the $\boldsymbol{z}_{i}$ vector. The variable DNWA (Do Not Want to Answer) represents the people in the income group that did not want to reveal their income. This variable is included in every model where Income is a variable.

However, its parameter estimates are of limited interest and they are thus not presented in the article. For a thorough discussion of the $D N W A$ variable, see section A.6.1.1.

## A.4.1. A Numerical Example - Utility for Farmed Cod

Assume we are interested in the utility for farmed cod. Farmed cod was assigned the number 2 , i.e., $i=2$. The $\boldsymbol{y}$ and $\boldsymbol{z}$-vectors thus have the form:

$$
\boldsymbol{y}_{2}=\boldsymbol{y}_{f c}=\left[\begin{array}{l}
0 \\
1 \\
0
\end{array}\right] \text { and } \boldsymbol{z}_{2}=\boldsymbol{z}_{f c}=\left[\begin{array}{l}
0 \\
1 \\
0 \\
0 \\
0
\end{array}\right]
$$

Expected utility obtained from Model 1 would be expressed as:

$$
E\left(U_{f c}\right)=\alpha \text { Price }+\boldsymbol{\mu}_{\eta}\left[\begin{array}{l}
0  \tag{11}\\
1 \\
0 \\
0 \\
0
\end{array}\right]+\boldsymbol{\delta}\left[\begin{array}{l}
0 \\
1 \\
0
\end{array}\right] \text { Tail, }
$$

and the expected utility obtained from Model 2 would be expressed as:

$$
E\left(U_{n, f c}\right)=\alpha \text { Price }+\boldsymbol{\mu}_{\eta}\left[\begin{array}{l}
0  \tag{12}\\
1 \\
0 \\
0 \\
0
\end{array}\right]+\boldsymbol{\delta}\left[\begin{array}{l}
0 \\
1 \\
0
\end{array}\right] \text { Tail }+\left[\begin{array}{l}
0 \\
1 \\
0 \\
0 \\
0
\end{array}\right]\left[\boldsymbol{\gamma}_{1} \text { Income }_{n}+\boldsymbol{\gamma}_{2} \text { Age }_{n}+\right.
$$

$$
\left.\boldsymbol{\gamma}_{3} \text { Single }_{n}+\boldsymbol{\gamma}_{4} D N W A_{n}\right]
$$

where $\boldsymbol{\delta}$, the $\boldsymbol{\gamma}$ s and $\boldsymbol{\mu}_{\eta}$ are the coefficient vectors, and $\alpha$ is the price coefficient. If we omit the vector notation in equation (11), i.e.: extract the coefficients for farmed cod in the vectors
$\boldsymbol{\mu}_{\eta}$ and $\boldsymbol{\delta}$, and setting $\boldsymbol{z}_{2}=\boldsymbol{z}_{f c}=\left[\begin{array}{l}0 \\ 1 \\ 0 \\ 0 \\ 0\end{array}\right]=1$ and $\boldsymbol{y}_{2}=\boldsymbol{y}_{f c}=\left[\begin{array}{l}0 \\ 1 \\ 0\end{array}\right]=1$, Model $l$ can be rewritten as:

$$
\begin{equation*}
E\left(U_{f c}\right)=\alpha \text { Price }+\mu_{\eta, f c}+\delta_{f c} \text { Tail. } \tag{13}
\end{equation*}
$$

By omitting the vector notation in equation (12), Model 2 can be rewritten as:

$$
\begin{align*}
& \quad E\left(U_{n, f c}\right)=\alpha \text { Price }+\mu_{\eta, f c}+\delta_{f c} \text { Tail }+\left[\gamma_{1 f c} \text { Income }_{n}+\gamma_{2 f c} D N W A_{n}+\right.  \tag{14}\\
& \left.\gamma_{3 f c} \text { Age }_{n}+\gamma_{4 f c} \text { Single }_{n}\right]
\end{align*}
$$

From the output of Model 1 in the article (Table 6.1) we obtain the following results:

$$
\begin{equation*}
E\left(U_{f c}\right)=-.219[\text { Price }]+2.36-1.434[\text { Tail }] \tag{15}
\end{equation*}
$$

This is the logarithm of the odds of choosing farmed cod, and is to be interpreted as utility. The standard deviation for the random parameter is 2.229 and it is significant at the $1 \%$ level. This suggests preference heterogeneity for farmed cod. Figure A.4.1 shows the estimated distribution for the Farmed_Cod parameter. It is assumed that the fillet is a loin $($ Tail $=0)$.


Figure A.4.1. The Random Parameter Distribution for Farmed_Cod in Model 1.

Following Hole (2007b), we can estimate the percentage of the population that is positive to farmed cod by the formula:

$$
\begin{equation*}
100 * \phi\left(-\frac{\mu_{n i}}{\sigma_{i}}\right) \tag{16}
\end{equation*}
$$

The estimated proportion of those who are positive to farmed cod, given that the fillet is a loin, is:

$$
\begin{equation*}
100^{*} \phi\left(-\frac{2.36}{2.229}\right)=85.5 \% \tag{17}
\end{equation*}
$$

which is the area to the right of the vertical line in Figure A.4.1. The estimated proportion being negative to farmed cod is $100 \%(1-0.855)=14.5 \%$. Note that by "positive to farmed cod," we mean that an individual would rather choose farmed cod than the NOT alternative. The distribution displayed in Figure A.4.1 is the same distribution presented for Farmed_Cod in the article, but in the article the numbers on the horizontal axis are converted to WTP values, while they are expressed as utility in the appendix.

From the output of Model 2 in the article we obtain the following results:

$$
\begin{gather*}
E\left(U_{n, f c}\right)=-.22[\text { Price }]+.93-1.449[\text { Tail }]+.86\left[\text { Income }_{n}\right]  \tag{18}\\
-.74\left[D N W A_{n}\right]+.023\left[\text { Age }_{n}\right]+.64\left[\text { Single }_{n}\right]
\end{gather*}
$$

The estimated utility a single 55 year old person with high income obtains from choosing a loin of farmed cod priced 9.83 EUR/KG is thus;

$$
-.22 * 9.83+.93 * 1-1.449 * 0+.86 * 1-.74 * 0+.023 * 55+.64 * 1=1.5324
$$

The information of interest is the sign (positive or negative), the significance level ( $p$-value) and the relative magnitude between the parameter estimates.

## A.5. Generalized Extreme Value Distribution

The error term $\varepsilon_{n i s}$ in the random utility function is assumed to be independent and identically distributed (iid) extreme value. For a thorough review of extreme value distributions, see Coles (2011) and Train (2009, p. 76-96).

## A.6. The Variables

The variables of interest in the RC experiment were the fish attributes and the consumer characteristics. The attributes of the five fish types were; the cut (tail or loin) and the price. Salmon, farmed cod and wild cod have loin as well as tail cuts. Monk and pangasius did not have tail cuts.

The consumer characteristics of interest were income, gender, age, marital status, children in the household, and education. Model 1 in the article focuses on fish attributes only. It is meant to examine what attributes that affect people's preferences when no consumer characteristics have been controlled for. The fish attributes are included in all models. Model 2 in the article
includes the consumer characteristics; income, age and marital status. Section A. 7 explains how the consumer characteristics included in Model 2 were chosen.

Gender, presence of children in the household and level of education, which were the other consumer characteristics of interest, were not analyzed in the article. Model 3 includes gender, presence of children in the household, and level of education in addition to the consumer characteristics from Model 2. This model is not presented in the article. See section A. 9 for an analysis of Model 3.

## A.6.1. Specification of the Consumer Characteristics

The consumer characteristics of interest were included by the following variables, that each was interacted with the variables representing the fish types:

## Income:

Income was included by a dummy variable called Income, taking the value 1 if the participant's household's gross monthly income was above 3000 EUR, and 0 otherwise. 43 out of the 178 participants ( $24 \%$ ) belonged to this group. Section A.6.1.1 explains why Income was chosen as a dummy variable, and not as a continuous variable.

Age:

Age was included as a continuous variable, called Age. The age of the participants ranged from 21 to 70 years.

## Single:

Marital status was included as a dummy variable, Single, taking the value 1 is the participant was living in a single household with or without children, and 0 otherwise. Of the 178 participants in the experiment, 51 (29\%) were single.

## Female:

Gender was included as a dummy variable, Female, taking the value 1 if a participant is female, and 0 otherwise. 103 out of the 178 (58\%) participants were women.

## Children:

Presence of children in the household was included by a dummy variable, Children, taking the value 1 if a participant had children under the age of 18 in the household, and 0 otherwise. 61 out of the 178 (34\%) participants had children in the household.

## Education:

The level of education was included by a dummy variable, Education, taking the value 1 if a participant had two or more years of higher education, and 0 otherwise. 66 out of the 178 participants (37\%) belonged to this group. An explanation of why Education was defined as a dummy variable rather than as a continuous variable is described in section A.6.1.2.

As explained in Section A.4, to capture the effects of the consumer characteristics, interaction terms with the $\boldsymbol{z}_{i}$-vectors were necessary. By including one extra consumer characteristic variable, the mixed logit model estimates five more parameters (since there were five fish types). Adding many consumer characteristics would quickly consume many degrees of freedom, and potentially reduce the robustness of the model. There are many observations in the dataset, since each participant made 16 choices in which there were 4 choice alternatives, but the number of participants (178 in total) is relatively small. Hence, it is desirable to limit the number of consumer characteristics to reduce the number of coefficients.

Not all variables were of a desired format. Particularly two variables of interest caused problems; Income and Education. The next two sections explain why Income and Education were defined as dummy variables rather than as continuous variables. In section A. 7 we explain why Income, Age and Single were the socio-demographic variables included in Model 2 presented in the article, in addition to Model 1.

## A.6.1.1. The Income Variable

A household's gross monthly income had four categories in the survey questionnaire (Appendix II, p. 9):

Table A.6.1.1.1. Income Groups of the Participants

| Income group | Definition | \# of Participants | Share |
| :---: | :--- | :---: | :---: |
| 1 | Less than 2000 EUR a month | 49 | $27.53 \%$ |
| 2 | From 2000 to 3000 EUR a month | 67 | $37.64 \%$ |
| 3 | Over 3000 EUR a month | 43 | $24.16 \%$ |
| 4 | Do not know/Do not want to answer | 19 | $10.67 \%$ |

As few as three income groups and one do-not-want-to-answer (DNWA) group, made this variable inappropriate as a continuous variable. Income group 3 could include all values above 3000 EUR a month. Large income variation within this group is a reasonable assumption. Participants in the $D N W A$ group could belong to any level on the income spectrum, which made the income variable even harder to interpret. To get a meaningful variable for income, there were two options; (1) to include a dummy variable for each income group, or (2) to divide income into two groups; low and high. The former option would consume five more degrees of freedom than the latter, and potentially reduce the robustness of the model. As a first step to decide which options to choose, we estimated a model that included a dummy variable for each income group (Model 4 - Income as Dummies for all Income Groups). That is, income group 1 was defined as the base category, and income groups 2, 3 and $D N W A$ were defined by a dummy variable each. No other consumer characteristics but income were included in Model 4. To test the overall significance of the different income groups, the following hypotheses were postulated:

[^5]
## Hypothesis 2

Salmon * IncGr 3
$=$ Farmed Cod $* \operatorname{IncGr} 3$
$=$ Wild Cod $*$ IncGr 3
$=$ Monk $*$ incGr 3
$=$ Pangasius $*$ IncGr 3
$=0$

## Hypothesis 3

Salmon * DNWA
$=$ Farmed Cod $*$ DNWA
$=$ Wild Cod $*$ DNWA
$=$ Monk $* D N W A$
$=$ Pangasius $*$ DNWA
$=0$

The Wald test results are:

Table A.6.1.1.2. Wald Tests for the Income Groups

| Hypothesis | Consumer Interaction Variable | Chi-Sq | Df | $p$-value |
| :---: | :--- | :---: | :---: | :---: |
| 1 | IncGr2 | 8.19 | 5 | 0.1461 |
| 2 | IncGr3 | 9.08 | 5 | 0.1060 |
| 3 | DNWA | 11.94 | 5 | 0.0356 |

The difference between income group 1 and 2 is insignificant at the $10 \%$ level ( $p$-value 0.146 ). The difference between group 1 and 3 is almost significant at the $10 \%$ level ( $p$-value 0.106 ). Since the income group $D N W A$ can contain both low and high income levels, its test result is of limited interest. It was included for the sake of clarity and tidiness.

The next step was to estimate a model that divided income into low and high by including a dummy variable for income group 3 only (Model 5 Income as Low and High). This dummy variable was called Income. A dummy variable for $D N W A$ was included to avoid income bias (see below in this section); however, its test result is not of importance. To test if there is a significant difference between the low and the high income groups, the following hypotheses were postulated:

## Hypothesis 1

Salmon * Income
$=$ Farmed Cod $*$ Income
$=$ Wild Cod $*$ Income
$=$ Monk $*$ income
$=$ Pangasius $*$ Income
$=0$

## Hypothesis 2

Salmon * DNWA
$=$ Farmed Cod $*$ DNWA
$=$ Wild Cod $*$ DNWA
$=$ Monk $*$ DNWA
$=$ Pangasius $* D N W A$
$=0$

The Wald test results are:

Table A.6.1.1.3. Wald Tests for Model 5- Income as High or Low

| Hypothesis | Consumer Interaction Variable | Chi-Sq | Df | $p$-value |
| :---: | :--- | :---: | :---: | :---: |
| 1 | Income (Above 3000 EUR a month) | 13.75 | 5 | 0.0173 |
| 2 | DNWA (Do not want to answer) | 9.45 | 5 | 0.0925 |

The variable Income is significant at the $5 \%$ level ( $p$-value 0.0173 ). The model with Income only, (Model 5 Income as Low and High) is a special case of the more complex model that includes a dummy variable for each income group (Model 4 - Income as Dummies for all Income Groups). To examine whether model Model 5 had a significantly poorer fit than Model 4, a likelihood ratio test between the models was conducted. The test statistic is given by:

$$
-2\left[\text { Log likelihood } \text { Model4 }- \text { Log likelihood }_{\text {Model5 }}\right]
$$

For large samples the test statistic has an approximate chi-squared distribution, with $d f$ equal to the difference in numbers of parameters between the two models (Agresti 2007, p. 86). The null hypothesis is: There is no significant difference between the models in explaining the data. The likelihood ratio test result is:

Table A.6.1.1.4. Likelihood Ratio Test between Model 4 and Model 5

|  | Log- <br> likelihood | \# of <br> coef. | AIC | Df | Chi-Sq <br> Statistic | $p-$ <br> value |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Model 4 - Income as Dummies (2,3,4) | -2517.5306 | 39 | 5113.061 |  |  |  |
| Model 5 - Income as Low and High | -2521.5856 | 34 | 5111.171 | 5 | 8.11 | 0.1502 |

The $p$-value of 0.1502 indicates that there is no significant improvement in model fit, by having each income group as a dummy variable of its own. Table A.6.1.1.4 also includes the AIC of the two models. AIC is a measure of model fit that penalizes a model for having many parameters (Agresti 2007, p. 141). A lower AIC number indicates a better fit. AIC is given by:

$$
\text { AIC }=-2[\text { Log likelihood }- \text { number of parameters in model }]
$$

Model 4 has the lower AIC of the two models, indicating a better fit. Thus we concluded to continue the analysis by having income as a dummy variable. The fact that income group 3 can include all values above 3000 EUR, also makes it natural to divide income into low and high.

The dummy variable for the income group $D N W A$ was included in both models (and in all other models where Income is included). Its estimated parameters and test statistics provide limited information since individuals from all income levels can belong to this group. A nonnegligible fraction of the participants ( $10.37 \%$ ) belonged to this group. Omitting this variable would reduce the dataset somewhat. It is impossible to know what type of people who did not want to reveal their income are. For all we know, this could be people with strong preferences for monk or pangasius, or something else. Thus leaving them out of the analysis could potentially cause a bias. Another option was to simply place the whole group into one of the other income groups, by assuming they would either have low or high income. This could also cause a potential bias. Assume they were to be placed in the high income group. The high income group would then consist of the people who had high income as well as everyone from the $D N W A$ group. If many respondents in the $D N W A$ group actually had low income, the Income variable (which is a dummy variable for high income) would be very imprecise. If many people from the $D N W A$ group had strong preferences for pangasius, it would seem like people with high income had strong preferences for pangasius. Hence, an income bias towards pangasius.

The $D N W A$ estimates are omitted from the outputs presented in the article, since their parameter estimates were of no importance for our analysis. If the parameter estimates of the $D N W A$ variable are of interest to the reader, see the regression outputs in section A. 12 .

## A.6.1.2. The Education Variable

The variable for education consisted of seven levels of dubious ordinal and cardinal order (Appendix II, p. 1). Therefore we found it inappropriate to use it as a continuous variable.

Table A.6.1.2.1. The Level of Education Among the Participants

| Education Level | \# of Participants | Share | Education <br> Dummy |
| :--- | :---: | :---: | :---: |
| No diploma | 7 | $3.93 \%$ | 0 |
| Brevet des colleges | 21 | $11.80 \%$ | 0 |
| CAP ou BEP | 36 | $20.22 \%$ | 0 |
| Baccaulaureat (BAC) | 48 | $26.97 \%$ | 0 |
| BAC +2 or 3 | 41 | $23.03 \%$ | 1 |
| BAC +3 or 4 | 22 | $12.36 \%$ | 1 |
| BAC +6 | 3 | $1.69 \%$ | 1 |

Having a dummy variable for each of the seven education levels (six dummy variables in addition to the base category) would consume many degrees of freedom, since one would have to estimate (\# of dummy variables) * (\# of fish types) $=6 * 5=30$ extra parameters. We therefore defined education as low and high, by creating a dummy variable, Education. This variable takes the value 1 if the participant has at least 2 years of education after completing BAC, and 0 otherwise. (That is, Education $=1$ if a participant has "BAC +2 or 3 " or "BAC +3 or 4 " or "BAC +6 ".) A total of 66 participants ( $37 \%$ ) belonged to this group. Education is expected to be correlated with income. $79.46 \%$ of the participants who belonged to the low education group (i.e.: lower than $\mathrm{BAC}+2$ or 3 ) also belonged to the low income group (i.e.: income less than 3000 EUR a month). Including both variables might not be necessary. To test for this we estimated a model that included both Income and Education, and postulated the following hypotheses:

[^6]
## Hypothesis 2

Salmon * DNWA
$=$ Farmed Cod $*$ DNWA
$=$ Wild Cod $*$ DNWA
$=$ Monk $*$ DNWA
= Pangasius $*$ DNWA
$=0$

## Hypothesis 3

Salmon * Education
$=$ Farmed Cod $*$ Education
$=$ Wild Cod $*$ Education
$=$ Monk $*$ Education
$=$ Pangasius $*$ Education
$=0$

The Wald test results are:

Table A.6.1.2.2. Wald Test for Income and Education

| Test | Consumer Characteristics Interaction | Chi-Sq | Df | $p$-value |
| :---: | :--- | :---: | :---: | :---: |
| 1 | Income | 14.15 | 5 | 0.0147 |
| 2 | DNWA | 10.78 | 5 | 0.056 |
| 3 | Education | 6.52 | 5 | 0.2590 |

The $p$-value of 0.2590 suggests that Education does not have a significant influence on the preferences for fish, when Income is controlled for. A likelihood ratio test to examine whether the more complex model, which includes both Income and Education, explains the data significantly better than its in-nested model, which includes only Income, gave the following result:

Table A.6.1.2.3. Likelihood Ratio Test between Model 5 and Model 6

| Model | Log- <br> likelihood | $\#$ of <br> coef. | AIC | Df | Chi-Sq. <br> Statistic | p-value |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Model 6- Income and Education | -2518.3447 | 39 | 5114.6894 |  |  |  |
| Model 5- Income | -2521.5856 | 34 | 5111.1712 | 5 | 6.4818 | 0.2621 |

The null hypothesis of no difference between the models is retained. Model 6 , which includes both Income and Education, does not describe the data significantly better than Model 5 that only includes Income. The AIC is lower for Model 5 than Model 6, suggesting a better fit.

The variable Education shows no significant effect on the preferences for fish, and it is likely to be collinear with income. It was of dubious ordinal and cardinal order, and had many levels, which made it inconvenient to divide it into separate dummy variables. On the basis of this, and the elimination procedure to be explained in the next section (A.7), Education was omitted as a variable in the article.

## A.7. Choice of Consumer Characteristics

Having defined the variables Income and Education in the previous sections (A.6.1.1 and A.6.1.2) we wanted to find the simplest model that best fits the data. That is, a model that would not have a significantly poorer fit than a larger model, but at the same time get a significantly poorer fit if a variable were to be removed. To find this model the following procedure was used: (1) Estimate a model with all relevant variables of interest, interacted with the variables representing the fish types, the $\boldsymbol{z}_{i}$-vectors. (2) Find the most insignificant variable interactions with a Wald-test. (3) Estimate a new model that excludes the most insignificant consumer characteristics interactions found in step (2). (4) Run a likelihood ratio test between the two models. (5) If there is no significant difference between the models, keep the simpler model, and repeat the procedure.

The procedure was repeated until we were left with a model that had a significantly better fit than its nested model and that did not have a significantly poorer fit than the model it was nested in. This procedure is quite similar to the technique of backward elimination (see Agresti (2007, p. 141)). The result of this procedure is found in Table A.7.1. Even though Education already has proven to be a poor explanatory variable, we included it in the procedure. Not surprisingly it was the first variable to be omitted.

Table A.7.1. Likelihood Ratio Tests to Find the Variables that Best Describe the Data

| Model | Consumer Characteristics | Variable with Highest Wald $p$ value | $\begin{gathered} \text { Log- } \\ \text { likelihood } \end{gathered}$ | \# of <br> Coef. | AIC | Models <br> Compared | Df | Chi-Sq <br> Statistic | $\begin{gathered} p- \\ \text { value } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Model 3 | Income, Age, <br> Single, Female, Children, Education | Education (0.9144) | -2505 | 59 | 5129 |  |  |  |  |
| Model 9 | Income, Age, Single, Female, Children | Children (0.8437) | -2506 | 54 | 5120 | 3-9 | 5 | 0.839 | 0.9745 |
| Model 8 | Income, Age, Single, Female | Female (0.5586) | -2507 | 49 | 5112 | 9-8 | 5 | 2.016 | 0.8469 |
| Model 2 | Income, Age, <br> Single | Single (0.0499) | -2509 | 44 | 5106 | 8-2 | 5 | 3.893 | 0.5649 |
| Model 7 | Income, Age |  | -2518 | 39 | 5114 | 2-7 | 5 | 18.264 | 0.0026 |

Children was the second variable to be removed. The likelihood ratio test between Model 9 and Model 8 had a $p$-value of 0.8469 . A model that includes the variable Children, does not explain the data significantly better than a model without it. Having children in the household is an ambiguous measure. As an example; there is a big difference between having a three year old girl and two teenage boys in the household. It is likely that the purchase pattern of the former family constellation is quite different from that of the latter. Purchase of fish is no exception. Hence, Children's limited impact on the preferences for fish comes as no surprise.

The next variable to be removed was Female. The likelihood ratio test between Model 8 and Model 2 gave a $p$-value of 0.5649 . Including Female does not significantly describe the data better, than omitting it.

After Female had been removed, Single was the variable with the highest $p$-value from the Wald-test $(p$-value $=0.0499)$. The likelihood ratio test showed that removing Single from the model would lead to a significantly poorer fit $(p$-value $=0.0026)$.

The variables left to be presented in the article were thus Income, Age and Single. Even if Education, Children and Female were omitted from the models presented in the article, some interesting findings were made. Neither higher education, the presence of children in the household or gender significantly affects the WTP for fish.

## A.8. Correlations in Preferences over Choice Scenarios

All models were estimated both with and without allowing for correlations in unobserved factors over choice scenarios. To test which of the two methods that gave the best fit, each model pair (one with and the other without correlations) was tested against each other with a likelihood ratio test. Since there are 10 coefficients on the off-diagonal part of the covariance matrix, the difference in estimated coefficients is 10 for each model pair. Hence, $d f=10$. The test results can be found below.

Table A.8.1. Likelihood Ratio Tests Between Models that Allow for Correlations in Unobserved Factors over Choice Scenarios and Models that do not Allow for Correlations in Unobserved Factors over Choice Scenarios

| Consumer Characteristics | Models <br> Compared | Log <br> Likelihood <br> Non-Corr | Log <br> Likelihood <br> Corr | Chi-Sq <br> Statistic | Df | $p$-value |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Fish Attributes Only | B. $1-1$ | -2620.026 | -2534.497 | 171.0580 | 10 | $1.68 \mathrm{E}-31$ |
| Income, Age, Single | B. $2-2$ | -2584.737 | -2508.769 | 151.9360 | 10 | $1.49 \mathrm{E}-27$ |
| Income, Age, Single, Female, | B.3-3 | -2578.085 | -2505.395 | 145.3800 | 10 | $3.32 \mathrm{E}-26$ |
| Children, Education <br> Income as dummies | B. $4-4$ | -2598.592 | -2517.531 | 162.1220 | 10 | $1.18 \mathrm{E}-29$ |
| (2, 3 and 4) | B.5-5 | -2600.926 | -2521.586 | 158.6800 | 10 | $6.07 \mathrm{E}-29$ |
| Income | B. $6-6$ | -2598.7433 | -2518.3447 | 160.7972 | 10 | $2.22 \mathrm{E}-29$ |
| Income, Education | B. $7-7$ | -2591.82 | -2517.901 | 147.8380 | 10 | $1.04 \mathrm{E}-26$ |
| Income, Age | B. $8-8$ | -2582.376 | -2506.822 | 151.1080 | 10 | $2.21 \mathrm{E}-27$ |
| Income, Age, Single, Female | B. $9-9$ | -2581.152 | -2505.814 | 150.6760 | 10 | $2.71 \mathrm{E}-27$ |
| Income, Age, Single, Female, |  |  |  |  |  |  |

Every single model that allowed for correlations in unobserved factors over choice scenarios explained the data significantly better than those that did not allow for correlations in unobserved factors over choice scenarios, even at the $1 \%$ level. The two models presented in the article allow for correlations over choice scenarios, providing more information about the structure of the data.

## A.9. Model 3 - Product Attribute and Full Set of Consumer Characteristics Interaction

Model

The output for Model 3 includes Female, Children and Education in addition to the consumer characteristics included in Model 2.

Table A.9.1. Empirical Estimates for Model 3 - Product Attribute and Full Set of Consumer Characteristics Model

| Variable Name | Estimate |  | Std.Err |
| :---: | :---: | :---: | :---: |
| Price | -0.219 | *** | (0.010) |
| Salmon | 3.335 | *** | (0.752) |
| Farmed_Cod | 1.144 |  | (1.113) |
| Wild_Cod | 1.532 |  | (0.952) |
| Monk | 1.428 |  | (1.408) |
| Pangasius | -1.663 |  | (1.757) |
| Tail * Salmon | -0.429 | *** | (0.118) |
| Tail * Farmed_Cod | -1.440 | *** | (0.193) |
| Tail * Wild_Cod | -0.661 | *** | (0.152) |
| Income * Salmon | 0.487 |  | (0.389) |
| Income * Farmed_Cod | 0.815 |  | (0.527) |
| Income * Wild_Cod | 1.093 | ** | (0.442) |
| Income * Monk | 0.094 |  | (0.653) |
| Income * Pangasius | 0.592 |  | (0.868) |
| DNWA * Salmon | -0.977 | ** | (0.483) |
| DNWA * Farmed_Cod | -0.713 |  | (0.704) |
| DNWA * Wild_Cod | -0.883 |  | (0.568) |
| DNWA * Monk | -2.414 | ** | (0.956) |
| DNWA * Pangasius | -0.868 |  | (1.013) |
| Age * Salmon | -0.017 |  | (0.013) |
| Age * Farmed_Cod | 0.024 |  | (0.018) |
| Age * Wild_Cod | 0.027 | * | (0.016) |
| Age * Monk | 0.037 |  | (0.023) |
| Age * Pangasius | -0.003 |  | (0.028) |
| Single * Salmon | 0.911 | ** | (0.362) |
| Single * Farmed_Cod | 0.514 |  | (0.483) |
| Single * Wild_Cod | 0.382 |  | (0.431) |
| Single * Monk | -0.073 |  | (0.635) |
| Single * Pangasius | 0.819 |  | (0.727) |
| Female * Salmon | 0.187 |  | (0.290) |
| Female * Farmed_Cod | -0.514 |  | (0.370) |
| Female * Wild_Cod | -0.292 |  | (0.320) |
| Female * Monk | -0.273 |  | (0.473) |
| Female * Pangasius | -0.178 |  | (0.601) |
| Children * Salmon | -0.042 |  | (0.322) |


| Children * Farmed_Cod | -0.032 | $(0.528)$ |  |
| :--- | ---: | :--- | ---: |
| Children * Wild_Cod | -0.457 | $(0.441)$ |  |
| Children * Monk | -0.511 | $(0.632)$ |  |
| Children * Pangasius | 0.038 | $(0.661)$ |  |
| High_Educ * Salmon | 0.241 | $(0.349)$ |  |
| High_Educ * Farmed_Cod | 0.287 | $(0.430)$ |  |
| High_Educ * Wild_Cod | 0.366 | $(0.351)$ |  |
| High_Educ * Monk | 0.356 | $(0.552)$ |  |
| High_Educ * Pangasius | 0.481 | $(0.783)$ |  |
| Standard Deviations |  |  |  |
| Salmon | 1.645 | $* * *$ | $(0.138)$ |
| Farmed_Cod | 2.134 | $* * *$ | $(0.231)$ |
| Wild_Cod | 1.872 | $* * *$ | $(0.189)$ |
| Monk | 2.448 | $* * *$ | $(0.328)$ |
| Pangasius | 2.967 | $* * *$ | $(0.403)$ |
|  |  |  |  |
| N |  |  |  |
| LR Chi-Squared |  | -2505.3946 |  |
| Log-likelihood |  | 5128.7892 |  |
| AIC |  |  | $\alpha=0.1 *$ |
| Significance codes: |  |  |  |

$57 \%$ of the participants were female and $34 \%$ of the participants had children less than 18 years of age in the household. $37 \%$ of the participants had higher education. Neither by t-tests, Wald-tests nor Likelihood Ratio tests do these variables have any significant impact on the utility obtained from choosing the different fish types. Women's utility obtained from choosing a given type of fish does not differ from that of men. Presence of children in the household does not affect the utility obtained from choosing the different fish types. Hence, farmed cod can appeal equally to men and woman, to families with and without children, and to individuals with or without higher education. Some of the DNWA (Do not want to answer) coefficient estimates are significant. The only information they provide is that people that did not want to reveal their income had significant negative parameter estimates for Salmon and Monk.

## A.10. Chi-squared Tests between Parameter Estimates

For both Model 1 and Model 2 we conducted several tests. The test results can be found in the matrices below. The chi-squared statistics and the corresponding $p$-value can be found for every test conducted. The null hypothesis in each test is that the sum of the coefficient for a certain type of fish and one or more consumer characteristic interactions with the same fish type is equal to a similar expression, but for another type of fish. For example, to test whether the coefficients for salmon and farmed cod in Model 1 are significantly different from each other the null hypothesis is:

$$
H_{0}: \mu_{\eta, s a}=\mu_{\eta, f c}
$$

The test result is found as the top left result in Table A.10.1, i.e., the chi-squared statistic is 5.69 and the corresponding $p$-value is 0.0171 . Hence we reject the null hypothesis that the coefficients are equal.

As another example; at the top left test result of Table A.10.5, the null hypothesis is:

$$
H_{0}: \mu_{\eta, s a}+\gamma_{\text {income }} * \mu_{\eta, s a}=\mu_{f c}+\gamma_{\text {income }} * \mu_{\eta, f c}
$$

The chi-squared statistic is 8.03 and the corresponding $p$-value is 0.005 . Hence the null hypothesis is rejected.

In the matrices below are all the Chi-Squared tests we conducted. The header of each table says which consumer characteristic interaction coefficients, in addition to the coefficients for the fish types, which were tested against each other (except Table A.10.2). For example, in Table A. 10.6 we test whether the sum of the coefficients for a fish type plus the coefficient for the interaction between Single and the same type of fish is equal to a similar expression for another fish type.

Table A.10.1. Chi-Squared Tests for Model 1

|  |  | Salmon | Farmed Cod | Wild Cod | Monk |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 5.69 |  |  |  |
|  | $p$-value | 0.0171 |  |  |  |
| Wild Cod | Chi-Sq. | 0.1 | 4.55 |  |  |
|  | $p$-value | 0.7503 | 0.0329 |  |  |
|  | Chi-Sq. | 0.37 | 1.58 | 0.19 |  |
|  | $p$-value | 0.5434 | 0.2081 | 0.6657 |  |
|  | Chi-Sq. | 75.06 | 46.43 | 62.7 | 51.51 |
|  | $p$-value | 0.0000 | 0.0000 | 0.0000 | 0.0000 |

Table A.10.2. Chi-Squared Tests for Model 1 The Tail Coefficients only

|  |  | Salmon | Farmed Cod |
| :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 51.51 |  |
|  | $p$-value | 0.0000 |  |
| Wild Cod | Chi-Sq. | 1.76 | 11.5 |
|  | $p$-value | 0.1848 | 0.0007 |

Table A.10.3. Chi-Squared Tests for Model 1 The Tail + Fish Coefficients

|  |  | Salmon | Farmed Cod |
| :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 41.69 |  |
|  | $p$-value | 0.0000 |  |
| Wild Cod | Chi-Sq. | 2.2 | 25.59 |
|  | $p$-value | 0.1384 | 0.0000 |

All the below matrices are Chi-Squared tests conducted for Model 2.

Table A.10.4. Model 2 Chi-Squared Tests for the Random Parameters only

|  |  | Salmon | Farmed Cod | Wild Cod | Monk |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 12.39 |  |  |  |
|  | $p$-value | 0.0000 |  |  |  |
| Wild Cod | Chi-Sq. | 13.93 | 0.07 |  |  |
|  | $p$-value | 0.0000 | 0.7930 |  |  |
|  | Chi-Sq. | 6.29 | 0.01 | 0.012 |  |
|  | $p$-value | 0.0010 | 0.9070 | 0.9350 |  |
|  | Chi-Sq. | 18.21 | 3.14 | 3.87 | 2.9 |
|  | $p$-value | 0.0000 | 0.0760 | 0.0490 | 0.0870 |

Table A.10.5. Model 2 Chi-Squared Tests Consumer Characteristics Interaction: Income

|  |  | Salmon | Farmed Cod | Wild Cod | Monk |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 8.03 |  |  |  |
|  | p-value | 0.005 |  |  |  |
| Wild Cod | Chi-Sq. | 6.42 | 0.55 |  |  |
|  | $p$-value | 0.011 | 0.457 |  |  |
|  | Chi-Sq. | 6.11 | 0.15 | 1.14 |  |
| Monk | pangasius | -value | 0.013 | 0.695 | 0.286 |
|  | Chi-Sq. | 17.47 | 3.32 | 5.62 | 1.75 |
|  | $p$-value | 0.0000 | 0.0690 | 0.0180 | 0.1860 |

Table A.10.6. Model 2 Chi-Squared Tests Consumer Characteristics Interaction: Single

|  |  | Salmon | Farmed Cod | Wild Cod | Monk |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 16.55 |  |  |  |
|  | $p$-value | 0.000 |  |  |  |
| Wild Cod | Chi-Sq. | 17.97 | 0.11 |  |  |
|  | $p$-value | 0.000 | 0.7403 |  |  |
|  | Chi-Sq. | 10.01 | 0.08 | 0.35 |  |
|  | $p$-value | 0.0016 | 0.7822 | 0.5523 |  |
|  | Chi-Sq. | 16.5 | 2.01 | 2.51 | 1.13 |
|  | $p$-value | 0.000 | 0.1563 | 0.1135 | 0.287 |

Table A.10.7. Model 2 Chi-Squared Tests Consumer Characteristics Interaction: Age

|  |  | Salmon | Farmed Cod | Wild Cod | Monk |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 12.39 |  |  |  |
|  | p-value | 0.0004 |  |  |  |
| Monk | Chi-Sq. | 13.85 | 0.08 |  |  |
|  | p-value | 0.0002 | 0.783 | 0.10 |  |
|  | Chi-Sq. | 6.22 | 0.02 | 0.9421 |  |
|  | pangasius | Chi-Sq. | 18.72 | 0.894 | 4.12 |
| 3 | p-value | 0.0000 | 3.33 | 0.11 |  |
|  |  |  | 0.0681 | 0.0424 | 0.0779 |

Table A.10.8. Model 2 Chi-Squared Tests Consumer Characteristics Interaction: Income and Age

|  |  | Salmon | Farmed Cod | Wild Cod | Monk |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 7.97 |  |  |  |
|  | $p$-value | 0.0047 |  |  |  |
| Wild Cod | Chi-Sq. | 6.27 | 0.58 |  |  |
|  | $p$-value | 0.0123 | 0.4451 |  |  |
|  | Chi-Sq. | 6.04 | 0.15 | 1.15 |  |
|  | $p$-value | 0.014 | 0.701 | 0.2833 |  |
|  | Chi-Sq. | 17.95 | 3.52 | 5.96 | 1.9 |
|  | $p$-value | 0.0000 | 0.0607 | 0.0146 | 0.1680 |

Table A.10.9. Model 2 Chi-Squared Tests Consumer Characteristics Interaction: Income and Single

|  |  | Salmon | Farmed Cod | Wild Cod | Monk |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 9.02 |  |  |  |
|  | $p$-value | 0.0027 |  |  |  |
| Wild Cod | Chi-Sq. | 7.25 | 0.57 |  |  |
|  | $p$-value | 0.0071 | 0.4494 |  |  |
| Monk | Chi-Sq. | 8.06 | 0.51 | 2.06 |  |
|  | $p$-value | 0.0045 | 0.4763 | 0.1513 |  |
|  | Chi-Sq. | 13.52 | 1.75 | 3.23 | 0.38 |
|  | $p$-value | 0.0002 | 0.1853 | 0.0724 | 0.5395 |

Table A.10.10. Model 2 Chi-Squared Tests Consumer Characteristics Interaction: Income, Age and Single

|  |  | Salmon | Farmed Cod | Wild Cod | Monk |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Farmed Cod | Chi-Sq. | 8.95 |  |  |  |
|  | $p$-value | 0.0028 |  |  |  |
| Wild Cod | Chi-Sq. | 7.09 | 0.6 |  |  |
|  | $p$-value | 0.0078 | 0.4389 |  |  |
|  | Chi-Sq. | 7.99 | 0.51 | 2.08 |  |
|  | $p$-value | 0.0047 | 0.4762 | 0.1488 |  |
|  | Chi-Sq. | 13.83 | 1.87 | 3.43 | 0.43 |
|  | $p$-value | 0.0002 | 0.1716 | 0.0642 | 0.514 |

## A.11. Survey Results - Definition of Agree and Disagree

The survey questions 2.1-2.8, 3.1-3.5, 6.1-6.8, 12.1-12.11, 13.1-13.11, 14.1-14.11, 15.1.15.11, and 17.1-17.10 are categorized from 1-10 or 1-10 in addition to 11, "Do not know".

As an example, question 17 is categorized like this:

## Attitudes toward fish farming and environmental aspects



Figure A.11.1. An Example of a Survey Question

We have defined "disagree" to be the values 1-4. We have defined "agree" to be the values 710. Whenever we write that a certain percentage of the respondents "agrees to the statement that..," we refer to the percentage of the respondents that have answered $7,8,9$ or 10 .
Equivalently, whenever we write that a certain percentage of the respondents "does not agree to the statement that..." we refer to the percentage of the respondents that answered $1,2,3$ or
4. We have followed this definition throughout the whole article. The definition is arbitrary. We chose it for convenience.

When we analyzed the survey results the cross section version of the dataset has been convenient (See section A.13.10). This implies that the dataset consists of one row for each participant, that is, 178 rows. The original dataset has 64 rows for each participant, since each participant made 16 choices, and there were 4 choice alternatives in each choice set ( 3 fillets of fish and NOT). Since consumer characteristics do not vary over choice scenarios, the cross section dataset has been suitable for the survey analysis.

The survey questions used in the article are summarized in the tables below. Each variable is tabulated in STATA 12, and from there the share (in percentage) of respondents in category 1-4 and 7-10 are summarized. Some questions are summarized conditional on income group or marital status.
Table A.11.1. Survey Responses on Fish Likings, Fish Buying Habits and Origin of the Fish

| R4_ |  | How often would you say you eat the following items for lunch or dinner at home? Check one box per line. | All respondents |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Twice a week or more \% | Once a week <br> \% | 2-3 times a month \% 3.93 | Once a month \% 3.93 | Every second month \% | 2-4 times a year \% | $\begin{gathered} \text { More seldom } \\ \% \end{gathered}$ | Never \% |  |  |  |
|  |  |  |  |  |  | 3.93 |  |  |  |  |  |  |  |
|  |  |  | All respondents |  |  | High income |  | Low income |  | Single |  | Married/Cohabiting |  |
|  |  | Fish likings | Like \% | Dislike \% | Do not know \% | Like \% | Dislike \% | Like \% | Dislike \% | Like \% | Dislike \% | Like \% | Dislike \% |
| q2_ | 1 | Salmon (non-smoked) | 78.65 | 10.11 |  |  |  |  |  |  |  |  |  |
| q2_ | 2 | Cod | 63.27 | 10.17 |  |  |  |  |  |  |  |  |  |
| q2 | 3 | Monk | 74.71 | 15.74 |  |  |  |  |  |  |  |  |  |
| q2 | 4 | Pangasius | 17.42 | 68.54 |  |  |  |  |  |  |  |  |  |
|  |  |  | All respondents |  |  | High income |  | Low income |  | Single |  | Married/Cohabiting |  |
|  |  | Buying fish | Agree \% | Disagree \% | Do not know \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% |
| q3_ | 1 | I have always decided which type of fish to buy before I go to the store | 24.15 | 58.42 |  |  |  |  |  |  |  |  |  |
| q3 | 2 | I prefer to buy pre-packed filets of fish | 25.83 | 48.33 |  |  |  |  |  |  |  |  |  |
| q3_ | 3 | I most offen choose the type of fish that is discounted | 56.74 | 21.35 |  |  |  |  |  |  |  |  |  |
| q3 | 4 | It is important to know where the fish has been caught/produced | 63.27 | 15.81 |  |  |  |  |  |  |  |  |  |
| q3 | 5 | I always ask how fresh the fish is before I make a decision | 48.87 | 32.03 |  |  |  |  |  |  |  |  |  |
|  |  |  | All respondents |  |  | High income |  | Low income |  | Single |  | Married/Cohabiting |  |
|  |  | Origin of the fish | Agree \% | Disagree \% | Do not know \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% |
|  |  | I have a very positive view of fresh farmed fish from: |  |  |  |  |  |  |  |  |  |  |  |
| q6_ | 1 | France | 75.57 | 5.68 |  |  |  |  |  |  |  |  |  |
| q6 | 2 | Countries in Northern Europe | 72.16 | 6.25 |  |  |  |  |  |  |  |  |  |
| q6_ | 3 | Countries in Southern Europe | 34.65 | 18.75 |  |  |  |  |  |  |  |  |  |
| q6 | 4 | Other developed countries | 14.45 | 23.7 |  |  |  |  |  |  |  |  |  |
| q6_ | 5 | Third world countries | 5.12 | 64.77 | 21.02 |  |  |  |  |  |  |  |  |
|  |  | I have a very positive view of fresh wild fish from: |  |  |  |  |  |  |  |  |  |  |  |
| q6 | 6 | The Atlantic North | 86.36 | 1.14 |  |  |  |  |  |  |  |  |  |
| q6_ | 7 | The Mediterranean | 63.07 | 3.98 |  |  |  |  |  |  |  |  |  |
| q6 | 8 | The Pacific | 54.54 | 10.79 | 12.5 |  |  |  |  |  |  |  |  |

Table A.11.2. Survey Responses on Attitudes Towards Fresh Salmon, Fresh Cod, Fresh Monk and Pangasius

|  |  |  | All respondents |  |  | High income |  | Low income |  | Single |  | Married/Cohabiting |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Attitudes towards fresh salmon | Agree \% | Disagree \% | Do not know \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% |
| q12_ | 1 | salmon tastes good. | 79.77 | 7.31 |  |  |  |  |  | 90.19 | 7.84 | 75.59 | 7.08 |
| q12 | 2 | salmon gives you good value for money | 68.54 | 8.43 |  |  |  |  |  | 70.58 | 9.8 | 67.72 | 7.87 |
| q12 | 3 | It is easy to make different dishes with salmon. | 84.83 | 6.74 |  |  |  |  |  | 84.32 | 3.92 | 85.05 | 7.87 |
| q12 | 4 | Salmon is healthy food | 83.52 | 1.71 |  |  |  |  |  |  |  |  |  |
| q12 | 5 | Salmon is fat food.. | 73.03 | 14.61 |  |  |  |  |  |  |  |  |  |
| q12_ | 6 | Salmon is safe to eat | 50.56 | 15.92 |  |  |  |  |  |  |  |  |  |
| q12 | 7 | Salmon is easy to prepare | 93.26 | 2.24 |  |  |  |  |  | 94.11 | 1.96 | 92.91 | 2.37 |
| q12 | 8 | Salmon is an expensive fish | 53.94 | 18.53 |  |  |  |  |  |  |  |  |  |
| q12 | 9 | The whole family likes salmon | 72 | 13.15 |  |  |  |  |  |  |  |  |  |
| q12 | 10 | Salmon can be served on special occasions | 75.27 | 17.41 |  |  |  |  |  |  |  |  |  |
| q12_ | 11 | Salmon is a Monday-to-Friday fish | 50.56 | 26.96 |  |  |  |  |  |  |  |  |  |
|  |  |  |  | All respond | ents | High | income | Low | income |  | ingle | Married/ | Cohabiting |
|  |  | Attitudes towards fresh cod | Agree \% | Disagree \% | Do not know \% | Agree \% | Disagree \% | Agree \% | Disagree\% | Agree \% | Disagree \% | Agree \% | Disagree \% |
| q13_ | 1 | Cod tastes good. | 67.79 | 10.73 |  | 76.73 | 9.31 | 64.34 | 11.31 | 62.75 | 11.76 | 69.85 | 10.31 |
| q13 | 2 | Cod gives you good value for money | 52.82 | 17.97 |  |  |  |  |  |  |  |  |  |
| q13_ | 3 | It is easy to make different dishes with cod. | 58.52 | 11.93 |  |  |  |  |  |  |  |  |  |
| q13 | 4 | Cod is healthy food | 74.16 | 0.56 |  | 88.38 | 0 | 68.11 | 0 |  |  |  |  |
| q13 | 5 | Cod is fat food.. | 11.85 | 51.41 |  |  |  |  |  |  |  |  |  |
| q13_ | 6 | Cod is safe to eat | 47.45 | 15.25 |  | 53.49 | 9.3 | 46.95 | 17.4 |  |  |  |  |
| q13_ | 7 | Cod is easy to prepare | 84.84 | 2.24 |  |  |  |  |  | 80.39 | 1.96 | 86.62 | 2.36 |
| q13_ | 8 | Cod is an expensive fish | 51.4 | 24.29 |  | 65.12 | 18.61 | 43.1 | 28.45 | 38 | 18 | 56.69 | 26.77 |
| q13_ | 9 | The whole family likes cod | 67.97 | 10.12 |  |  |  |  |  |  |  |  |  |
| q13_ | 10 | Cod can be served on special occasions | 26.86 | 49.72 |  |  |  |  |  |  |  |  |  |
| q13_ | 11 | Cod is a Monday-to-Friday fish | 71.19 | 16.94 |  |  |  |  |  |  |  |  |  |
|  |  |  | All res | pondents |  |  |  |  |  |  |  |  |  |
|  |  | Attitudes towards fresh monk | Agree \% | Disagree\% |  |  |  |  |  |  |  |  |  |
| q14_ | 7 | Monk is easy to prepare | 42.86 | 24.58 |  |  |  |  |  |  |  |  |  |
|  |  |  |  | All respond | ents | High | income | Low | income |  | ingle | Married/ | Cohabiting |
|  |  | Attitudes towards pangasius | Agree \% | Disagree \% | Do not know \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% |
| q15_ | 1 | Pangasius tastes good | 18.54 | 60.68 | 10.11 |  |  |  |  |  |  |  |  |
| q15_ | 2 | Pangasius gives you good value for money | 32.76 | 21.47 | 37.85 |  |  |  |  |  |  |  |  |
| q15_ | 3 | It is easy to make different dishes with Pangasius | 11.86 | 24.29 | 53.67 |  |  |  |  |  |  |  |  |
| q15 | 4 | Pangasius is healthy food | 18.63 | 23.72 | 47.46 |  |  |  |  |  |  |  |  |
| q15_ | 5 | Pangasius is fat food | 12.42 | 28.24 | 49.15 |  |  |  |  |  |  |  |  |
| q15_ | 6 | Pangasius is safe to eat | 10.22 | 32.39 | 49.43 |  |  |  |  |  |  |  |  |
| q15_ | 7 | Pangasius is easy to prepare | 32.59 | 11.24 | 47.75 |  |  |  |  |  |  |  |  |
| q15_ | 8 | Pangasius is an expensive fish | 10.22 | 42.61 | 38.07 |  |  |  |  |  |  |  |  |
| q15_ | 9 | The whole family likes Pangasius | 11.86 | 32.19 | 45.76 |  |  |  |  |  |  |  |  |
| q15_ | 10 | Pangasius can be served on special occasions | 12.42 | 48.58 | 31.07 |  |  |  |  |  |  |  |  |
| q15_ | 11 | Pangasius is a Monday-to-Friday fish | 30.34 | 28.65 | 32.02 |  |  |  |  |  |  |  |  |

Table A.11.3. Survey Responses on Attitudes Toward Fish Farming and Environmental Aspects

|  |  |  | All respondents |  |  | High income |  | Low income |  | Single |  | Married/Cohabiting |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Attitudes towards fish farming and environmental aspects | Agree \% | Disagree \% | Do not know \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% | Agree \% | Disagree \% |
| q17 | 1 | Farmed fish is healthy food | 30.9 | 30.9 |  |  |  |  |  |  |  |  |  |
| q17 | 2 | Wild fish is healthy food | 83.14 | 2.81 |  |  |  |  |  |  |  |  |  |
| q17_ | 3 | Farmed fish is safe to eat | 33.14 | 33.15 | 11.8 |  |  |  |  |  |  |  |  |
| q17_ | 4 | Wild fish is safe to eat | 60.45 | 10.72 |  |  |  |  |  |  |  |  |  |
| q17 | 5 | I am concerned about the environmental impact of the production of farmed fish | 59.55 | 10.11 |  |  |  |  |  |  |  |  |  |
| q17_ | 6 | I am concerned about the environmental impact of cashing wild fish | 76.96 | 5.06 |  |  |  |  |  |  |  |  |  |
| q17_ | 7 | I am concerned about the environmental sustainability of fish farming | 69.66 | 9.54 |  |  |  |  |  |  |  |  |  |
| q17_ | 8 | I am concerned about the environmental sustainability of fisheries of wild fish | 76.27 | 6.21 |  |  |  |  |  |  |  |  |  |
| q17_ | 9 | I am concerned about the welfare of farmed fish | 51.68 | 20.22 |  |  |  |  |  |  |  |  |  |
| q17 | 10 | I am concerned about the welfare of wild caught fish | 56.74 | 19.11 |  |  |  |  |  |  |  |  |  |

## A.11.1. Welch Two Sample t-test

A Welch Two Sample $t$-test was conducted to test whether the participants significantly differed in their concerns about the welfare of farmed and wild fish. The STATA output for the test follows in Table A.11.1.1.

Table A.11.1.1. Welch Two Sample t-test Regarding Concerns on Animal Welfare

| Two-sample t test with unequal variances |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Variable | Obs | Mean | Std. Err. | Std. Dev | [95\% Con | Interval] |
| q17 9xx | 166 | 6.698795 | . 2093942 | 2.697855 | 6.285358 | 7.112233 |
| q17_10xx | 167 | 6.958084 | . 2058507 | 2.660178 | 6.551661 | 7.364507 |
| combined | 333 | 6.828829 | . 146761 | 2.678137 | 6.54013 | 7.117528 |
| diff |  | -. 2592887 | . 2936332 |  | -. 8368993 | . 318322 |
| diff $=$ mean(q17_9xx) - mean (q17_10xx) $t=-0.8830$ |  |  |  |  |  |  |
| Ho: diff $=0$ Welch's degrees of freedom $=332.865$ |  |  |  |  |  |  |
| Ha: diff < 0 |  | Ha: diff ! = 0 |  |  | Ha: diff > 0 |  |
| $\operatorname{Pr}(\mathrm{T}<\mathrm{t}$ | . 1889 | $\operatorname{Pr}(\|T\|>\|t\|)=0.3779$ |  |  | $\operatorname{Pr}(\mathrm{T}>\mathrm{t})=0.8111$ |  |

## A.12. Regression Outputs

This section contains all the regressions we ran. For some relevant models the corresponding covariance matrix and standard deviations follow. Model 1 and Model 2 are the models presented in the article. These two models, and Model 3 are the only models that have been given names in addition to a model number.

Model 1 - Product Attribute Model
Model 2 - Product Attribute and Consumer Characteristics Interaction Model

Model 3 - Product Attribute and Full Set of Consumer Characteristics Interaction Model

All models but Model 1 contain consumer characteristics interaction terms. The consumer characteristics that are included in the models are given in the header. For example, in Model 6 - Income and Education, the fish types, represented by the $\boldsymbol{z}_{i}$-vectors, are interacted with the consumer characteristics Income and Education.

Each model was estimated by both allowing for correlations in unobserved factors over choice scenarios and not allowing for it. A model that allows for it is simply assigned a number, e.g.: Model 4 or Model 5. The corresponding model that does not allow for correlations in unobserved factors over choice scenarios is given the same number, but with a " $B$ " in front of the number, e.g.: Model B. 4 or Model B.5. All the models that allow for correlations in unobserved factors over choice scenarios are presented first. The corresponding models that do not allow for it follow after. The variable names in the regression outputs are different from those presented in the article. Table A. 12.1 presents the variable names used in STATA for each variable and interaction term.

We used STATA 12 to analyze the data. We used two extensions to STATA 12, namely mixlogit and wtp. Both extensions are created by Arne Risa Hole (see Hole (2007b) and Hole (2007a)). The wtp extension was "delta" by default (see Hole (2007a)). We chose to use 500 Halton draws for the simulations by the mixlogit program. Hole (2007b) suggests using 500 Halton draws for the final model.

Table A.12.1. The Variable Names and Interaction Terms as Defined in STATA

| Article Output Label | STATA Output Label | Article Output Label | STATA Output Label |
| :---: | :---: | :---: | :---: |
| Price | p1000 | Age * Salmon | sa_age |
| Salmon | sa | Age * Farmed_Cod | fc_age |
| Farmed_Cod | fc | Age * Wild_Cod | wc_age |
| Wild_Cod | wc | Age * Monk | mo_age |
| Monk | mo | Age * Pangasius | pa_age |
| Pangasius | pa | Single * Salmon | sa_single |
| Tail * Salmon | sa_tail | Single * Farmed_Cod | fc_single |
| Tail * Farmed_Cod | wc_tail | Single * Wild_Cod | wc_single |
| Tail * Wild_Cod | fc_tail | Single * Monk | mo_single |
| Income * Salmon | sa_inc3 | Single * Pangasius | pa_single |
| Income * Farmed_Cod | wc_inc3 | $\begin{aligned} & \text { High_Educ * Salmon } \\ & \text { High_Educ * } \end{aligned}$ | sa_educUNI |
| Income * Wild_Cod | fc_inc3 | Farmed Cod High_Educ * | fc_educUNI |
| Income * Monk | mo_inc3 | Wild_Cod | wc_educUNI |
| Income * Pangasius | pa_inc3 | $\begin{aligned} & \text { High_Educ } * \text { Monk } \\ & \text { High_Educ }^{*} \end{aligned}$ | mo_educUNI |
| DNWA * Salmon | sa_nonInc | Pangasius | pa_educUNI |
| DNWA * Farmed_Cod | fc_nonInc | Female * Salmon | femaleSA |
| DNWA * Wild_Cod | wc_nonInc | Female * Farmed_Cod | femaleFC |
| DNWA * Monk | mo_nonInc | Female * Wild_Cod | femaleWC |
| DNWA * Pangasius | pa_nonInc | Female * Monk | femaleMO |
| $\begin{aligned} & \text { IncGroup } 2 * \text { Salmon } \\ & \text { IncGroup } 2 * \end{aligned}$ | sa_inc_2 | Female * Pangasius | femalePA |
| Farmed_Cod | fc_inc_2 | $\begin{aligned} & \text { Children * Salmon } \\ & \text { Children * } \end{aligned}$ | childrenSA |
| IncGroup 2 * Wild_Cod | wc_inc_2 | Farmed_Cod | childrenFC |
| IncGroup 2 * Monk IncGroup 2 * | mo_inc_2 | Children * Wild_Cod | childrenWC |
| Pangasius | pa_inc_2 | Children * Monk <br> Children * Pangasius | childrenMO <br> childrenPA |

## A.12.1. Model 1 - Product Attribute Model

Table A.12.1.1. Model 1 - Product Attribute Model


Table A.12.1.2. Covariance Matrix Model 1

| y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| v11 | 3.095808 | . 5139164 | 6.02 | 0.000 | 2.08855 | 4.103066 |
| v21 | 2.006123 | . 3854116 | 5.21 | 0.000 | 1.25073 | 2.761515 |
| v31 | 1.519497 | . 3147228 | 4.83 | 0.000 | . 9026516 | 2.136342 |
| v41 | 2.70131 | . 7478071 | 3.61 | 0.000 | 1.235635 | 4.166985 |
| v51 | . 9953472 | . 4599499 | 2.16 | 0.030 | . 0938619 | 1.896833 |
| v22 | 4.970637 | . 9918143 | 5.01 | 0.000 | 3.026716 | 6.914557 |
| v32 | 3.579057 | . 622302 | 5.75 | 0.000 | 2.359367 | 4.798746 |
| v42 | 1.253135 | . 9090898 | 1.38 | 0.168 | -. 5286479 | 3.034919 |
| v52 | 3.757281 | . 950221 | 3.95 | 0.000 | 1.894882 | 5.61968 |
| v33 | 4.379695 | . 7231669 | 6.06 | 0.000 | 2.962314 | 5.797077 |
| v43 | . 6614561 | . 8292742 | 0.80 | 0.425 | -. 9638915 | 2.286804 |
| v53 | 4.358352 | 1.000834 | 4.35 | 0.000 | 2.396754 | 6.319951 |
| v44 | 10.62544 | 3.041254 | 3.49 | 0.000 | 4.664696 | 16.58619 |
| v54 | 1.277392 | 1.250936 | 1.02 | 0.307 | -1.174397 | 3.72918 |
| v55 | 7.03414 | 1.922035 | 3.66 | 0.000 | 3.267021 | 10.80126 |

Table A.12.1.3. Standard Deviations Model 1

| Y | Coef. | Std. Err. | z | $P>\|z\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| sa | 1.759491 | .1460412 | 12.05 | 0.000 | 1.473255 | 2.045726 |
| fc | 2.229492 | . 2224305 | 10.02 | 0.000 | 1.793537 | 2.665448 |
| wc | 2.092772 | . 1727773 | 12.11 | 0.000 | 1.754135 | 2.431409 |
| pa | 3.259669 | . 4664973 | 6.99 | 0.000 | 2.345351 | 4.173987 |
| mo | 2.652195 | . 3623479 | 7.32 | 0.000 | 1.942007 | 3.362384 |

Table A.12.1.4. WTP Table Model 1
wtp p1000 sa fc wc mo pa sa_tail fc_tail wc_tail

|  | sa | fc | WC | mo | pa | sa tail | fc tail | wc tail |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| wtp | 13.687511 | 10.780963 | 13.350375 | 12.760482 | -6.9386952 | -1.9653784 | -6.5537952 | -3.0482884 |
| 11 | 12.19247 | 8.3596955 | 11.454519 | 10.028299 | -11.723679 | -2.9614462 | -8.1472395 | -4.330633 |
| ul | 15.182552 | 13.202231 | 15.246232 | 15.492666 | -2.1537117 | -. 96931054 | -4.9603509 | -1.7659438 |

Table A.12.2.1. Model 2 - Product Attribute and Consumer Characteristics Interaction Model


Table A.12.2.2. Covariance Matrix Model 2

| Y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| v11 | 2.822967 | . 4518939 | 6.25 | 0.000 | 1.937271 | 3.708663 |
| v21 | 1.784098 | . 353448 | 5.05 | 0.000 | 1.091353 | 2.476844 |
| v31 | 1.276213 | . 3063345 | 4.17 | 0.000 | . 6758079 | 1.876617 |
| v41 | 2.422724 | . 6872761 | 3.53 | 0.000 | 1.075688 | 3.769761 |
| v51 | . 5462072 | . 4020077 | 1.36 | 0.174 | -. 2417135 | 1.334128 |
| v22 | 4.572038 | . 8510785 | 5.37 | 0.000 | 2.903955 | 6.240121 |
| v32 | 3.153135 | . 5249271 | 6.01 | 0.000 | 2.124297 | 4.181973 |
| v42 | 1.199411 | . 6952826 | 1.73 | 0.085 | -. 1633184 | 2.562139 |
| v52 | 3.113368 | . 7121832 | 4.37 | 0.000 | 1.717515 | 4.509222 |
| v33 | 3.564117 | . 604374 | 5.90 | 0.000 | 2.379566 | 4.748669 |
| v43 | . 4433143 | . 7474009 | 0.59 | 0.553 | -1.021565 | 1.908193 |
| v53 | 3.802049 | . 692995 | 5.49 | 0.000 | 2.443804 | 5.160295 |
| v44 | 10.15173 | 3.054783 | 3.32 | 0.001 | 4.16447 | 16.139 |
| v54 | . 8192685 | 1.165263 | 0.70 | 0.482 | -1.464605 | 3.103142 |
| v55 | 6.244267 | 1.436091 | 4.35 | 0.000 | 3.429581 | 9.058952 |

Table A.12.2.3. Standard Deviation Model 2

| y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| sa | 1.680169 | . 1344787 | 12.49 | 0.000 | 1.416595 | 1.943742 |
| fc | 2.138232 | . 1990145 | 10.74 | 0.000 | 1.748171 | 2.528294 |
| wc | 1.887887 | . 1600662 | 11.79 | 0.000 | 1.574163 | 2.201611 |
| pa | 3.186179 | . 4793804 | 6.65 | 0.000 | 2.24661 | 4.125747 |
| mo | 2.498853 | . 2873499 | 8.70 | 0.000 | 1.935658 | 3.062049 |

## Table A.12.2.4. WTP Model 2

|  | Mean | Lower <br> limit | Upper <br> limit |
| :---: | :---: | :---: | :---: |
| sa | 16.937258 | 11.959101 | 21.915415 |
| fc | 4.2175381 | $-3.103475$ | 11.538551 |
| wc | 5.0945647 | 0.92297773 | 11.112107 |
| mo | 4.7723843 | -4.3102845 | 13.855053 |
| pa | -6.9310872 | $-18.30411$ | 4.4419357 |
| sa_inc3 | 2.4026267 | 0.75510633 | 5.5603596 |
| fc_inc3 | 3.9177922 | 0.31656666 | 8.152151 |
| wc_inc3 | 5.6221293 | 2.1085425 | 9.135716 |
| mo_inc3 | 1.3717002 | -3.839846 | 6.5832465 |
| pa_inc3 | 4.1993857 | -2.113995 | 10.512766 |
| sa*nonInc | -4.507488 | -9.2861818 | 0.27120574 |
| $\mathrm{fc}^{*}$ nonInc | -3.4048105 | -10.125023 | 3.3154015 |
| wc*nonInc | -3.9535512 | -9.3440448 | 1.4369423 |
| mo*nonInc | -11.197443 | -19.836947 | -2.55794 |
| $\mathrm{pa*}$ nonInc | -4.716034 | -18.283771 | 8.8517026 |
| sa*age | -0.1019663 | $0.20121554$ | 0.00271706 |
| fc*age | 0.10382196 | 0.02863731 | 0.23628123 |
| wc*age | 0.13281048 | 0.01897163 | 0.24664934 |
| mo*age | $0.17370465$ | $0.00273429$ | 0.344675 |
| pa*age | 0.04726644 | 0.25905983 | 0.16452695 |
| sa*single | 4.4254843 | $1.5645142$ | 7.2864545 |
| fc*single | 2.905658 | $0.85362495$ | 6.664941 |
| wc*single | 3.0743863 | 0.16789544 | 6.3166681 |
| mo*single | 1.0826956 | $-3.6172375$ | 5.7826288 |
| pa*single | 5.511584 | 0.39932593 | $11.422494$ |
| sa*tail | -1.9607626 | -2.9533725 | 0.96815265 |
| fc*tail | -6.5867473 | -8.1686491 | -5.0048456 |
| wc*tail | -3.0564578 | -4.3302702 | -1.7826454 |

A.12.3. Model 3 - Product Attribute and Full Set of Consumer Characteristics Interaction Model

Table A.12.3.1. Model 3 - Product Attribute and Full Set of Consumer Characteristics Interaction Model

| Mixed logit model <br> Log likelihood = -2505.3946 |  |  |  | $\begin{aligned} & \text { Number of obs } \\ & \text { LR chi2 (15) } \\ & \text { Prob > chi2 } \end{aligned}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Con | Interval] |
| p1000 | -. 2194887 | . 010346 | -21.21 | 0.000 | -. 2397665 | -. 1992109 |
| sa_inc3 | . 4867803 | . 3889622 | 1.25 | 0.211 | -. 2755716 | 1.249132 |
| fc_inc3 | . 8151295 | . 526986 | 1.55 | 0.122 | -. 217744 | 1.848003 |
| wc_inc3 | 1.093393 | . 4415729 | 2.48 | 0.013 | . 2279258 | 1.95886 |
| mo-inc3 | . 093985 | . 6532151 | 0.14 | 0.886 | -1.186293 | 1.374263 |
| pa_inc3 | . 5917403 | . 8679535 | 0.68 | 0.495 | -1.109417 | 2.292898 |
| sa_nonInc | -. 9771875 | . 4833835 | -2.02 | 0.043 | -1.924602 | -. 0297733 |
| fc_nonInc | -. 7126866 | . 7042876 | -1.01 | 0.312 | -2.093065 | . 6676916 |
| wc_nonInc | -. 8825129 | . 5675379 | -1.55 | 0.120 | -1.994867 | . 2298409 |
| mo_nonInc | -2.413946 | . 955722 | -2.53 | 0.012 | -4.287126 | -. 540765 |
| pa_nonInc | -. 8682049 | 1.013254 | -0.86 | 0.392 | -2.854147 | 1.117737 |
| childrenSA | -. 0423645 | . 3224722 | -0.13 | 0.895 | -. 6743985 | . 5896695 |
| childrenFC | -. 0322402 | . 5283185 | -0.06 | 0.951 | -1.067725 | 1.003245 |
| childrenWC | -. 457222 | . 4413726 | -1.04 | 0.300 | -1.322296 | . 4078524 |
| childrenMO | -. 5114613 | . 6321563 | -0.81 | 0.418 | -1.750465 | . 7275423 |
| childrenPA | . 0383729 | . 6608988 | 0.06 | 0.954 | -1.256965 | 1.333711 |
| femaleSA | . 1866528 | . 2903716 | 0.64 | 0.520 | -. 3824651 | . 7557707 |
| femaleFC | -. 5143591 | . 3704646 | -1.39 | 0.165 | -1.240456 | . 2117381 |
| femaleWC | -. 2924441 | . 3198976 | -0.91 | 0.361 | -. 9194319 | . 3345438 |
| femaleMO | -. 2733139 | . 4728027 | -0.58 | 0.563 | -1.19999 | . 6533624 |
| femalePA | -. 1784797 | . 6014022 | -0.30 | 0.767 | -1.357206 | 1.000247 |
| sa_age | -. 0170753 | . 0125096 | -1.36 | 0.172 | -. 0415937 | . 0074431 |
| fc_age | . 0238793 | . 0180901 | 1.32 | 0.187 | -. 0115766 | . 0593352 |
| wc_age | . 0270251 | . 015737 | 1.72 | 0.086 | -. 0038189 | . 0578691 |
| mo_age | . 0374012 | . 0231314 | 1.62 | 0.106 | -. 0079354 | . 0827379 |
| pa_age | -. 0027111 | . 028183 | -0.10 | 0.923 | -. 0579487 | . 0525265 |
| sa_single | . 9107461 | . 3615915 | 2.52 | 0.012 | . 2020399 | 1. 619452 |
| fc_single | . 5142746 | . 4827878 | 1.07 | 0.287 | -. 4319721 | 1.460521 |
| wc_single | . 3818803 | . 4308048 | 0.89 | 0.375 | -. 4624817 | 1.226242 |
| mo_single | -. 073355 | . 6348796 | -0.12 | 0.908 | -1.317696 | 1.170986 |
| pa_single | . 8193952 | .7271502 | 1.13 | 0.260 | -. 605793 | 2.244583 |
| sa_educUNI | . 2408917 | . 3488066 | 0.69 | 0.490 | -. 4427566 | . 9245401 |
| fc_educUNI | . 2873476 | . 429614 | 0.67 | 0.504 | -. 5546803 | 1.129375 |
| wc_educUnI | . 36613 | . 3512717 | 1.04 | 0.297 | -. 3223498 | 1.05461 |
| mo_educUNI | . 3559302 | . 5522768 | 0.64 | 0.519 | -. 7265125 | 1.438373 |
| pa_educUNI | . 4806885 | . 7825609 | 0.61 | 0.539 | -1.053103 | 2.01448 |
| -sa_tail | -. 4292298 | . 1176437 | -3.65 | 0.000 | -. 6598072 | -. 1986524 |
| $\mathrm{fc}_{-}^{-}$tail | -1.440224 | . 1928167 | -7.47 | 0.000 | -1.818138 | -1.062311 |
| wc_tail | -. 6605438 | . 1517246 | -4.35 | 0.000 | -. 9579186 | -. 363169 |
| sa | 3.334567 | . 7519284 | 4.43 | 0.000 | 1.860815 | 4.80832 |
| fc | 1.144092 | 1.11339 | 1.03 | 0.304 | -1.038113 | 3.326298 |
|  | 1.531906 | . 9517272 | 1.61 | 0.107 | -. 3334447 | 3.397257 |
| pa | -1.662705 | 1.756757 | -0.95 | 0.344 | -5.105886 | 1.780476 |
| mo | 1.427556 | 1.407913 | 1.01 | 0.311 | -1.331904 | 4.187015 |
| /111 | 1.64493 | . 1376219 | 11.95 | 0.000 | 1.375196 | 1.914664 |
| /121 | 1.090176 | . 2158634 | 5.05 | 0.000 | . 6670913 | 1.51326 |
| /131 | . 804167 | . 1915132 | 4.20 | 0.000 | . 428808 | 1.179526 |
| /141 | 1.190066 | . 4578423 | 2.60 | 0.009 | . 2927117 | 2.08742 |
| /151 | . 3675639 | . 257511 | 1.43 | 0.153 | -. 1371484 | . 8722763 |
| /122 | 1.835032 | . 1925291 | 9.53 | 0.000 | 1.457682 | 2.212382 |
| /132 | 1.242848 | . 1714953 | 7.25 | 0.000 | . 9067232 | 1.578973 |
| /142 | -. 2806761 | . 4464544 | -0.63 | 0.530 | -1.155711 | . 5943584 |
| /152 | 1.482551 | . 2785503 | 5.32 | 0.000 | . 9366023 | 2.028499 |
| /133 | 1.145908 | . 1529339 | 7.49 | 0.000 | . 8461625 | 1.445653 |
| /143 | -. 7408717 | . 3854691 | -1.92 | 0.055 | -1.496377 | . 0146338 |
| /153 | 1.348948 | . 3066513 | 4.40 | 0.000 | . 7479225 | 1.949974 |
| /144 | 2.599921 | . 3769456 | 6.90 | 0.000 | 1.861121 | 3.33872 |
| /154 | . 4064174 | . 2911178 | 1.40 | 0.163 | -. 164163 | . 9769977 |
| /155 | 1.29493 | . 3328663 | 3.89 | 0.000 | . 6425239 | 1.947336 |

## A.12.4. Model 4 - Income as Dummies for all Income Groups

Table A.12.4.1. Model 4 - Income as Dummies for all Income Groups

| Mixed logit model <br> Log likelihood $=-2517.5306$ |  |  |  | Number of obs LR chi2(15) Prob > chi2 |  | $\begin{array}{r} 11380 \\ 897.63 \\ 0.0000 \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| y I | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Con | Interval] |
| p1000 | -. 2195032 | . 0103191 | -21.27 | 0.000 | -. 2397282 | -. 1992781 |
| sa_inc_2 \| | -. 7769536 | . 3384176 | -2.30 | 0.022 | -1.44024 | -. 1136672 |
| fc_inc_2 \| | -. 0612313 | . 452569 | -0.14 | 0.892 | -. 9482502 | . 8257876 |
| wc_inc_2 \| | -. 3891837 | . 4057793 | -0.96 | 0.338 | -1.184496 | . 4061291 |
| mo_inc_2 \| | . 4058495 | . 5763738 | 0.70 | 0.481 | -. 7238224 | 1.535521 |
| pa_inc_2 \| | -. 8115087 | . 7646674 | -1.06 | 0.289 | -2.310229 | . 6872118 |
| sa_inc_3 \| | -. 3330455 | . 4038338 | -0.82 | 0.410 | -1.124545 | . 4584541 |
| fc_inc_3 \| | . 8657054 | . 5353571 | 1.62 | 0.106 | -. 1835751 | 1.914986 |
| wc_inc_3 \| | 1.038777 | . 4409634 | 2.36 | 0.018 | . 1745044 | 1.903049 |
| mo_inc_3 \| | . 7465031 | . 690635 | 1.08 | 0.280 | -. 6071167 | 2.100123 |
| pa_inc_3 \| | -. 0228213 | . 8905516 | -0.03 | 0.980 | -1.76827 | 1.722628 |
| sa_nonInc \| | -1.498905 | . 5034964 | -2.98 | 0.003 | -2.485739 | -. 5120698 |
| fc_nonInc \| | -. 8279735 | . 8363707 | -0.99 | 0.322 | -2.46723 | . 8112828 |
| wc_nonInc \| | -1.074317 | . 6855174 | -1.57 | 0.117 | -2.417906 | . 2692726 |
| mo_nonInc \| | -2.165486 | 1.106372 | -1.96 | 0.050 | -4.333935 | . 0029635 |
| pa_nonInc \| | -1.680657 | 1.244749 | -1.35 | 0.177 | -4.120321 | . 7590072 |
| sa_tail \| | -. 4325366 | . 1173489 | -3.69 | 0.000 | -. 6625362 | -. 202537 |
| fc_tail \| | -1.435866 | . 1928649 | -7.44 | 0.000 | -1.813875 | -1.057858 |
| wc_tail \| | -. 6613276 | . 1517063 | -4.36 | 0.000 | -. 9586664 | -. 3639888 |
| sa \| | 3.50687 | . 3177517 | 11.04 | 0.000 | 2.884088 | 4.129652 |
| fc \| | 2.26044 | . 4333511 | 5.22 | 0.000 | 1.411087 | 3.109792 |
| wc \| | 2.915238 | . 3795725 | 7.68 | 0.000 | 2.17129 | 3.659187 |
| pa \| | -. 9377152 | . 6343454 | -1.48 | 0.139 | -2.181009 | . 3055788 |
| mo \| | 2.657833 | . 5360118 | 4.96 | 0.000 | 1.607269 | 3.708397 |
| /111 \| | 1.693189 | . 1359409 | 12.46 | 0.000 | 1.42675 | 1.959628 |
| /121 \| | 1.128895 | . 1896172 | 5.95 | 0.000 | . 7572525 | 1.500538 |
| 1131 \| | . 7689937 | . 1810062 | 4.25 | 0.000 | . 4142281 | 1.123759 |
| 1141 \| | 1.296033 | . 3934924 | 3.29 | 0.001 | . 5248022 | 2.067264 |
| 1151 \| | . 4136651 | . 2652549 | 1.56 | 0.119 | -. 106225 | . 9335551 |
| /122 \| | 1.918423 | . 1919702 | 9.99 | 0.000 | 1.542168 | 2.294677 |
| 1132 \| | 1.320833 | . 1778442 | 7.43 | 0.000 | . 9722653 | 1.669402 |
| 1142 \| | -. 3203226 | . 4305534 | -0.74 | 0.457 | -1.164192 | . 5235466 |
| /152 \| | 1.607423 | . 2750429 | 5.84 | 0.000 | 1.068349 | 2.146497 |
| /133 \| | 1.256884 | . 1520302 | 8.27 | 0.000 | . 9589107 | 1.554858 |
| /143 \| | -. 7159321 | . 4823603 | -1.48 | 0.138 | -1.661341 | . 2294767 |
| /153 \| | 1.494459 | . 2766224 | 5.40 | 0.000 | . 9522894 | 2.036629 |
| /144 \| | 2.6189 | . 4115005 | 6.36 | 0.000 | 1.812374 | 3.425426 |
| /154 \| | . 2747836 | . 2700698 | 1.02 | 0.309 | -. 2545435 | . 8041107 |
| /155 \| | 1.340948 | . 3087385 | 4.34 | 0.000 | . 7358322 | 1.946065 |

## A.12.5. Model 5 - Income as Low and High

Table A.12.5.1. Model 5 - Income as Low and High

| Mixed logit model <br> Log likelihood $=-2521.5856$ |  |  |  | Number of obs <br> LR chi2(15) <br> Prob > chi2 |  | $\begin{array}{r} 11380 \\ 902.57 \\ 0.0000 \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Con | Interval] |
| p1000 | -. 218674 | . 0103005 | -21.23 | 0.000 | -. 2388626 | -. 1984854 |
| sa_inc_3 | . 1101639 | . 3616211 | 0.30 | 0.761 | -. 5986005 | . 8189283 |
| fc_inc_3 | . 8770205 | . 4730104 | 1.85 | 0.064 | -. 0500628 | 1.804104 |
| wc_inc_3 | 1.236088 | . 3782625 | 3.27 | 0.001 | . 4947067 | 1.977468 |
| mo_inc_3 | . 4897637 | . 5919796 | 0.83 | 0.408 | -. 670495 | 1.650022 |
| pa_inc_3 | . 474978 | . 7421392 | 0.64 | 0.522 | -. 9795881 | 1.929544 |
| sa_nonInc | -1.043095 | . 4669096 | -2.23 | 0.025 | -1.958221 | -. 1279686 |
| fc_nonInc | -. 794113 | . 7931808 | -1.00 | 0.317 | -2.348719 | . 7604928 |
| wc_nonInc | -. 8742903 | . 6460249 | -1.35 | 0.176 | -2.140476 | . 3918952 |
| mo-nonInc | -2.415986 | 1.05691 | -2.29 | 0.022 | -4.487491 | -. 3444803 |
| pa_nonInc | -1.128544 | 1.089826 | -1.04 | 0.300 | -3.264564 | 1.007475 |
| sa_tail | -. 430767 | . 1172108 | -3.68 | 0.000 | -. 660496 | -. 201038 |
| fc_tail | -1.430036 | . 1926131 | -7.42 | 0.000 | -1.80755 | -1.052521 |
| wc_tail | -. 6610077 | . 1516627 | -4.36 | 0.000 | -. 9582612 | -. 3637543 |
| sa | 3.046131 | . 2533444 | 12.02 | 0.000 | 2.549585 | 3.542677 |
| fc | 2.228137 | . 3579618 | 6.22 | 0.000 | 1.526544 | 2.929729 |
| wc | 2.697999 | . 3066245 | 8.80 | 0.000 | 2.097026 | 3.298972 |
|  | $-1.402489$ | $.5454625$ | -2.57 | 0.010 | $-2.471576$ | -. 3334023 |
| mo | 2.93149 | . 4015716 | 7.30 | 0.000 | $2.144424$ | 3.718556 |
| /111 | 1.70157 | . 1373827 | 12.39 | 0.000 | 1.432304 | 1.970835 |
| /121 | 1.135433 | . 1843261 | 6.16 | 0.000 | . 7741603 | 1.496705 |
| /131 | . 8115959 | . 1727103 | 4.70 | 0.000 | . 4730899 | 1.150102 |
| /141 | 1.408446 | . 4118279 | 3.42 | 0.001 | . 6012777 | 2.215613 |
| /151 | . 4179797 | . 2755433 | 1.52 | 0.129 | -. 1220752 | . 9580347 |
| /122 | 1.905966 | . 1910407 | 9.98 | 0.000 | 1.531533 | 2.280399 |
| /132 | 1.308158 | . 1787839 | 7.32 | 0.000 | . 957748 | 1.658568 |
| /142 | -. 2603717 | . 3507424 | -0.74 | 0.458 | -. 9478142 | . 4270708 |
| /152 | 1.59451 | . 2871397 | 5.55 | 0.000 | 1.031726 | 2.157293 |
| /133 | 1.259543 | . 1520411 | 8.28 | 0.000 | . 9615477 | 1.557538 |
| /143 | -. 6958762 | . 4744297 | -1.47 | 0.142 | -1.625741 | . 233989 |
| 1153 | 1.432153 | . 3091793 | 4.63 | 0.000 | . 826173 | 2.038134 |
| /144 | 2.552167 | . 3516405 | 7.26 | 0.000 | 1.862964 | 3.24137 |
| 1154 | . 3714819 | . 3251436 | 1.14 | 0.253 | -. 2657879 | 1.008752 |
| /155 | 1.311247 | . 3268253 | 4.01 | 0.000 | . 6706813 | 1.951813 |

## A.12.6. Model 6 - Income and Education

Table A.12.6.1. Model 6 - Income and Education

| Mixed logit model |  |  |  | Number of obs LR chi2(15) |  | $\begin{array}{r} 11380 \\ 897.81 \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Log likelihood | $=-2518.344$ |  |  |  |  | 0.0000 |
| Y | Coef. | Std. Err | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Con | Interval] |
| p1000 | -. 2196156 | . 0103334 | -21.25 | 0.000 | -. 2398687 | -. 1993624 |
| sa_inc3 | -. 0019033 | . 3769359 | -0.01 | 0.996 | -. 7406841 | . 7368775 |
| fc_inc3 | . 8979831 | . 4882901 | 1.84 | 0.066 | -. 0590479 | 1.855014 |
| wc_inc3 | 1.251629 | . 3962915 | 3.16 | 0.002 | . 474912 | 2.028346 |
| mo-inc3 | . 4714741 | . 5890864 | 0.80 | 0.424 | -. 6831141 | 1.626062 |
| pa_inc3 | . 29817 | . 9351084 | 0.32 | 0.750 | -1.534609 | 2.130949 |
| sa_nōnInc | -1.15059 | . 4704385 | -2.45 | 0.014 | -2.072632 | -. 2285475 |
| fc_nonInc | -. 8981883 | . 7865899 | -1.14 | 0.254 | -2.439876 | . 6434996 |
| wc_nonInc | -. 9567357 | . 6312696 | -1.52 | 0.130 | -2.194001 | . 2805299 |
| mo_nonInc | -2.473886 | 1.039379 | -2.38 | 0.017 | -4.511031 | -. 4367417 |
| pa_nonInc | -. 9889031 | . 9649234 | -1.02 | 0.305 | -2.880118 | . 902312 |
| sa_educUNI | . 6808361 | . 2959201 | 2.30 | 0.021 | . 1008434 | 1.260829 |
| fc_educUNI | . 2644023 | . 4034684 | 0.66 | 0.512 | -. 5263812 | 1.055186 |
| wc_educUNI | . 1390699 | . 3415994 | 0.41 | 0.684 | -. 5304526 | . 8085924 |
| mo_educUNI | -. 0344771 | . 5334476 | -0.06 | 0.948 | -1.080015 | 1.011061 |
| pa_educUNI | 1.038035 | . 6327745 | 1.64 | 0.101 | -. 2021804 | 2.27825 |
| sa_tail | -. 433119 | . 1174052 | -3.69 | 0.000 | -. 663229 | -. 2030091 |
| fc_tail | -1.434653 | . 1925815 | -7.45 | 0.000 | -1.812106 | -1.0572 |
| wc_tail | -. 6640525 | . 151663 | -4.38 | 0.000 | -. 9613066 | -. 3667985 |
| sa | 2.868314 | . 2598693 | 11.04 | 0.000 | 2.35898 | 3.377648 |
| fc | 2.180555 | . 3869034 | 5.64 | 0.000 | 1.422238 | 2.938872 |
| wc | 2.683641 | . 3341939 | 8.03 | 0.000 | 2.028633 | 3.338649 |
| pa | -1.909587 | . 6960943 | -2.74 | 0.006 | -3.273907 | -. 545267 |
| mo | 2.960718 | . 4338914 | 6.82 | 0.000 | 2.110307 | 3.81113 |
| /111 | 1.704832 | . 1368482 | 12.46 | 0.000 | 1.436614 | 1.97305 |
| /121 | 1.115313 | . 1865238 | 5.98 | 0.000 | . 7497327 | 1.480893 |
| /131 | . 7752552 | . 1775589 | 4.37 | 0.000 | . 4272461 | 1.123264 |
| /141 | 1.421549 | . 372526 | 3.82 | 0.000 | . 6914121 | 2.151687 |
| /151 | . 350716 | . 2741945 | 1.28 | 0.201 | -. 1866954 | . 8881275 |
| /122 | 1.914757 | . 1928645 | 9.93 | 0.000 | 1.53675 | 2.292765 |
| /132 | 1.310011 | . 1729235 | 7.58 | 0.000 | . 9710868 | 1.648934 |
| /142 | -. 2298261 | . 316432 | -0.73 | 0.468 | -. 8500215 | . 3903692 |
| /152 | 1.60265 | . 2787617 | 5.75 | 0.000 | 1.056287 | 2.149013 |
| /133 | 1.251521 | . 1506425 | 8.31 | 0.000 | . 956267 | 1.546775 |
| /143 | -. 8857156 | . 3800422 | -2.33 | 0.020 | -1.630585 | -. 1408465 |
| /153 | 1.431985 | . 2817696 | 5.08 | 0.000 | . 8797267 | 1.984243 |
| /144 | 2.665041 | . 3584521 | 7.43 | 0.000 | 1.962488 | 3.367595 |
| /154 | . 326595 | . 2941819 | 1.11 | 0.267 | -. 249991 | . 903181 |
| /155 | 1.351829 | . 3207825 | 4.21 | 0.000 | . 7231066 | 1.980551 |

A.12.7. Model 7 - Income and Age

Table A.12.7.1. Model 7 - Income and Age

| Mixed logit model <br> Log likelihood $=-2517.9007$ |  |  |  | Number of obs LR chi2(15) Prob > chi2 |  | $\begin{array}{r} 11380 \\ 861.84 \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Con | Interval] |
| p1000 | -. 2151495 | . 010208 | -21.08 | 0.000 | -. 2351568 | -. 1951422 |
| sa_inc3 \| | . 084258 | . 3358955 | 0.25 | 0.802 | -. 5740851 | . 7426011 |
| fc_inc3 \| | . 5862758 | . 4463335 | 1.31 | 0.189 | -. 2885219 | 1.461073 |
| wc_inc3 \| | . 9592396 | . 3709346 | 2.59 | 0.010 | . 2322211 | 1.686258 |
| mo-inc3 \| | . 1019458 | . 5292575 | 0.19 | 0.847 | -. 9353798 | 1.139271 |
| pa_inc3 \| | . 1590509 | . 6663165 | 0.24 | 0.811 | -1.146905 | 1.465007 |
| sa_nōnInc \| | -1.052047 | . 4649799 | -2.26 | 0.024 | -1.96339 | -. 1407026 |
| fc_nonInc | -. 7570505 | . 6869432 | -1.10 | 0.270 | -2.103434 | . 5893335 |
| wc_nonInc \| | -. 9076407 | . 5611077 | -1.62 | 0.106 | -2.007392 | . 1921102 |
| mo_nonInc \| | -2.350384 | . 961045 | -2.45 | 0.014 | -4.233998 | -. 4667707 |
| pa_nonInc \| | -1.151207 | 1.135602 | -1.01 | 0.311 | -3.376946 | 1.074531 |
| sa_age \| | -. 0229618 | . 0110245 | -2.08 | 0.037 | -. 0445694 | -. 0013543 |
| fc_age \| | . 0204164 | . 0145689 | 1.40 | 0.161 | -. 0081381 | . 0489708 |
| wc_age \| | . 0246466 | . 0125158 | 1.97 | 0.049 | . 0001161 | . 0491771 |
| mo_age \| | . 0266071 | . 0183842 | 1.45 | 0.148 | -. 0094252 | . 0626395 |
| pa_age \| | -. 0091755 | . 0214981 | -0.43 | 0.670 | -. 051311 | . 0329599 |
| sa_Eail \| | -. 4163712 | . 1169928 | -3.56 | 0.000 | -. 6456729 | -. 1870695 |
| fc_tail \| | -1.414978 | . 1915506 | -7.39 | 0.000 | -1.79041 | -1.039546 |
| wc_tail \| | -. 6390379 | . 1506201 | -4.24 | 0.000 | -. 9342479 | -. 3438279 |
| sa | 4.022472 | . 5688519 | 7.07 | 0.000 | 2.907543 | 5.137402 |
| fc \| | 1.258656 | . 7752828 | 1.62 | 0.104 | -. 2608707 | 2.778182 |
| wc \| | 1.527926 | . 6512656 | 2.35 | 0.019 | . 2514686 | 2.804383 |
| pa \| | -1.003957 | 1.064403 | -0.94 | 0.346 | -3.090149 | 1.082236 |
| mo \| | 1.75693 | . 9416493 | 1.87 | 0.062 | -. 0886686 | 3.602529 |
| /111 \| | 1.704267 | . 1313796 | 12.97 | 0.000 | 1.446767 | 1.961766 |
| /121 \| | 1.004769 | . 2032118 | 4.94 | 0.000 | . 6064813 | 1.403057 |
| 1131 \| | . 8360079 | . 1906484 | 4.39 | 0.000 | . 4623439 | 1.209672 |
| 1141 \| | 1.497435 | . 3737926 | 4.01 | 0.000 | . 7648146 | 2.230055 |
| 1151 \| | . 2194099 | . 2978745 | 0.74 | 0.461 | -. 3644133 | . 8032331 |
| /122 \| | 1.795573 | . 1867867 | 9.61 | 0.000 | 1.429478 | 2.161668 |
| /132 | 1.168871 | . 1807858 | 6.47 | 0.000 | . 8145377 | 1.523205 |
| 1142 \| | -. 2323023 | . 4188673 | -0.55 | 0.579 | -1.053267 | . 5886626 |
| /152 | 1.571337 | . 3555452 | 4.42 | 0.000 | . 874481 | 2.268192 |
| 1133 \| | 1.143012 | . 1406888 | 8.12 | 0.000 | . 8672675 | 1.418757 |
| /143 | -. 4350394 | . 4027205 | -1.08 | 0.280 | -1.224357 | . 3542783 |
| /153 | 1.370541 | . 3054398 | 4.49 | 0.000 | . 77189 | 1.969192 |
| /144 | 2.738865 | . 5018014 | 5.46 | 0.000 | 1.755353 | 3.722378 |
| /154 \| | . 6665772 | . 341615 | 1.95 | 0.051 | -. 0029759 | 1.33613 |
| /155 \| | -. 770024 | . 5474523 | -1.41 | 0.160 | -1.843011 | . 3029627 |

Table A.12.7.2. Covariance Matrix Model 7

| y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| v11 | 2.904525 | . 4478116 | 6.49 | 0.000 | 2.02683 | 3.782219 |
| v21 | 1.712394 | . 3958005 | 4.33 | 0.000 | . 9366397 | 2.488149 |
| v31 | 1.42478 | . 3631854 | 3.92 | 0.000 | . 7129501 | 2.136611 |
| v41 | 2.552028 | . 7077866 | 3.61 | 0.000 | 1.164791 | 3.939264 |
| v51 | . 373933 | . 5102991 | 0.73 | 0.464 | -. 6262348 | 1.374101 |
| v22 | 4.233643 | . 7980914 | 5.30 | 0.000 | 2.669413 | 5.797874 |
| v32 | 2.938789 | . 4711008 | 6.24 | 0.000 | 2.015448 | 3.86213 |
| v42 | 1.08746 | . 6804656 | 1.60 | 0.110 | -. 2462279 | 2.421148 |
| v52 | 3.041906 | . 6769487 | 4.49 | 0.000 | 1.715111 | 4.368701 |
| v33 | 3.371647 | . 5593267 | 6.03 | 0.000 | 2.275387 | 4.467907 |
| v43 | . 4830803 | . 6641188 | 0.73 | 0.467 | -. 8185686 | 1.784729 |
| v53 | 3.586664 | . 6682135 | 5.37 | 0.000 | 2.27699 | 4.896339 |
| v44 | 9.986917 | 3.35447 | 2.98 | 0.003 | 3.412276 | 16.56156 |
| v54 | 1.192953 | 1.297021 | 0.92 | 0.358 | -1.349162 | 3.735068 |
| v55 | 5.432885 | 1.490515 | 3.64 | 0.000 | 2.511529 | 8.354241 |

Table A.12.7.3. Standard Deviations Model 7

| y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| sa | 1.704267 | . 1313796 | 12.97 | 0.000 | 1.446767 | 1.961766 |
| fc | 2.057582 | . 1939392 | 10.61 | 0.000 | 1.677468 | 2.437696 |
| wc | 1.836204 | . 1523051 | 12.06 | 0.000 | 1.537692 | 2.134717 |
| pa | 3.160208 | . 5307356 | 5.95 | 0.000 | 2.119986 | 4.200431 |
| mo | 2.330855 | . 3197357 | 7.29 | 0.000 | 1.704185 | 2.957525 |

## A.12.8. Model 8 - Income, Age, Single and Female

Table A.12.8.1. Model 8 - Income, Age, Single and Female

| Mixed logit model | Number of obs | $=$ | 11380 |
| :--- | :--- | :--- | :--- |
|  | LR chi2 (15) | $=$ | 843.48 |
| Log likelihood $=-2506.8224$ |  | Prob $>$ chi2 | $=$ |


| Y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| p1000 | -. 2197579 | . 010317 | -21.30 | 0.000 | -. 2399788 | -. 1995371 |
| femaleSA | . 123874 | . 2741471 | 0.45 | 0.651 | -. 4134444 | . 6611924 |
| femaleWC | -. 2918723 | . 3121699 | -0.93 | 0.350 | -. 9037141 | . 3199695 |
| femaleFC | -. 489887 | . 3538451 | -1.38 | 0.166 | -1.183411 | . 2036366 |
| femalePA | -. 5004762 | . 6189649 | -0.81 | 0.419 | -1.713625 | . 7126727 |
| femaleMO | -. 2102739 | . 4811883 | -0.44 | 0.662 | -1.153386 | . 7328379 |
| sa_inc3 | . 5477731 | . 3571898 | 1.53 | 0.125 | -. 152306 | 1.247852 |
| fc_inc3 | . 7990615 | . 4708823 | 1.70 | 0.090 | -. 1238509 | 1.721974 |
| wc_inc3 | 1.182974 | . 385728 | 3.07 | 0.002 | . 426961 | 1.938987 |
| mo_inc3 | . 2510835 | . 5856089 | 0.43 | 0.668 | -. 8966888 | 1.398856 |
| pa_inc3 | . 9620616 | . 7125767 | 1.35 | 0.177 | -. 434563 | 2.358686 |
| sa_nōInc | -. 9733456 | . 5079825 | -1.92 | 0.055 | -1.968973 | . 0222818 |
| fc_nonInc | -. 7136065 | . 7043145 | -1.01 | 0.311 | -2.094038 | . 6668245 |
| wc_nonInc | -. 8569513 | . 582229 | -1.47 | 0.141 | -1.998099 | . 2841966 |
| mo_nonInc | -2.457975 | . 9724547 | -2.53 | 0.011 | -4.363951 | -. 5519988 |
| pa_nonInc | -. 9859304 | 1.403494 | -0.70 | 0.482 | -3.736728 | 1.764868 |
| sa_age | -. 0204871 | . 0111652 | -1.83 | 0.067 | -. 0423705 | . 0013963 |
| fc_age | . 0219603 | . 0144203 | 1.52 | 0.128 | -. 006303 | . 0502236 |
| wc_age | . 0286662 | . 0128544 | 2.23 | 0.026 | . 0034721 | . 0538604 |
| mo_age | . 0385091 | . 0195852 | 1.97 | 0.049 | . 0001227 | . 0768954 |
| pa_age | -. 0168309 | . 0248467 | -0.68 | 0.498 | -. 0655296 | . 0318678 |
| sa_sin̄gle | . 9894323 | . 3199252 | 3.09 | 0.002 | . 3623905 | 1.616474 |
| fc_single | . 5867889 | . 416105 | 1.41 | 0.158 | -. 228762 | 1.40234 |
| wc_single | . 646393 | . 3637666 | 1.78 | 0.076 | -. 0665765 | 1.359362 |
| mo_single | . 2287842 | . 5304664 | 0.43 | 0.666 | -. 8109108 | 1.268479 |
| pa_single | 1.327733 | . 6594871 | 2.01 | 0.044 | . 0351623 | 2.620304 |
| sa_tail | -. 4309934 | . 117598 | -3.66 | 0.000 | -. 6614812 | -. 2005056 |
| fc_tail | -1.446216 | . 1925078 | -7.51 | 0.000 | -1.823524 | -1.068908 |
| wc_tail | -. 670536 | . 1515113 | -4.43 | 0.000 | -. 9674926 | -. 3735793 |
| sa | 3.555328 | . 6165059 | 5.77 | 0.000 | 2.346999 | 4.763657 |
| fc | 1.289367 | . 834272 | 1.55 | 0.122 | -. 3457757 | 2.92451 |
| wc | 1.337882 | . 7309043 | 1.83 | 0.067 | -. 0946645 | 2.770428 |
| pa | -1.02774 | 1.343611 | -0.76 | 0.444 | -3.661169 | 1.605689 |
| mo | 1.178493 | 1.13043 | 1.04 | 0.297 | -1.03711 | 3.394096 |
| /111 | 1.671968 | . 1342072 | 12.46 | 0.000 | 1.408926 | 1.935009 |
| /121 | 1.044116 | . 1703658 | 6.13 | 0.000 | . 7102052 | 1.378027 |
| /131 | . 7654823 | . 1520872 | 5.03 | 0.000 | . 4673969 | 1.063568 |
| /141 | 1.476313 | . 3613438 | 4.09 | 0.000 | . 7680923 | 2.184534 |
| /151 | . 3503324 | . 2324277 | 1.51 | 0.132 | -. 1052175 | . 8058823 |
| /122 | 1.820779 | . 1795855 | 10.14 | 0.000 | 1.468798 | 2.17276 |
| /132 | 1.244632 | . 1613029 | 7.72 | 0.000 | . 9284837 | 1.56078 |
| /142 | -. 2071591 | . 3209852 | -0.65 | 0.519 | -. 8362784 | . 4219603 |
| /152 | 1.494319 | . 2785776 | 5.36 | 0.000 | . 9483166 | 2.040321 |
| /133 | 1.187707 | . 1445744 | 8.22 | 0.000 | . 9043465 | 1.471068 |
| /143 | -. 3611484 | . 4134508 | -0.87 | 0.382 | -1.171497 | . 4492003 |
| /153 | 1.395301 | . 3157498 | 4.42 | 0.000 | . 7764427 | 2.014159 |
| /144 | 2.823938 | . 4231249 | 6.67 | 0.000 | 1.994629 | 3.653248 |
| /154 | . 3863853 | . 3015725 | 1.28 | 0.200 | -. 204686 | . 9774566 |
| /155 | 1.339671 | . 2977958 | 4.50 | 0.000 | . 7560024 | 1.92334 |

## A.12.9 Model 9 - Income, Age, Single, Female and Children

Table A.12.9.1. Model 9 - Income, Age, Single, Female and Children

| Mixed logit model <br> Log likelihood = -2505.8142 |  |  |  | Number of obs <br> LR chi2(15) <br> Prob $>$ chi2 |  | $\begin{array}{r} 11380 \\ 839.06 \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Con | Interval] |
| p1000 | -. 2198589 | . 0103257 | -21.29 | 0.000 | -. 2400968 | -. 1996209 |
| femaleSA | . 1454464 | . 277937 | 0.52 | 0.601 | -. 3993002 | . 690193 |
| femaleWC | -. 3010792 | . 3105315 | -0.97 | 0.332 | -. 9097097 | . 3075514 |
| femaleFC | -. 4930745 | . 3567683 | -1.38 | 0.167 | -1.192328 | . 2061786 |
| femalePA | -. 4955184 | . 6307272 | -0.79 | 0.432 | -1.731721 | . 7406842 |
| femalemo | -. 2295455 | . 4746401 | -0.48 | 0.629 | -1.159823 | . 700732 |
| childrenSA | -. 0353057 | . 3098367 | -0.11 | 0.909 | -. 6425745 | . 571963 |
| childrenWC | -. 4000821 | . 39037 | -1.02 | 0.305 | -1.165193 | . 3650289 |
| childrenFC | . 0392318 | . 4543832 | 0.09 | 0.931 | -. 8513428 | . 9298065 |
| childrenMO | -. 466433 | . 6006921 | -0.78 | 0.437 | -1.643768 | . 7109019 |
| childrenPA | . 0419306 | . 6703024 | 0.06 | 0.950 | -1.271838 | 1.355699 |
| sa_inc3 | . 5174595 | . 3637826 | 1.42 | 0.155 | -. 1955414 | 1.23046 |
| fc-inc3 | . 7992007 | . 4901997 | 1.63 | 0.103 | -. 1615731 | 1.759974 |
| wc_inc3 | 1.090697 | . 4073208 | 2.68 | 0.007 | . 292363 | 1.889031 |
| mo_inc3 | . 1529481 | . 6022155 | 0.25 | 0.800 | -1.027373 | 1.333269 |
| pa_inc3 | . 9002013 | . 7251458 | 1.24 | 0.214 | -. 5210584 | 2.321461 |
| sa_nonInc | -. 9846223 | . 5257765 | -1.87 | 0.061 | -2.015125 | . 0458807 |
| fc_nonInc | -. 6888907 | . 7003389 | -0.98 | 0.325 | -2.06153 | . 6837483 |
| wc_nonInc | -. 8673521 | . 5883962 | -1.47 | 0.140 | -2.020588 | . 2858834 |
| mo_nonInc | -2.423003 | . 9853925 | -2.46 | 0.014 | -4.354337 | -. 491669 |
| pa_nonInc | -. 9948132 | 1.533486 | -0.65 | 0.517 | -4.00039 | 2.010764 |
| sa_age | -. 0204262 | . 0114437 | -1.78 | 0.074 | -. 0428555 | . 0020032 |
| fc_age | . 0223059 | . 0151533 | 1.47 | 0.141 | -. 0073939 | . 0520058 |
| wc_age | . 0264057 | . 0131649 | 2.01 | 0.045 | . 000603 | . 0522084 |
| mo_age | . 0360115 | . 0201087 | 1.79 | 0.073 | -. 0034009 | . 0754238 |
| pa_age | -. 0178802 | . 0252722 | -0.71 | 0.479 | -. 0674128 | . 0316525 |
| sa_single | . 9777347 | . 3396624 | 2.88 | 0.004 | . 3120086 | 1.643461 |
| fc_single | . 5545705 | . 4583864 | 1.21 | 0.226 | -. 3438503 | 1.452991 |
| wc_single | . 4419415 | . 4145492 | 1.07 | 0.286 | -. 37056 | 1.254443 |
| mo_single | -. 012237 | . 5993753 | -0.02 | 0.984 | -1.186991 | 1.162517 |
| pa_single | 1.360751 | . 7326088 | 1.86 | 0.063 | -. 0751361 | 2.796637 |
| sa_tail | -. 4313691 | . 1176515 | -3.67 | 0.000 | -. 6619618 | -. 2007764 |
| $\mathrm{fc}_{-}^{-}$tail | -1.449361 | . 1925822 | -7.53 | 0.000 | -1.826815 | -1.071906 |
| wc_tail | -. 6720018 | . 1516148 | -4.43 | 0.000 | -. 9691613 | -. 3748424 |
| - sa | 3.578647 | . 6699721 | 5.34 | 0.000 | 2.265526 | 4.891768 |
| fc | 1.285368 | . 9302425 | 1.38 | 0.167 | -. 5378733 | 3.10861 |
| wc | 1.679327 | . 7879962 | 2.13 | 0.033 | . 1348828 | 3.223771 |
| pa | -. 9807675 | 1.454173 | -0.67 | 0.500 | -3.830895 | 1.86936 |
| mo | 1.561865 | 1.226682 | 1.27 | 0.203 | -. 8423887 | 3.966118 |
| 1111 | 1.672621 | . 1363033 | 12.27 | 0.000 | 1.405471 | 1.93977 |
| /121 | 1.037736 | . 1786289 | 5.81 | 0.000 | . 6876297 | 1.387842 |
| /131 | . 7820718 | . 1635336 | 4.78 | 0.000 | . 4615519 | 1.102592 |
| /141 | 1.47799 | . 3718436 | 3.97 | 0.000 | . 7491897 | 2.20679 |
| 1151 | . 372054 | . 2402698 | 1.55 | 0.122 | -. 0988661 | . 8429742 |
| /122 | 1.819196 | . 1811708 | 10.04 | 0.000 | 1.464108 | 2.174285 |
| /132 | 1.256885 | . 1603577 | 7.84 | 0.000 | . 9425899 | 1.571181 |
| /142 | -. 2241423 | . 3241843 | -0.69 | 0.489 | -. 8595319 | . 4112474 |
| /152 | 1.515707 | . 279291 | 5.43 | 0.000 | . 9683065 | 2.063107 |
| /133 | 1.156679 | . 1482599 | 7.80 | 0.000 | . 8660946 | 1.447263 |
| /143 | -. 3139335 | . 4294761 | -0.73 | 0.465 | -1.155691 | . 5278241 |
| /153 | 1.326243 | . 3222482 | 4.12 | 0.000 | . 6946482 | 1.957838 |
| /144 | 2.823529 | . 415281 | 6.80 | 0.000 | 2.009593 | 3.637465 |
| /154 | . 3867206 | . 2928387 | 1.32 | 0.187 | -. 1872327 | . 9606738 |
| /155 | 1.354235 | . 2972452 | 4.56 | 0.000 | . 7716448 | 1.936825 |

Table A.12.9.2. Covariance Matrix Model 9

| y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| v11 | 2.79766 | . 4559675 | 6.14 | 0.000 | 1.90398 | 3.69134 |
| v21 | 1.735739 | . 3503841 | 4.95 | 0.000 | 1.048999 | 2.422479 |
| v31 | 1.30811 | . 3077424 | 4.25 | 0.000 | . 7049456 | 1.911273 |
| v41 | 2.472116 | . 6911148 | 3.58 | 0.000 | 1.117556 | 3.826676 |
| v51 | . 6223053 | . 4073731 | 1.53 | 0.127 | -. 1761312 | 1.420742 |
| v22 | 4.386371 | . 8383722 | 5.23 | 0.000 | 2.743192 | 6.029551 |
| v32 | 3.098105 | . 530076 | 5.84 | 0.000 | 2.059175 | 4.137035 |
| v42 | 1.126004 | . 6408416 | 1.76 | 0.079 | -. 1300221 | 2.382031 |
| v52 | 3.143462 | . 7224766 | 4.35 | 0.000 | 1.727434 | 4.55949 |
| v33 | 3.529302 | . 631298 | 5.59 | 0.000 | 2.291981 | 4.766624 |
| v43 | . 5110527 | . 6893628 | 0.74 | 0.458 | -. 8400735 | 1.862179 |
| v53 | 3.73008 | . 7434865 | 5.02 | 0.000 | 2.272873 | 5.187286 |
| v44 | 10.30556 | 2.914619 | 3.54 | 0.000 | 4.593017 | 16.01811 |
| v54 | . 8857226 | 1.116484 | 0.79 | 0.428 | -1.302545 | 3.07399 |
| v55 | 6.178217 | 1.518608 | 4.07 | 0.000 | 3.2018 | 9.154634 |

Table A.12.9.3. Standard Deviations Model 9

| Y | Coef. | Std. Err. | z | $P>\|z\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| sa | 1.672621 | . 1363033 | 12.27 | 0.000 | 1.405471 | 1.93977 |
| fc | 2.094367 | . 2001493 | 10.46 | 0.000 | 1.702081 | 2.486652 |
| wc | 1.878644 | . 1680196 | 11.18 | 0.000 | 1.549331 | 2.207956 |
| pa | 3.210228 | . 4539582 | 7.07 | 0.000 | 2.320486 | 4.09997 |
| mo | 2.485602 | . 305481 | 8.14 | 0.000 | 1.88687 | 3.084334 |

Below are the regression outputs for the models that did not allow for correlations in unobserved factors over choice scenarios.

## A.12.10. Model B. 1 - Product Attribute Model (No Correlation)

Table A.12.10.1. Model B. 1 - Product Attribute Model (No Correlation)

26. The sign of the estimated standard deviations is irrelevant: interpret them as 7. being positive
A.12.11. Model B. 2 - Product Attribute and Consumer Characteristics Interaction Model (No Correlation)

Table A.12.11.1. Model B. 2 - Product Attribute and Consumer Characteristics Interaction Model (No Correlation)

| Mixed logit model | Number of obs | $=$ | 11380 |
| :--- | :--- | :--- | :--- |
|  | LR chi2 (5) | $=$ | 702.28 |
| Log likelihood $=-2584.7369$ | Prob $>$ chi2 | $=$ | 0.0000 |


| y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Mean |  |  |  |  |  |  |
| p1000 | -. 2212292 | . 010227 | -21.63 | 0.000 | -. 2412738 | -. 2011846 |
| sa_inc3 | . 3073738 | . 3178855 | 0.97 | 0.334 | -. 3156704 | . 930418 |
| fc_inc3 | . 3155598 | . 4656275 | 0.68 | 0.498 | -. 5970534 | 1.228173 |
| wc_inc3 | 1.468236 | . 3909246 | 3.76 | 0.000 | . 702038 | 2.234434 |
| mo_inc3 | . 1343882 | . 5690834 | 0.24 | 0.813 | -. 9809947 | 1.249771 |
| pa_inc3 | . 6131615 | . 6661159 | 0.92 | 0.357 | -. 6924016 | 1.918725 |
| sa_nōnInc | -1.108815 | . 4207675 | -2.64 | 0.008 | -1.933505 | -. 2841261 |
| fc_nonInc | -. 9502941 | . 6198928 | -1.53 | 0.125 | -2.165262 | . 2646734 |
| wc_nonInc | -. 7946081 | . 4957299 | -1.60 | 0.109 | -1.766221 | . 1770046 |
| mo_nonInc | -3.206712 | 1.157803 | -2.77 | 0.006 | -5.475963 | -. 9374608 |
| pa_nonInc | -1.262539 | 1.068464 | -1.18 | 0.237 | -3.356689 | . 8316115 |
| sa_age | -. 0246155 | . 0099949 | -2.46 | 0.014 | -. 0442052 | -. 0050259 |
| fc_age | . 0245569 | . 0144714 | 1.70 | 0.090 | -. 0038065 | . 0529204 |
| wc_age | . 0245614 | . 0118089 | 2.08 | 0.038 | . 0014164 | . 0477063 |
| mo_age | . 0438991 | . 0188639 | 2.33 | 0.020 | . 0069266 | . 0808716 |
| pa_age | . 0178701 | . 0238914 | 0.75 | 0.454 | -. 0289562 | . 0646964 |
| sa_single | . 7466991 | . 3018717 | 2.47 | 0.013 | . 1550414 | 1.338357 |
| fc_single | . 1833902 | . 4422272 | 0.41 | 0.678 | -. 6833593 | 1.05014 |
| wc_single | . 8986283 | . 3250911 | 2.76 | 0.006 | . 2614615 | 1.535795 |
| mo_single | . 5236316 | . 5420452 | 0.97 | 0.334 | -. 5387574 | 1.586021 |
| pa_single | . 7205517 | . 6089905 | 1.18 | 0.237 | -. 4730477 | 1.914151 |
| sa_tail | -. 4493796 | . 1176557 | -3.82 | 0.000 | -. 6799806 | -. 2187786 |
| fc_tail | -1.411387 | . 1946752 | -7.25 | 0.000 | -1.792943 | -1.029831 |
| wc_tail | -. 7021076 | . 1523444 | -4.61 | 0.000 | -1.000697 | -. 403518 |
| sa | 3.887584 | . 5398025 | 7.20 | 0.000 | 2.829591 | 4.945578 |
| fc | . 93478 | . 7386395 | 1.27 | 0.206 | -. 5129268 | 2.382487 |
| wc | 1.322184 | . 6336908 | 2.09 | 0.037 | . 0801733 | 2.564195 |
| pa | -2.786808 | 1.436408 | -1.94 | 0.052 | -5.602115 | . 0284993 |
| mo | . 7562164 | 1.036525 | 0.73 | 0.466 | -1.275335 | 2.787768 |
| SD |  |  |  |  |  |  |
| sa | 1.589231 | . 1292428 | 12.30 | 0.000 | 1.33592 | 1.842543 |
| fc | 2.018708 | . 1987104 | 10.16 | 0.000 | 1.629243 | 2.408173 |
| wc | 1.691765 | . 1591437 | 10.63 | 0.000 | 1.379849 | 2.003681 |
| pa | 3.213147 | . 4172448 | 7.70 | 0.000 | 2.395363 | 4.030932 |
| mo | 2.346026 | . 2697847 | 8.70 | 0.000 | 1.817258 | 2.874794 |

The sign of the estimated standard deviations is irrelevant: interpret them as being positive

# A.12.12. Model B. 3 - Product Attribute and Full Set of Consumer Characteristics Interaction Model (No Correlation) 

Table A.12.12.1. Model B. 3 - Product Attribute and Full Set of Consumer Characteristics Interaction Model (No Correlation)

| Mixed logit model | Number of obs | $=$ | 11380 |
| :--- | :--- | :--- | :--- |
|  | LR chi2 (5) | $=$ | 684.76 |
| Log likelihood $=-2578.0853$ | Prob $>$ chi2 | $=$ | 0.0000 |


| Y | Coef. | Std. Err | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Con | Interval] |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Mean |  |  |  |  |  |  |
| p1000 | -. 2215241 | . 0102425 | -21.63 | 0.000 | -. 2415989 | -. 2014493 |
| femaleSA | . 1523377 | .2579517 | 0.59 | 0.555 | -. 3532384 | . 6579137 |
| femaleWC | -. 3599552 | . 2855836 | -1.26 | 0.208 | -. 9196888 | .1997783 |
| femaleFC | -. 5276143 | . 3560344 | -1.48 | 0.138 | -1.225429 | .1702003 |
| femalePA | -. 2763563 | . 6935106 | -0.40 | 0.690 | -1.635612 | 1.082899 |
| femaleMO | -. 2224424 | . 4746677 | -0.47 | 0.639 | -1.152774 | . 7078892 |
| childrenSA | . 0963526 | . 2775494 | 0.35 | 0.728 | -. 4476342 | . 6403394 |
| childrenWC | -. 3353401 | . 3292031 | -1.02 | 0.308 | -. 9805664 | . 3098861 |
| childrenFC | . 3582823 | . 4253617 | 0.84 | 0.400 | -. 4754113 | 1.191976 |
| childrenMO | -. 0993182 | . 5422773 | -0.18 | 0.855 | -1.162162 | . 9635258 |
| childrenPA | -. 1393065 | . 5944195 | -0.23 | 0.815 | -1.304347 | 1.025734 |
| sa_inc3 | . 2697931 | . 330294 | 0.82 | 0.414 | -. 3775713 | . 9171576 |
| fc_inc3 | . 5062941 | . 481602 | 1.05 | 0.293 | -. 4376284 | 1.450217 |
| wc-inc3 | 1.158924 | . 394308 | 2.94 | 0.003 | . 3860946 | 1.931754 |
| mo_inc3 | -. 0908324 | . 5971402 | -0.15 | 0.879 | -1.261206 | 1.079541 |
| pa-inc3 | . 4734794 | . 8208183 | 0.58 | 0.564 | -1.135295 | 2.082254 |
| sa_nōnInc | -1.151898 | . 4286213 | -2.69 | 0.007 | -1.99198 | -. 3118151 |
| fc_nonInc | -. 7199096 | . 6016407 | -1.20 | 0.231 | -1.899104 | . 4592845 |
| wc_nonInc | -. 9616222 | . 5064621 | -1.90 | 0.058 | -1.95427 | . 0310252 |
| mo_nonInc | -3.28597 | 1.165868 | -2.82 | 0.005 | -5.571029 | -1.000911 |
| pa_nonInc | -1.431559 | 1.12145 | -1.28 | 0.202 | -3.62956 | . 7664424 |
| -sa_age | -. 0223535 | . 011377 | -1.96 | 0.049 | -. 044652 | -. 0000549 |
| fc_age | . 0263532 | . 0158416 | 1.66 | 0.096 | -. 0046957 | . 0574021 |
| wc_age | . 0305456 | . 0128128 | 2.38 | 0.017 | . 005433 | . 0556583 |
| mo_age | . 0492182 | . 0208535 | 2.36 | 0.018 | . 0083462 | .0900902 |
| pa_age | .0200493 | . 0312931 | 0.64 | 0.522 | -. 041284 | . 0813826 |
| sa_single | . 7603386 | . 3132374 | 2.43 | 0.015 | .1464047 | 1.374273 |
| fc_single | . 4164262 | . 4698776 | 0.89 | 0.375 | -. 5045171 | 1.337369 |
| wc_single | . 6787391 | . 3336888 | 2.03 | 0.042 | . 024721 | 1.332757 |
| mo_single | . 301425 | . 5871436 | 0.51 | 0.608 | -. 8493553 | 1.452205 |
| pa_single | . 6236462 | . 6621538 | 0.94 | 0.346 | -. 6741515 | 1.921444 |
| sa_educUNI | .1683996 | . 2940309 | 0.57 | 0.567 | -. 4078904 | . 7446897 |
| fc_educUNI | .1105826 | . 3894419 | 0.28 | 0.776 | -. 6527094 | . 8738747 |
| wc_educUNI | . 7045497 | . 3181663 | 2.21 | 0.027 | . 0809552 | 1.328144 |
| mo_educUNI | . 5494157 | . 5195217 | 1.06 | 0.290 | -. 4688281 | 1.56766 |
| pa_educUNI | . 3616289 | . 8779494 | 0.41 | 0.680 | -1.35912 | 2.082378 |
| sa_tail | -. 4505488 | . 1177751 | -3.83 | 0.000 | -. 6813838 | -. 2197139 |
| fc_tail | -1.415659 | . 1947666 | -7.27 | 0.000 | -1.797394 | -1.033923 |
| wc_tail | -. 6965502 | . 152476 | -4.57 | 0.000 | -. 9953976 | -. 3977028 |
| sa | 3.638391 | . 694099 | 5.24 | 0.000 | 2.277982 | 4.9988 |
| fc | . 8676914 | . 9791223 | 0.89 | 0.376 | -1.051353 | 2.786736 |
| wc | 1.224319 | . 7854852 | 1.56 | 0.119 | -. 3152033 | 2.763842 |
| pa | -2.759517 | 1.995734 | -1.38 | 0.167 | -6.671084 | 1.152051 |
| mo | . 62174 | 1.307349 | 0.48 | 0.634 | -1.940617 | 3.184097 |
| SD |  |  |  |  |  |  |
| sa | 1.59422 | . 1304183 | 12.22 | 0.000 | 1.338605 | 1.849835 |
| fc | 1.910997 | . 1842664 | 10.37 | 0.000 | 1.549842 | 2.272153 |
| wC | 1.669651 | .1554217 | 10.74 | 0.000 | 1.36503 | 1.974272 |
| pa | 3.198915 | . 4190679 | 7.63 | 0.000 | 2.377557 | 4.020273 |
| mo | 2.351558 | . 2721999 | 8.64 | 0.000 | 1.818056 | 2.88506 |

## A.12.13. Model B. 4 - Income as Dummies for all Income Groups (No Correlation)

Table A.12.13.1. Model B. 4 - Income as Dummies for all Income Groups (No Correlation)


The sign of the estimated standard deviations is irrelevant: interpret them as being positive

## A.12.14. Model B. 5 - Income as Low and High (No Correlation)

Table A.12.14.1. Model B. 5 - Income as Low and High (No Correlation)

| Mixed logit model |  |  |  | Number of obs LR chi2(5) Prob > chi2 |  | $\begin{array}{r} 11380 \\ 743.89 \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Log likelihood | -2600.92 |  |  |  |  | 0.0000 |
| y | Coef. | Std. Err | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Con | Interval] |
| Mean |  |  |  |  |  |  |
| p1000 | -. 2195481 | . 0101798 | -21.57 | 0.000 | -. 2395002 | -. 199596 |
| sa_inc_3 | -. 0216635 | . 3077729 | -0.07 | 0.944 | -. 6248873 | . 5815604 |
| fc_inc_3 | . 4074428 | . 4064897 | 1.00 | 0.316 | -. 3892624 | 1.204148 |
| wc_inc-3 | 1.244787 | . 3522805 | 3.53 | 0.000 | . 5543299 | 1.935244 |
| mo_inc_3 | . 3338106 | . 5276616 | 0.63 | 0.527 | -. 7003872 | 1.368008 |
| pa_inc_3 | . 3785514 | . 6348926 | 0.60 | 0.551 | -. 8658152 | 1.622918 |
| sa_nonInc | -1.078228 | . 4633911 | -2.33 | 0.020 | -1.986458 | -. 1699982 |
| fc_nonInc | -. 9684699 | . 6443967 | -1.50 | 0.133 | -2.231464 | . 2945245 |
| wc_nonInc | -. 7896183 | . 5018907 | -1.57 | 0.116 | -1.773306 | . 1940694 |
| mo_nonInc | -3.287969 | 1.127289 | -2.92 | 0.004 | -5.497416 | -1.078522 |
| pa_nonInc | -1.56151 | 1.109459 | -1.41 | 0.159 | -3.736009 | . 6129886 |
| sa_tail | -. 4484326 | . 1175008 | -3.82 | 0.000 | -. 67873 | -. 2181352 |
| fc_tail | -1.402456 | . 1943569 | -7.22 | 0.000 | -1.783389 | -1.021523 |
| wc_tail | -. 7007592 | . 1522095 | -4.60 | 0.000 | -. 9990843 | -. 4024341 |
| sa | 2.989869 | . 2484562 | 12.03 | 0.000 | 2.502904 | 3.476834 |
| fc | 2.170754 | . 3483924 | 6.23 | 0.000 | 1.487918 | 2.853591 |
| wc | 2.742532 | . 292987 | 9.36 | 0.000 | 2.168288 | 3.316776 |
| pa | -1.622422 | . 5547049 | -2.92 | 0.003 | -2.709624 | -. 5352208 |
| mo | 2.917453 | . 3764427 | 7.75 | 0.000 | 2.179639 | 3.655268 |
| SD |  |  |  |  |  |  |
| sa | 1.628168 | . 1347938 | 12.08 | 0.000 | 1.363977 | 1.892359 |
| fc | 2.025007 | . 1957236 | 10.35 | 0.000 | 1.641396 | 2.408618 |
| wc | 1.74891 | . 1655058 | 10.57 | 0.000 | 1.424525 | 2.073296 |
| pa | 3.32771 | . 4225854 | 7.87 | 0.000 | 2.499458 | 4.155963 |
| mo | 2.356607 | . 2856305 | 8.25 | 0.000 | 1.796782 | 2.916433 |

The sign of the estimated standard deviations is irrelevant: interpret them as being positive
A.12.15. Model B. 6 - Income and Education (No Correlation)

Table A.12.15.1. Model B. 6 - Income and Education (No Correlation)

| Mixed logit model | Number of obs | $=$ | 11380 |
| :--- | :--- | :--- | :--- |
|  | LR chi2 (5) | $=$ | 737.01 |
| Log likelihood $=-2598.7433$ | Prob $>\operatorname{chi2}$ | $=0.0000$ |  |


| Y | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Mean |  |  |  |  |  |  |
| p1000 | -. 2203148 | . 0101673 | -21.67 | 0.000 | -. 2402424 | -. 2003872 |
| sa_inc3 | . 078514 | . 2979559 | 0.26 | 0.792 | -. 5054688 | . 6624968 |
| fc_inc3 | . 3532082 | . 3768075 | 0.94 | 0.349 | -. 385321 | 1.091737 |
| wc_inc3 | 1.037666 | . 3339709 | 3.11 | 0.002 | . 3830955 | 1.692237 |
| mo_inc3 | . 4419055 | . 5391451 | 0.82 | 0.412 | -. 6147994 | 1.49861 |
| pa_inc3 | . 0100179 | . 6907666 | 0.01 | 0.988 | -1.34386 | 1.363895 |
| sa_nonInc | -1.084602 | . 4902046 | -2.21 | 0.027 | -2.045385 | -. 1238186 |
| fc_nonInc | -. 6200244 | . 5565419 | -1.11 | 0.265 | -1.710827 | . 4707778 |
| wc_nonInc | -. 8757121 | . 5059726 | -1.73 | 0.083 | -1.8674 | . 1159761 |
| mo_nonInc | -3.948041 | 1.479365 | -2.67 | 0.008 | -6.847542 | -1.048539 |
| pa_nonInc | -1.702773 | 1.422221 | -1.20 | 0.231 | -4.490276 | 1.084729 |
| sa_educUNI | . 4180674 | . 2720922 | 1.54 | 0.124 | -. 1152234 | . 9513583 |
| fc_educUNI | -. 0418893 | . 3434526 | -0.12 | 0.903 | -. 7150441 | . 6312655 |
| wc_educUNI | . 2259001 | . 3111504 | 0.73 | 0.468 | -. 3839435 | . 8357438 |
| mo_educUNI | . 5017562 | . 5174348 | 0.97 | 0.332 | -. 5123974 | 1.51591 |
| pa_educUNI | . 9041553 | . 639378 | 1.41 | 0.157 | -. 3490025 | 2.157313 |
| sa_tail | -. 436933 | . 1173321 | -3.72 | 0.000 | -. 6668996 | -. 2069663 |
| fc_tail | -1.413537 | . 193639 | -7.30 | 0.000 | -1.793063 | -1.034012 |
| wc_tail | -. 6967348 | . 152745 | -4.56 | 0.000 | -. 9961094 | -. 3973601 |
| sa | 2.774276 | . 2551275 | 10.87 | 0.000 | 2.274236 | 3.274317 |
| fc | 2.172743 | . 3408331 | 6.37 | 0.000 | 1.504723 | 2.840764 |
| wc | 2.639184 | . 3202668 | 8.24 | 0.000 | 2.011473 | 3.266896 |
| mo | 2.777158 | . 4372442 | 6.35 | 0.000 | 1.920175 | 3.63414 |
| pa | -2.058678 | . 6883157 | -2.99 | 0.003 | -3.407752 | -. 7096038 |
| SD |  |  |  |  |  |  |
| sa | 1.619028 | . 1290835 | 12.54 | 0.000 | 1.366029 | 1.872027 |
| fc | 1.752635 | . 1538991 | 11.39 | 0.000 | 1.450998 | 2.054272 |
| wC | 1.733166 | . 1456031 | 11.90 | 0.000 | 1.447789 | 2.018543 |
| mo | 2.646609 | . 3470805 | 7.63 | 0.000 | 1.966344 | 3.326874 |
| pa | 3.50803 | . 4823208 | 7.27 | 0.000 | 2.562698 | 4.453361 |

## A.12.16. Model B. 7 - Income and Age (No Correlation)

Table A.12.16.1. Model B. 7 - Income and Age (No Correlation)

| Mixed logit model <br> Log likelihood $=-2591.8197$ |  |  |  | Number of obs LR chi2(5) Prob > chi2 |  | $\begin{array}{r} 11380 \\ 714.00 \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | 0.0000 |
| y \| | Coef. | Std. Err. | z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% Con | Interval] |
| Mean |  |  |  |  |  |  |
| p1000 \| | -. 2201406 | . 0101888 | -21.61 | 0.000 | -. 2401103 | -. 200171 |
| sa_inc3 \| | . 0522675 | . 3068436 | 0.17 | 0.865 | -. 5491349 | . 65367 |
| fc_inc3 \| | . 3011282 | . 4372212 | 0.69 | 0.491 | -. 5558097 | 1.158066 |
| wc_inc3 \| | 1.110882 | . 3598264 | 3.09 | 0.002 | . 4056354 | 1.816129 |
| mo_inc3 \| | -. 0347357 | . 5346324 | -0.06 | 0.948 | -1.082596 | 1.013125 |
| pa_inc3 \| | . 299782 | . 7123258 | 0.42 | 0.674 | -1.096351 | 1.695915 |
| sa_nonInc \| | -1.096145 | . 4471846 | -2.45 | 0.014 | -1.97261 | -. 219679 |
| fc_nonInc \| | -. 9388788 | . 6151484 | -1.53 | 0.127 | -2.144548 | . 26679 |
| wc_nonInc \| | -. 8862412 | . 49811 | -1.78 | 0.075 | -1.862519 | . 0900364 |
| mo_nonInc \| | -3.16453 | 1.139066 | -2.78 | 0.005 | -5.397059 | -. 9320013 |
| pa_nonInc \| | -1.464951 | 1.081396 | -1.35 | 0.176 | -3.584448 | . 654545 |
| sa_age \| | -. 0258103 | . 0105043 | -2.46 | 0.014 | -. 0463983 | -. 0052223 |
| fc_age \| | . 0165338 | . 0153313 | 1.08 | 0.281 | -. 013515 | . 0465826 |
| wc_age \| | . 0253774 | . 0130292 | 1.95 | 0.051 | -. 0001594 | . 0509141 |
| mo_age \| | . 0399938 | . 0185816 | 2.15 | 0.031 | . 0035746 | . 076413 |
| pa_age \| | . 0094541 | . 0313285 | 0.30 | 0.763 | -. 0519486 | . 0708567 |
| sa_tail \| | -. 446634 | . 1174765 | -3.80 | 0.000 | -. 6768836 | -. 2163844 |
| fc_tail \| | -1.403053 | . 1942859 | -7.22 | 0.000 | -1.783846 | -1.022259 |
| wc_tail \| | -. 7017746 | . 1522861 | -4.61 | 0.000 | -1.00025 | -. 4032994 |
| sa \| | 4.173266 | . 5559755 | 7.51 | 0.000 | 3.083574 | 5.262958 |
| fc \| | 1.41413 | . 7622067 | 1.86 | 0.064 | -. 0797672 | 2.908028 |
| wc \| | 1.62943 | . 6926236 | 2.35 | 0.019 | . 2719126 | 2.986947 |
| pa \| | -2.109756 | 1.866338 | -1.13 | 0.258 | -5.767711 | 1.548199 |
| mo \| | 1.124268 | . 9691943 | 1.16 | 0.246 | -. 7753182 | 3.023854 |
| SD |  |  |  |  |  |  |
| sa \| | 1.605394 | . 1303048 | 12.32 | 0.000 | 1.350002 | 1.860787 |
| fc \| | 1.989765 | . 1971604 | 10.09 | 0.000 | 1.603338 | 2.376193 |
| wc \| | 1.685913 | . 1635229 | 10.31 | 0.000 | 1.365414 | 2.006412 |
| pa \| | 3.295532 | . 4185922 | 7.87 | 0.000 | 2.475106 | 4.115957 |
| mo \| | 2.335818 | . 2788206 | 8.38 | 0.000 | 1.78934 | 2.882297 |
| The sign of the estimated standard deviations is irrelevant: interpret them as being positive |  |  |  |  |  |  |

## A.12.17. Model B. 8 - Income, Age, Single and Female

Table A.12.17.1. Model B. 8 - Income, Age, Single and Female

| Mixed logit model |  |  |  | Number of obs LR chi2(5) |  |  | $\begin{array}{r} 11380 \\ 692.37 \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Log likelihood | -2582.375 |  |  |  |  |  | 0.0000 |
| Y I | Coef. | Std. Err | Z | $\mathrm{P}>\|\mathrm{z}\|$ | [95\% |  | Interval] |
| Mean |  |  |  |  |  |  |  |
| p1000 | -. 2208827 | . 0102114 | -21.63 | 0.000 | -. 2408 |  | -. 2008688 |
| femaleSA \| | . 1632118 | . 2572894 | 0.63 | 0.526 | -. 3410 |  | . 6674898 |
| femaleWC \| | -. 3971561 | . 2888523 | -1.37 | 0.169 | -. 9632 |  | . 168984 |
| femaleFC | -. 4653565 | . 3897295 | -1.19 | 0.232 | -1.229 |  | . 2984992 |
| femalePA | -. 2499069 | . 6455539 | -0.39 | 0.699 | -1.515 |  | 1.015355 |
| femaleMO | -. 2575745 | . 4706646 | -0.55 | 0.584 | -1.18 |  | . 6649111 |
| sa_inc3 | . 3075114 | . 3154067 | 0.97 | 0.330 | -. 3106 |  | . 9256971 |
| fc_inc3 | . 3720397 | .451242 | 0.82 | 0.410 | -. 5123 |  | 1.256458 |
| wc_inc3 | 1.387898 | . 3885756 | 3.57 | 0.000 | . 6263 |  | 2.149492 |
| mo_inc3 | . 1108292 | . 5723893 | 0.19 | 0.846 | -1.011 |  | 1.232692 |
| pa_inc3 \| | . 6517143 | . 6598894 | 0.99 | 0.323 | -. 6416 |  | 1.945074 |
| sa_nonInc | -1.138633 | . 4229212 | -2.69 | 0.007 | -1.967 |  | -. 3097227 |
| fc_nonInc \| | -. 8088386 | . 6066503 | -1.33 | 0.182 | -1.997 |  | . 3801741 |
| wc_nonInc \| | -. 756146 | . 5009442 | -1.51 | 0.131 | -1.737 |  | . 2256866 |
| mo_nonInc \| | -3.32061 | 1.147699 | -2.89 | 0.004 | -5.57 |  | -1.071161 |
| pa_nonInc \| | -1.261126 | 1.07545 | -1.17 | 0.241 | -3.368 |  | . 8467169 |
| sa_age | -. 0232244 | . 0101892 | -2.28 | 0.023 | -. 0431 |  | -. 003254 |
| fc_age \| | . 0194326 | . 0158936 | 1.22 | 0.221 | -. 0117 |  | .0505834 |
| wc_age \| | . 0235107 | .0118653 | 1.98 | 0.048 | . 0002 |  | .0467662 |
| mo_age \| | . 0415803 | . 0188602 | 2.20 | 0.027 | . 0046 |  | . 0785456 |
| pa_age \| | . 0164734 | . 0273374 | 0.60 | 0.547 | -. 037 |  | . 0700538 |
| sa_single | . 7574746 | . 2957595 | 2.56 | 0.010 | . 1777 |  | 1.337153 |
| fc_single \| | . 2346478 | . 4454963 | 0.53 | 0.598 | -. 6385 |  | 1.107804 |
| wc_single | . 8614223 | . 3248169 | 2.65 | 0.008 | . 2247 |  | 1.498052 |
| mo_single \| | . 4750061 | . 5400075 | 0.88 | 0.379 | -. 5833 |  | 1.533401 |
| pa_single \| | .7251331 | . 6278271 | 1.15 | 0.248 | -. 5053 |  | 1.955652 |
| sa_tail \| | -. 4488372 | . 1176786 | -3.81 | 0.000 | -. 679 |  | -. 2181915 |
| fc_tail \| | -1.408806 | .1945507 | -7.24 | 0.000 | -1.790 |  | -1.027494 |
| wc_tail \| | -. 7013171 | . 1522455 | -4.61 | 0.000 | -. 9997 |  | -. 4029213 |
| sa \| | 3.738725 | . 5824984 | 6.42 | 0.000 | 2.597 |  | 4.8804 |
| fc \| | 1.401539 | .870079 | 1.61 | 0.107 | -. 3037 |  | 3.106862 |
| wc \| | 1.597881 | . 6736916 | 2.37 | 0.018 | . 27 |  | 2.918293 |
| pa \| | -2.62466 | 1.739215 | -1.51 | 0.131 | -6.033 |  | . 7841384 |
| mo \| | 1.030966 | 1.104598 | 0.93 | 0.351 | -1.134 |  | 3.195938 |
| SD \| |  |  |  |  |  |  |  |
| sa \| | 1.599647 | .130603 | 12.25 | 0.000 | 1.34 |  | 1.855624 |
| fc \| | 1.945857 | . 1964055 | 9.91 | 0.000 | 1.560 |  | 2.330805 |
| wc | 1.662154 | . 1562513 | 10.64 | 0.000 | 1.355 |  | 1.968401 |
| pa \| | 3.198289 | . 4170934 | 7.67 | 0.000 | 2.380 |  | 4.015778 |
| mo \| | 2.365593 | .2710986 | 8.73 | 0.000 | 1.83 |  | 2.896937 |
| The sign of the estimated standard deviations is irrelevant: interpret them as being positive |  |  |  |  |  |  |  |

## A.12.18. Model B. 9 - Income, Age, Single, Female and Children

Table A.12.18.1. Model B. 9 - Income, Age, Single, Female and Children


## A.13. STATA Do-Files

Below follow all the STATA Do-Files. They follow in the same order as the regression outputs did. Note that the Do-files for the regression outputs that did not allow for correlations in unobserved factors over choice scenarios are not included. They are exactly equal to the Do-files below except that the option "corr" is missing in the mixlogit command line.

## A.13.1. STATA Do-File Model 1 - Product Attribute Model

Table 13.1.1. STATA Do-File Model 1 - Product Attribute Model

```
clear
use H:\Masteroppgave\RCData\RCdata.dta
* We create a new variable for group where we combine id and set to idset -
group(idset)
gen idset = 100*id+set
******************** fish ************************
* Generate dummy variable for children. 1 = family have children under 18, 0
otherwise:
gen children = 0
replace children = 1 if q18_2 > 0
* Generate tail interactions with sa, fc and wc
gen sa_tail = sa*tail
gen fc_tail = fc*tail
gen wc_tail = wc*tail
* Four observations have missing values. We drop those:
drop if choice == 9
* Generate global random variables:
global randvars "sa fc wc pa mo"
* Define mixlogit model:
mixlogit y pl000 sa_tail fc_tail wc_tail, rand($randvars) group(idset) id(id)
nrep(500) corr
estimates store fish
test sa = wc
test sa = mo
test sa = pa
test wc = mo
test wc = pa
test mo = pa
test sa_tail = fc_tail
test sa tail = wc tail
test fc-tail = wc-
test sa + sa_tail = fc + fc_tail
test sa + sa tail = wc + wc tail
test fc + fc_tail = wc + wc_tail
mixlcov
mixlcov, sd
wtp p1000 sa fc wc mo pa sa tail fc tail wc tail
*************************** fish SLUTT ****若***********************
```


## A.13.2 STATA Do-File Model 2 - Product Attribute and Consumer Characteristics Interaction Model

Table A.13.2.1 STATA Do-File Model 2 - Product Attribute and Consumer Characteristics Interaction Model

```
clear
use H:\Masteroppgave\RCData\RCdata.dta
* We create a new variable for group where we combine id and set to idset -
group(idset)
gen idset = 100*id+set
******************** fish_inc3000_age_single **************************
* Generate dummy for "do no want to answer":
replace q20 = 0 if q20==4
replace q20 = 0 if q20==.
gen nonInc = 0
replace nonInc = 1 if q20 == 0
* Generate dummy for income group 3:
gen inc_3 = 0
replace inc_3 = 1 if q20 == 3
* Generate income variables for do not want to answer
gen sa_nonInc = sa*nonInc
gen fc nonInc = fc*nonInc
gen wc_nonInc = wc*nonInc
gen mo_nonInc = mo*nonInc
gen pa_nonInc = pa*nonInc
* Generate income dummy for income group 3
gen sa_inc3 = sa*inc_3
gen fc inc3 = fc*inc 3
gen wc_inc3 = wc*inc_3
gen mo inc3 = mo*inc 3
gen pa_inc3 = pa*inc_3
* Generate age variables
gen sa age = sa*age
gen fc_age = fc*age
gen mo_age = mo*age
gen wc_age = wc*age
gen pa_age = pa*age
* Generate single variable
gen single = 0
replace single = 1 if d3 == 1
gen sa_single = sa*single
gen fc_single = fc*single
gen mo_single = mo*single
gen wc_single = wc*single
gen pa_single = pa*single
* Generate tail interactions with sa, fc and wc
gen sa tail = sa*tail
gen fc_tail = fc*tail
gen wc_tail = wc*tail
* Four observations have missing values. We drop those:
drop if choice == 9
* Generate global random variables:
global randvars "sa fc wc pa mo"
```

```
* Define mixlogit model:
mixlogit y p1000 sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3 sa_nonInc fc_nonInc
wc nonInc mo nonInc pa nonInc sa age fc age wc age mo age pa age sa single
fc_single wc_single mo_single pa_single sa_tail fc_tail wc_tail, rand($randvars)
group(idset) id(id) nrep(500) corr
estimates store Model2_WTP
******************** TEST OVERALL SIGNIFICACE *****************
test sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3
test sa nonInc fc nonInc wc nonInc mo nonInc pa nonInc
test sa_age fc_age wc_age mo_age pa_age
test sa_single fc_single wc_single mo_single pa_single
*******\overline{*}*********\overline{*}*********\overline{*}**** SINGLE TESTS *\overline{*}**********************
test sa = wc
test sa = mo
test sa = pa
test wc = mo
test wc = pa
test mo = pa
**************************** INCOME TESTS ****************************
test sa+sa inc3 = fc+fc inc3
test sa+sa_inc3 = wc+wc_inc3
test sa+sa_inc3 = mo+mo_inc3
test sa+sa inc3 = pa+pa inc3
test fc+fc_inc3 = wc+wc_inc3
test fc+fc inc3 = mo+mo inc3
test fc+fc_inc3 = pa+pa_inc3
test wc+wc inc3 = mo+mo inc3
test wc+wc_inc3 = pa+pa_inc3
test mo+mo inc3 = pa+pa inc3
**********\overline{*}************\overline{*}***** INCOME AGE TESTS ************************
test sa+sa inc3+sa age = fc+fc inc3+fc age
test sa+sa-inc3+sa age = wc+wc inc3+wc age
test sa+sa_inc3+sa_age = mo+mo_inc3+mo_age
test sa+sa inc3+sa age = pa+pa inc3+pa age
test fc+fc_inc3+fc_age = wc+wc_inc3+wc_age
test fc+fc inc3+fc age = mo+mo inc3+mo age
test fc+fc_inc3+fc_age = pa+pa_inc3+pa_age
test wc+wc inc3+wc age = mo+mo inc3+mo age
test wc+wc_inc3+wc_age = pa+pa_inc3+pa_age
test mo+mo inc3+mo age = pa+pa inc3+pa age
**********\overline{*}*******\overline{*}********** \overline{INCOME S INGLE TESTS *************************}
test sa+sa inc3+sa single = fc+fc inc3+fc single
test sa+sa-inc3+sa-single = wc+wc_inc3+wc-single
test sa+sa_inc3+sa_single = mo+mo_inc3+mo_single
test sa+sa inc3+sa single = pa+pa inc3+pa single
test fc+fc_inc3+fc_single = wc+wc_inc3+wc_single
test fc+fc inc3+fc single = mo+mo inc3+mo single
test fc+fc_inc3+fc_single = pa+pa_inc3+pa_single
test wc+wc inc3+wc single = mo+mo inc3+mo single
test wc+wc_inc3+wc_single = pa+pa_inc3+pa_single
test mo+mo_inc3+mo_single = pa+pa_inc3+pa_single
********************************* TNCOME AGE SINGTE TESTS *******************
test sa+sa_inc3+sa_age+sa_single = fc+fc_inc3+fc_age+fc_single
test sa+sa_inc3+sa_age+sa_single = wc+wc_inc3+wc_age+wc_single
test sa+sa inc3+sa age+sa single = mo+mo inc3+mo age+mo single
test sa+sa_inc3+sa_age+sa_single = pa+pa_inc3+pa_age+pa_single
test fc+fc inc3+fc age+fc single = wc+wc inc3+wc age+wc single
test fc+fc_inc3+fc__age+fc_single = mo+mo_inc3+mo_age+mo_single
test fc+fc inc3+fc age+fc single = pa+pa inc3+pa age+pa single
test wc+wc_inc3+wc_age+wc_single = mo+mo_inc3+mo_age+mo_single
test wc+wc_inc3+wc_age+wc_single = pa+pa_inc3+pa_age+pa_single
test mo+mo-inc3+mo_age+mo_single = pa+pa-inc3+pa-age+pa-single
**************** ESTMATE WILLINGNESS TO PAY *****************************
wtp p1000 sa fc wc mo pa sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3 sa_nonInc
fc nonInc wc nonInc mo nonInc pa nonInc sa age fc age wc age mo age pa age
```

sa_single fc_single wc_single mo_single pa_single sa_tail fc_tail wc_tail
 wc_age mo_age pa_age sa_single fc_single wc_single mo_single pa_single sa_tail fc_tail wc_tail

* nlcom
$\left(\left(\_b[s a]+\_b[f c]+\_b[w c]+\_b[m o]+\_b[p a]+\_b\left[s a \_i n c 3\right]+\_b\left[f c \_i n c 3\right]+\_b\left[w c \_i n c 3\right]+\_b\left[m o \_i n c 3\right.\right.\right.$

 +_b[wc_singlē]+_b[mo_single]+_b[pa_single]+_b[sa_tail]+_b[fc_tail]+_b[wc_tail])/(_b [ $\overline{\mathrm{p}} 1000 \mathrm{]}$ ))
mixlcov
mixlcov, sd

$\star * * * * * * * * * * * * * * * * * * * * * * * * * * ~$


## A.13.3. STATA Do-File Model 3 - Product Attribute and Full Set of Consumer Characteristics Interaction Model

Table A.13.3.1. STATA Do-File Model 3 - Product Attribute and Full Set of Consumer Characteristics Interaction Model

```
clear
use H:\Masteroppgave\RCData\RCdata.dta
* We create a new variable for group where we combine id and set to idset -
group(idset)
gen idset = 100*id+set
******************** fish_inc3000_female_children_age_single_educDummy
************************
* Generate dummy variable for children. 1 = family have children udder 18, 0
otherwise:
gen children = 0
replace children = 1 if q18_2 > 0
* Generate dummy for "do no want to answer":
replace q20 = 0 if q20==4
replace q20 = 0 if q20==.
gen nonInc = 0
replace nonInc = 1 if q20 == 0
* Generate dummy for income group 3:
gen inc 3 = 0
replace inc_3 = 1 if q20 == 3
* Generate income variables for do not want to answer
gen sa_nonInc = sa*nonInc
gen fc_nonInc = fc*nonInc
gen wc_nonInc = wc*nonInc
gen mo nonInc = mo*nonInc
gen pa_nonInc = pa*nonInc
* Generate income dummy for income group 3
gen sa_inc3 = sa*inc_3
gen fccinc3 = fc*inc-3
gen wc_inc3 = wc*inc_3
gen mo_inc3 = mo*inc_3
gen pa_inc3 = pa*inc_3
* Generate children variable
gen childrenSA = children*sa
gen childrenWC = children*Wc
gen childrenFC = children*fc
gen childrenPA = children*pa
gen childrenMO = children*mo
* Generate gender variables
gen femaleSA = female*sa
gen femaleWC = female*wc
gen femaleFC = female*fc
gen femalePA = female*pa
gen femaleMO = female*mo
* Generate age variables
gen sa_age = sa*age
gen fc_age = fc*age
gen mo_age = mo*age
gen wc_age = wc*age
gen pa_age = pa*age
* Generate single variable
gen single = 0
```

```
replace single = 1 if d3 == 1
gen sa_single = sa*single
gen fc_single = fc*single
gen mo_single = mo*single
gen wc_single = wc*single
gen pa_single = pa*single
* Generate education dummy for university degree :
gen educUNI = 0
replace educUNI = 1 if d7 > 4
gen sa_educUNI = sa*educUNI
gen fc_educUNI = fc*educUNI
gen mo_educUNI = mo*educUNI
gen wc_educUNI = wc*educUNI
gen pa_educUNI = pa*educUNI
* Generate tail interactions with sa, fc and wc
gen sa_tail = sa*tail
gen fc_tail = fc*tail
gen wc_tail = wc*tail
* Four observations have missing values. We drop those:
drop if choice == 9
* Generate global random variables:
global randvars "sa fc wc pa mo"
* Define mixlogit model:
mixlogit y p1000 femaleSA femaleWC femaleFC femalePA femaleMO childrenSA childrenWC
childrenFC childrenMO childrenPA sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3 sa_nonInc
```



```
sa_single fc__single wc_single mo_single pa_single-sa_edu\overline{c}UNI fc_e_educUNI wc_educUNI
mo_educUNI p\overline{a_educUNI sa_tail fc_tail wc_tāil, rand($randvars) g}\textrm{g}roup(idset) id(id)
nrep(500) corr
estimates store fish inc3000 female children age single
test femaleSA femaleWC femaleFC femalePA femaleMO
test childrenSA childrenWC childrenFC childrenMO childrenPA
test sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3
test sa_nonInc f`c_nonInc\overline{co_nonInc mo_nōnInc pa_nonInc}
test sa_age fc_agè wc_age mō_age pa_ag}
test sa_single fc_single wc_single mo_single pa_single
test sa_educUNI f
**************************** fish_inc3000_female_children_age_single_educDummy SLUTT
***************************
```


## A.13.4. STATA Do-File Model 4 - Income as Dummies for all Income Groups

Table A.13.4.1. STATA Do-File Model 4 - Income as Dummies for all Income Groups

```
clear
use H:\Masteroppgave\RCData\RCdata.dta
* We create a new variable for group where we combine id and set to idset -
group(idset)
gen idset = 100*id+set
************************** fish inc dummy **********************************
* Generate dummy for "do no want to answer":
replace q20 = 0 if q20==4
replace q20 = 0 if q20==.
gen nonInc = 0
replace nonInc = 1 if q20 == 0
* Generate dummy for income group 1:
gen inc_1 = 0
replace inc_1 = 1 if q20 == 1
*Generate dummy for income group 2:
gen inc_2 = 0
replace inc_2 = 1 if q20 == 2
*Generate dummy for income group 3:
gen inc_3 = 0
replace inc_3 = 1 if q20 == 3
* Generate income variablesgen sa_no_answ = sa*no_answ
gen sa_nonInc = sa*nonInc
gen fc_nonInc = fc*nonInc
gen wc_nonInc = wc*nonInc
gen mo_nonInc = mo*nonInc
gen pa_nonInc = pa*nonInc
* Generate income dummy for group 1
*gen sa_inc_1 = sa*inc_1
*gen fc_inc-1 = fc*inc-1
*gen wc_inc_1 = wc*inc_1
*gen mo_inc_1 = mo*inc_1
*gen pa_inc_1 = pa*inc_1
* Generate income dummy for group 2
gen sa_inc_2 = sa*inc_2
gen fc_inc-2 = fc*inc-2
gen wc_inc_2 = wc*inc_2
gen mo_inc_2 = mo*inc- 2
gen pa_inc_2 = pa*inc_2
* Generate income dummy for group 3
gen sa_inc_3 = sa*inc_3
gen fc-inc-3 = fc*inc-3
gen wc_inc_3 = wc*inc_3
gen mo-inc_3 = mo*inc_3
gen pa_inc_3 = pa*inc_3
* Generate tail interactions with sa, fc and wc
gen sa_tail = sa*tail
gen fc_tail = fc*tail
gen wc-tail = wc*tail
* Four observations have missing values. We drop those:
drop if choice == 9
* Generate global random variables:
```

global randvars "sa fc wc pa mo"

* Define mixlogit model:
mixlogit y plo00 sa_inc_2 fc_inc_2 wc_inc_2 mo_inc_2 pa_inc_2 sa_inc_3 fc_inc_3


estimates store fish_inc_dummy
test sa inc 2 fc inc ${ }^{-} 2 \mathrm{w} \overline{\mathrm{C}}$ inc 2 mo inc 2 pa inc 2

test sa_nonInc fc_nonInc wc_nonInc mo_nonInc pa_nonInc



## A.13.5. STATA Do-File Model 5 - Income as Low and High

## Table A.13.5.1. STATA Do-File Model 5 - Income as Low and High

```
clear
use H:\Masteroppgave\RCData\RCdata.dta
* We create a new variable for group where we combine id and set to idset -
group(idset)
gen idset = 100*id+set
************************** fish inc dummy 3000EUR
* Generate dummy for "do not want to answer":
replace q20 = 0 if q20==4
replace q20 = 0 if q20==.
gen nonInc = 0
replace nonInc = 1 if q20 == 0
* Generate dummy for income group :
gen inc_3 = 0
replace inc 3 = 1 if q20 == 3
* Generate income variablesgen sa_no_answ = sa*no_answ
gen sa_nonInc = sa*nonInc
gen fc nonInc = fc*nonInc
gen wc_nonInc = wc*nonInc
gen mo_nonInc = mo*nonInc
gen pa_nonInc = pa*nonInc
* Generate income dummy for group 3
gen sa inc 3 = sa*inc 3
gen fc_inc_3 = fc*inc_3
gen wc inc 3 = wc*inc 3
gen mo_inc_3 = mo*inc_3
gen pa_inc_3 = pa*inc 3
* Generate tail interactions with sa, fc and wc
gen sa_tail = sa*tail
gen fc_tail = fc*tail
gen wc tail = wc*tail
* Four observations have missing values. We drop those:
drop if choice == 9
* Generate global random variables:
global randvars "sa fc wc pa mo"
* Define mixlogit model:
mixlogit y p1000 sa_inc_3 fc_inc_3 wc_inc_3 mo_inc_3 pa_inc_3 sa_nonInc fc_nonInc
wc_nonInc mo_nonInc pa_nonInc sa_tail fc_tail wc_tail, rand($randvars) group(idset)
id(id) nrep(500) corr
estimates store fish_inc_dummy_3000EUR
test sa_inc_3 fc_inc_3 wc_inc_3 mo_inc_3 pa_inc_3
```



```
*************************** fish inc SLUTT ****************************
```


## A.13.6. STATA Do-File Model 6 - Income and Education

Table A.13.6.1. STATA Do-File Model 6 - Income and Education

```
clear
use H:\Masteroppgave\RCData\RCdata.dta
* We create a new variable for group where we combine id and set to idset -
group(idset)
gen idset = 100*id+set
******************** fish inc3000 educDummy ************************
* Generate dummy variable for children. 1 = family have children under 18, 0
otherwise:
gen children = 0
replace children = 1 if q18_2 > 0
* Generate dummy for "do not want to answer":
replace q20 = 0 if q20==4
replace q20 = 0 if q20==.
gen nonInc = 0
replace nonInc = 1 if q20 == 0
* Generate dummy for income group 3:
gen inc_3 = 0
replace inc_3 = 1 if q20 == 3
* Generate income variables for do not want to answer
gen sa_nonInc = sa*nonInc
gen fc_nonInc = fc*nonInc
gen wc_nonInc = wc*nonInc
gen mo nonInc = mo*nonInc
gen pa_nonInc = pa*nonInc
* Generate income dummy for income group 3
gen sa_inc3 = sa*inc_3
gen fc_inc3 = fc*inc_3
gen wc_inc3 = wc*inc_3
gen mo_inc3 = mo*inc_3
gen pa_inc3 = pa*inc_3
* Generate education dummy for university degree :
gen educUNI = 0
replace educUNI = 1 if d7 > 4
gen sa_educUNI = sa*educUNI
gen fc_educUNI = fc*educUNI
gen mo_educUNI = mo*educUNI
gen wc_educUNI = wc*educUNI
gen pa_educUNI = pa*educUNI
* Generate tail interactions with sa, fc and wc
gen sa_tail = sa*tail
gen fcctail = fc*tail
gen wc_tail = wc*tail
* Four observations have missing values. We drop those:
drop if choice == 9
* Generate global random variables:
global randvars "sa fc wc pa mo"
* Define mixlogit model:
mixlogit y p1000 sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3 sa_nonInc fc_nonInc
wc_nonInc mo_nonInc pa_nonInc sa_educUNI fc_educUNI wc_educUNI mo_educUNI
pa_educUNI sā_tail fc__tail wc_ta\overline{i}, rand($rāndvars) grōup(idset) i
corr
estimates store fish inc3000 female_children age single
```

```
test sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3
test sa_nonInc \overline{fc_nonInc}\mp@subsup{\overline{c}}{~}{wc_nonInc mo_nōnInc pa_nonInc}
test sa_educUNI fc_educUNI wc_educUNI mo_educUNI pa_educUNI
*************************** fish_inc3000_educDummy SLUTT
***************************
```


## A.13.7. STATA Do-File Model 7 - Income and Age

Table A.13.7.1. STATA Do-File Model 7 - Income and Age

```
clear
use H:\Masteroppgave\RCData\RCdata.dta
* We create a new variable for group where we combine id and set to idset -
group(idset)
gen idset = 100*id+set
******************** fish inc3000 age *************************
* Generate dummy variable for children. 1 = family have children under 18, 0
otherwise:
gen children = 0
replace children = 1 if q18_2 > 0
* Generate dummy for "do not want to answer":
replace q20 = 0 if q20==4
replace q20 = 0 if q20==.
gen nonInc = 0
replace nonInc = 1 if q20 == 0
* Generate dummy for income group 3:
gen inc_3 = 0
replace inc_3 = 1 if q20 == 3
* Generate income variables for do not want to answer
gen sa_nonInc = sa*nonInc
gen fc_nonInc = fc*nonInc
gen wc_nonInc = wc*nonInc
gen mo nonInc = mo*nonInc
gen pa_nonInc = pa*nonInc
* Generate income dummy for income group 3
gen sa_inc3 = sa*inc_3
gen fc_inc3 = fc*inc_3
gen wc_inc3 = wc*inc_3
gen mo-inc3 = mo*inc_3
gen pa_inc3 = pa*inc_3
* Generate age variables
gen sa_age = sa*age
gen fc_age = fc*age
gen mo_age = mo*age
gen wc_age = wc*age
gen pa_age = pa*age
* Generate tail interactions with sa, fc and wc
gen sa_tail = sa*tail
gen fc_tail = fc*tail
gen wc_tail = wc*tail
* Four observations have missing values. We drop those:
drop if choice == 9
* Generate global random variables:
global randvars "sa fc wc pa mo"
* Define mixlogit model:
mixlogit y p1000 sa inc3 fc inc3 wc inc3 mo inc3 pa inc3 sa nonInc fc nonInc
wc_nonInc mo_nonInc pa_nonIñ sa_age\overline{e}fc_age- wc_age mo_age pà_age sa_sinngle sa_tail
```



```
estimates store fish_inc3000_female_children_age_single
test sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3
test sa nonInc \overline{fc nonInc}\mathrm{ wc nonInc mo nōnInc pa nonInc}
```

[^7]
## A.13.8. STATA Do-File Model $\boldsymbol{8}$ - Income, Age, Single and Female

Table A.13.8.1. STATA Do-File Model 8 - Income, Age, Single and Female

```
clear
use H:\Masteroppgave\RCData\RCdata.dta
* We create a new variable for group where we combine id and set to idset -
group(idset)
gen idset = 100*id+set
******************** fish_inc3000_female_age_single ************************
* Generate dummy variable for children. \overline{1 = family have children under 18, 0}
otherwise:
gen children = 0
replace children = 1 if q18_2 > 0
* Generate dummy for "do not want to answer":
replace q20 = 0 if q20==4
replace q20 = 0 if q20==.
gen nonInc = 0
replace nonInc = 1 if q20 == 0
* Generate dummy for income group 3:
gen inc_3 = 0
replace inc_3 = 1 if q20 == 3
* Generate income variables for do not want to answer
gen sa_nonInc = sa*nonInc
gen fc-nonInc = fc*nonInc
gen wc_nonInc = wc*nonInc
gen mo_nonInc = mo*nonInc
gen pa_nonInc = pa*nonInc
* Generate income dummy for income group 3
gen sa_inc3 = sa*inc_3
gen fc_inc3 = fc*inc_3
gen wc_inc3 = wc*inc_3
gen mo_inc3 = mo*inc_3
gen pa_inc3 = pa*inc_3
* Generate gender variables
gen femaleSA = female*sa
gen femaleWC = female*wc
gen femaleFC = female*fc
gen femalePA = female*pa
gen femaleMO = female*mo
* Generate age variables
gen sa_age = sa*age
gen fc_age = fc*age
gen mo_age = mo*age
gen wc_age = wc*age
gen pa_age = pa*age
* Generate single variable
gen single = 0
replace single = 1 if d3 == 1
gen sa_single = sa*single
gen fc_single = fc*single
gen mo_single = mo*single
gen wc_single = wc*single
gen pa_single = pa*single
* Generate tail interactions with sa, fc and wc
gen sa_tail = sa*tail
```

```
gen fc tail = fc*tail
gen wc_tail = wc*tail
* Four observations have missing values. We drop those:
drop if choice == 9
* Generate global random variables:
global randvars "sa fc wc pa mo"
* Define mixlogit model:
mixlogit y p1000 femaleSA femaleWC femaleFC femalePA femaleMO sa_inc3 fc_inc3
wc_inc3 mo_inc3 pa_inc3 sa_nonInc fc_nonInc wc_nonInc mo_nonInc pa_nonInc sa_age
```



```
sa_tail fc_tail wc_tail, rand($randvars) group(idset) id(id) nrep(500) corr
estimates store fi\overline{sh_inc3000_female_children_age_single}
test femaleSA female\overline{WC femaleFC femalePA femaleMO}
test sa inc3 fc inc3 wc inc3 mo inc3 pa inc3
test sa_nonInc f
test sa_age fc_age wc_age mo_age pa_age
test sa_single fc_sing
*************************** fish_inc3000_female_age_single SLUTT
***************************
```


## A.13.9. STATA Do-File Model 9 - Income, Age, Single, Female and Children

Table A.13.9.1. STATA Do-File Model 9 - Income, Age, Single, Female and Children

```
clear
use H:\Masteroppgave\RCData\RCdata.dta
* We create a new variable for group where we combine id and set to idset -
group(idset)
gen idset = 100*id+set
******************** fish_inc3000_female_children_age_single educDummy
************************
* Generate dummy variable for children. 1 = family have children under 18, 0
otherwise:
gen children = 0
replace children = 1 if q18_2 > 0
* Generate dummy for "do not want to answer":
replace q20 = 0 if q20==4
replace q20 = 0 if q20==.
gen nonInc = 0
replace nonInc = 1 if q20 == 0
* Generate dummy for income group 3:
gen inc_3 = 0
replace inc_3 = 1 if q20 == 3
* Generate income variables for do not want to answer
gen sa_nonInc = sa*nonInc
gen fc_nonInc = fc*nonInc
gen wc nonInc = wc*nonInc
gen mo_nonInc = mo*nonInc
gen pa_nonInc = pa*nonInc
* Generate income dummy for income group 3
gen sa_inc3 = sa*inc_3
gen fc_inc3 = fc*inc_3
gen wc-inc3 = wc*inc-3
gen mo_inc3 = mo*inc_3
gen pa_inc3 = pa*inc_3
* Generate children variable
gen childrenSA = children*sa
gen childrenWC = children*wc
gen childrenFC = children*fc
gen childrenPA = children*pa
gen childrenMO = children*mo
* Generate gender variables
gen femaleSA = female*sa
gen femaleWC = female*WC
gen femaleFC = female*fc
gen femalePA = female*pa
gen femaleMO = female*mo
* Generate age variables
gen sa_age = sa*age
gen fc_age = fc*age
gen mo_age = mo*age
gen wc_age = wc*age
gen pa_age = pa*age
* Generate single variable
gen single = 0
replace single = 1 if d3 == 1
gen sa_single = sa*single
gen fc_single = fc*single
```

```
gen mo_single = mo*single
gen wc_single = wc*single
gen pa_single = pa*single
* Generate tail interactions with sa, fc and wc
gen sa_tail = sa*tail
gen fc tail = fc*tail
gen wc_tail = wc*tail
* Four observations have missing values. We drop those:
drop if choice == 9
* Generate global random variables:
global randvars "sa fc wc pa mo"
* Define mixlogit model:
mixlogit y p1000 femaleSA femaleWC femaleFC femalePA femaleMO childrenSA childrenWC
childrenFC childrenMO childrenPA sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3 sa_nonInc
fc_nonInc wc_nonInc mo_nonInc pa_no\overline{nInc sa_āge fc_age wc_agè mo_age pa_age}
sa_single fc_single wc_single mo_single pa_single sa_tail fc_tail wc_tail,
ra\overline{n}d($randvars) group(\overline{i}dset) id(\overline{i}d) nrep(5\overline{0}0) corr
estimates store fish_inc3000_female_children_age_single
test femaleSA female\overline{WC}
test childrenSA childrenWC childrenFC childrenMO childrenPA
test sa_inc3 fc_inc3 wc_inc3 mo_inc3 pa_inc3
test sa_nonInc fc_nonInc wc_nonInc mo_nonInc pa_nonInc
test sa_age fc_age wc_age mo_age pa_age
test sa_single fc_sing}le wc_\overline{single m
***************************** fish_inc3000_female_children_age_single_educDummy SLUTT
***************************
```


## A.13.10. STATA Do-File to Convert the Dataset to a Cross Section Dataset

Table A.13.10.1. Do-File to Convert the Dataset to a Cross Section Dataset

```
clear
use H:\Masteroppgave\RCData\RCdataCross\RCDataCross
*drop block set alt choice y sa fc wc pa mo tail p300 p1000
sort id
quietly by id: gen dup = cond(_N==1,0,_n)
drop if dup >1
* Generate dummy variable for children. 1 = family have children under 18, 0
otherwise:
gen children = 0
replace children = 1 if q18_2 > 0
* Generate single variable
gen single = 0
replace single = 1 if d3 == 1
* Generate inc3000 variable
gen inc_3 = 0
replace inc_3 = 1 if q20 == 3
* Generate educUni *****
gen educUNI = 0
replace educUNI = 1 if d7 > 4
* young
gen young = 0
replace young = 1 if age < 36
* middle
gen middle = 0
replace middle = 1 if age > 35 & age < 56
gen midAge = age if middle == 1
* old
gen old = 0
replace old = 1 if age > 55
```


## A.13.11. STATA Do-File Unpaired Unequal Welch Tests

Table A.13.11.1. STATA Do-File Unpaired Unequal Welch Tests

```
clear
use H:\Masteroppgave\RCData\RCdataCross\RCDataCross
*drop block set alt choice y sa fc wc pa mo tail p300 p1000
sort id
quietly by id: gen dup = cond(_N==1,0,_n)
drop if dup >1
* Generelle sammenligninger av svar fra spørreundersøkelsen
gen q17_1xx = q17_1 if q17_1!=11
gen q17_2xx = q17__2 if q17__ 2!=11
ttest q17_1xx=q17_2xx, unpaired unequal welch
gen q6 1x\overline{x}= q6 1- if q6 1!=11
gen q6_2xx = q6_2 if q6_2!=11
ttest q}\mp@subsup{6}{_}{\prime}1xx=q6_2xx, unp\overline{p}aired unequal welc
gen q6_6xx = q6_6 if q6_6!=11
ttest q6_2xx=q6_6xx, unpaired unequal welch
gen q6_7xx = q6_7 if q6_7!=11
ttest q6 2xx=q6 7xx, unpaired unequal welch
gen q6_8xx = q6_8 if q6_8!=11
ttest q6 2xx=q6 8xx, unpaired unequal welch
ttest q1\overline{7}_1xx=q\overline{1}7_2xx, unpaired unequal welch
gen q17 3xx = q17 3 if q17 3!=11
ttest q\overline{1}7_1xx=q17_3xx, unp\overline{a}ired unequal welch
gen q17_4xx = q17_4 if q17_4!=11
ttest q\overline{1}7 3xx=q17 }\mp@subsup{}{}{-}4xx, unp\overline{a}ired unequal welc
gen q17_5xx = q17_5 if q17_5!=11
gen q17_7xx = q17_7 if q17_7!=11
ttest q17_5xx=q17_7xx, unpaired unequal welch
gen q17 8xx = q17 8 if q17 8!=11
ttest q17_8xx=q17_7xx, unpāired unequal welch
ttest q17 7xx=q17_8xx, unpaired unequal welch
ttest q17-}\mp@subsup{}{}{-}7=q17_8-if q17_7!=11 & q17 8!=11
ttest q17 7xx==q17 8xx, unpaired unequal welch
ttest q17 -}7xx==q17-8xx, unpaired unequal welc
ttest q17_7xx==q17_8xx, unpaired unequal
gen q17 6xx = q17 6 if q17 6!=11
ttest q\overline{17_6xx==q1\overline{7}_8xx, unp}\mp@subsup{\}{1}{\prime}red unequal welch
gen q17 9xx = q17 9 if q17 9!=11
gen q17_10xx = q17_10 if q17_10!=11
ttest q\overline{1}7 9xx==q17 -10xx, unpāired unequal welch
```


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## Appendix II

The survey questions answered by the participants are presented in this appendix.

## FISH IN FRANCE MAY2008

## Information from the database (D)

## D1. What is your gender?

Male


Female

## D2. How old are you?

Date of birth:

```
D3. Family status?
Single (with or without children)
Married or cohabiting (with or without children)
```



```
Living in collective home (residence for students or elderly) \(\square_{3}\)
```

D7. Education level
No diploma

Brevet des collèges

CAP ou BEP

Baccaulauréat (BAC)

BAC +2 or 3....................................................... $\square_{5}$
$\mathrm{BAC}+4$ or 5
$\square 6$
BAC +6
$\square_{7}$

## Questions to be asked at the recruitment stage (R)



R2. Who in your household is deciding what food to shop?
Check one or more boxes.
(R2.1) Yourself.
(R2.1) Your partner
$\square$
(P21) Someone els

| R3. How often would you say you buy the following fresh grocery products YOURSELF? Check one box per line. |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Check one box per li | line. <br> Twice a week or more | Once a week | 2-3 times a month | Once a month | Every second month | 2-4 <br> times <br> a year | More seldom | Never |
| (R3.1) | Poultry.. |  |  |  |  |  |  |  | $\square_{8}$ |
|  | Beef.................... |  |  |  | 4 | 5 |  |  |  |
|  | Pork .................... |  |  |  |  | 5 |  |  |  |
| (R3.4) | Lamb .................. |  |  |  |  |  |  |  |  |
| (R3.5) | Fresh Fish ............ |  | $\square_{2}$ | 3 | $\square_{4}$ | 5 |  | $\square$ | $\square_{8}$ |


|  | How often would Check one box per | you say y line. <br> Twice a week or more | Once a week | e followin <br> 2-3 times a month | items <br> Once <br> a month | home? <br> Every second month | 2-4 <br> times <br> a year | More seldom | Never |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (R4.1 <br> (R4.2 <br> (R4.3 <br> (R4.4) | Fish in a readymade meal Canned fish Frozen fish. Fresh fish. |  | $\square_{2}$ $\square_{2}$ $\square_{2}$ $\square_{2}$ |  |  | $\square_{5}$ $\square_{5}$ $\square_{5}$ $\square_{5}$ |  | $\square_{7}$ $\square_{7}$ $\square_{7}$ $\square_{7}$ |  |

If they eat fish less than once a month and buy fresh fish less than every second month, then STOP.

## FISH IN FRANCE

Thank you for taking part in this study of French consumers fish habits and preferences. The focus of this study is home consumption of fresh fish, and if noting else is specified we are asking about fresh fish for home consumption.

## Fish likings

1. What is your favourite fish?
2. On a scale from 1 to 10 , where 1 means dislike very much and 10 means just as good as my favourite fish, how would you rate the following fish types? If you have never tasted a fish or do not remember how it tasted, please check Do not know.
Check one box per line.


## Buying fish

3. On a scale from 1 to 10 , where 1 means you very strongly DISAGREE and 10 means you very strongly AGREE, how much do you agree with the following statements?
Check one box per line.


| In what kind of store do you normally purchase the salmon, cod, monk and pangasius for consumption in your household? <br> Check one box per column. |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Salmon | Cod | Monk | Pangasius |
| Fish shop .................................. | $\square_{1}$ | $\square \square_{1}$ | $\square_{1}$ | $\square_{1}$ |
| Super- or Hypermarket................. |  | 12 | 2 | $\square_{2}$ |
| Traditional wet market ................. | $\square 3$ | $\square_{3}$ | $\square$ | $\square 3$ |
| Other.. | $\square$ | $\square$ |  | $\square \square_{4}$ |
| Don't know / Can't remember ...... | 5 | $\square 5$ | $\square$ | $\square_{5}$ |
| Never bought.............................. | $\square_{6}$ | $\square 6$ | $\square_{6}$ | $\square$ |

5. In what form do you usually buy salmon, cod, monk and pangasius for home consumption? Check one box per column.

|  | Salmon | Cod | Monk | Pangasius |
| :---: | :---: | :---: | :---: | :---: |
| Fresh whole fish ......................... | $\square_{1}$ | $\square 1$ | $\square_{1}$ | $\square_{1}$ |
| Fresh filets................................ | $\square_{2}$ | $\square \square_{2}$ | $\square_{2}$ | $\square{ }^{2}$ |
| Frozen filets............................... | $\square_{3}$ | $\square 3$ | $\square_{3}$ | $\square \square_{3}$ |
| Other........................................ | $\square \square_{4}$ | $\square 4$ | $\square \square_{4}$ | $\square \square_{4}$ |
| Don't know/ Can't remember ........ | $\square_{5}$ | $\square$ | $\square 5$ | $\square 5$ |
| Never bought............................. | $\square 6$ | $\square 6$ | $\square \square_{6}$ | $\square 6$ |

## Origin of the fish

6. On a scale from 1 to 10 , where 1 means you very strongly DISAGREE and 10 means you very strongly AGREE, how much do you agree with following statements about the origin of fish?
Check one box per line.

| Very <br> strongly <br> disagree |  |  |  |  |  | Very <br> strongly <br> agree |  |  |  | Do <br> not <br> know |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 |

I have a very positive view of fresh farmed fish from:


## Last time you ate various types of fish

| Do you remember when the last time you ate fish (whatever the fish and the place)? Check one box |  |
| :---: | :---: |
| Less than 1 week ago .................. | $\square_{1}$ |
| 1-2 weeks ago............................. |  |
| 2-4 weeks ago............................ | $\square$ |
| 5-12 weeks ago.......................... | $\square_{4}$ |
| More than 3 months ago............... | $\square \square_{5}$ |
| Can't remember ......................... | $\square 6$ |


|  | Salmon | Cod | Monk | Pangasius |
| :---: | :---: | :---: | :---: | :---: |
| Less than 2 weeks ago.................. | 1 | $\square_{1}$ | $\square_{1}$ | $\square 1$ |
| 2-4 weeks ago............................. | 2 | $\square_{2}$ | $\square_{2}$ | $\square_{2}$ |
| 5-12 weeks ago........................... | $\square$ | $\square \square_{3}$ | $\square_{3}$ | $\square_{3}$ |
| More than 3 months ago................ | $\square_{4}$ | $\square_{4}$ | $\square_{4}$ | $\square 4$ |
| Can't remember $\qquad$ <br> Never tasted $\qquad$ | $\begin{aligned} & \square_{5} \\ & \square_{6} \end{aligned}$ | $\square_{5}^{5}$ | $\square_{5}$ | $\square_{5}^{5}$ |

9. Where did you last time eat salmon, cod, monk and pangasius?

Check one box per column.

|  | Salmon | Cod | Monk | Pangasius |
| :---: | :---: | :---: | :---: | :---: |
| At home. | $\square_{1}$ | $\square \square_{1}$ | $\square \square_{1}$ | $\square \square_{1}$ |
| At friends or family..................... | $\square_{2}$ | $\square_{2}$ | $\square_{2}$ | $\square_{2}$ |
| At a brasserie or restaurant............ | $\square_{3}$ | $\square_{3}$ | $\square_{3}$ | $\square 3$ |
| At a cafeteria or staff canteen......... | $\square_{4}$ | $\square_{4}$ | $\square_{4}$ | $\square_{4}$ |
| Other.................................. | $\square_{5}$ | $\square 5$ | $\square_{5}$ | $\square_{5}$ |
| Can't remember ........................... | $\square_{6}$ | $\square \square_{6}$ | $\square_{6}$ | $\square_{6}$ |
| Never tasted ............................... | $\square_{7}$ | $\square_{7}$ | $\square_{7}$ | $\square_{7}$ |

10. How often would you say you have salmon, cod, monk and pangasius at home? Check one box per line.

Twice Once 2-3 times Once Every 2-4 More Never a week a week a month a month second times seldom or more month a year

| Salmon ................ | $\square_{1}$ | $\square_{2}$ | $\square \square_{3}$ | $\square \square_{4}$ | $\square_{5}$ | $\square_{6}$ | $\square_{7}$ | $\square$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Cod. |  | $\square_{2}$ | $\square \square_{3}$ | $\square_{4}$ | $\square_{5}$ | $\square \square_{6}$ | $\square$ | $\square$ |
| Monk .............. |  | $\square_{2}$ | $\square 3$ | $\square_{4}$ | $\square 5$ | $\square 6$ | $\square_{7}$ | $\square$ |
| Pangasius............. |  | $\square_{2}$ | $\square 3$ | $\square_{4}$ | $\square 5$ | $\square_{6}$ | 7 | $\square$ |

11. Which of the following statements best describe your fish consumption at home? Check one box.
I eat fish at home mainly on Monday to Thursday
I eat fish at home mainly Fridays, Saturdays and/or Sundays
I eat fish at home regularly all days of the week (Monday to Sunday)
I rarely eat fish at home


## Attitudes towards fresh salmon



## Attitudes towards fresh cod



## Attitudes towards fresh monk



## Attitudes towards Pangasius



## Fish prices

16. What is your best guess at the average market price for one kilogram of fresh salmon, cod, monk and pangasius fillets this week?
Check one box per line.
Price per kilogram fillet

|  |  | Price per kilogram fillet |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $€ 3$ | $€ 6$ | €9 | $€ 12$ | $€ 15$ | €18 | $€ 21$ | €24 | €27 | €30 | €33 | $€ 36$ |
|  | Salmon ....... |  |  |  |  |  |  |  |  |  |  |  | $\square$ |
| (16.2) | Cod............ |  |  | $\square$ | $\square$ | $\square$ | $\square$ | $\square$ | $\square$ | $\square$ | $\square$ | $\square$ | $\square$ |
|  | Monk......... Pangasius.... |  |  |  |  |  |  |  |  | $\square$ |  |  | $\square$ |

## Attitudes toward fish farming and environmental aspects



## Demographics

18. How many persons live in your household, - included yourself?
(18.1) Number of adults (>18):
(18.2) Number of children ( $\leq 18$ ):
19. What is your current occupational situation?

Check one box.

|  | Paid work full time |
| :---: | :---: |
|  | Paid work part time |
|  | Unemployed for less than 3 month |
|  | Unemployed for more than 3 month |
|  | Housewife |
|  | Student |
|  | Retired or not able to work through illness.. |
|  | Civil servant |
|  | Working pensioner. |
|  | Other . |

20. What is your household's gross monthly income? Check one box.

Less than 2000 euros per month. $\qquad$


From 2000 to 3000 euros per month
Over 3000 euros per month


Do not now / Do not want to answer


## Appendix III

The attached file is a part of a larger table that was sent to us personally us from Johan Kvalheim. Johan Kvalheim is currently a Representative of the Norwegian Seafood Council in France and the UK ${ }^{1}$. The file was sent to us on e-mail the $20^{\text {th }}$ of April 2012.

The table consists of unrevised and preliminary data for Norwegian seafood exports for 2010 and 2011. The sources for the export statistics are from Statistics Norway (SSB) and Norwegian Customs (TAD). An explanation of the table is given in Norwegian. Since the data are preliminary and unrevised, they must be interpreted with caution.

[^8]
## EKSPORTSTATISTIKK

Vedlagt følger statistikk over sjømateksporten for siste måned. Statistikken er basert på urettede og foreløpige tall. Den inneholder tabeller med detaljert eksportstatistikk av hovedprodukter. I tillegg til enkeltprodukter er også vedlagt samlestatistikk for noen produktgrupper.

Statistikken inneholder også importstatistikk som viser norsk import for de største produktgrupper og de viktigste land.

Importstatistikken inneholder tall som er to måneder gammel.
Datakilde for eksportstatistikken er Statistisk sentralbyrå (SSB) og Toll- og avgiftsdirektoratet (TAD) (siste måned). Statistikkgrunnlaget er data fra TAD med informasjon om eksporten av fiskeprodukter. Datasettene inneholder eksportmengde og verdi spesifisert på produkttype og eksportland. Denne statistikken inneholder tall for siste måned der datakilden er Toll- og avgiftsdirektoratet (urevidert) og hittil i år der kilden er SSB. Statistikken bør derfor brukes med varsomhet.

Norges sjømatråd samarbeider med SSB for å få statistikken så korrekt som mulig, men er i tillegg avhengig av eksportørene for å få en pålitelig statistikk. Vi ber brukere som finner feil eller urimeligheter i statistikken straks å ta dette opp med Norges sjømatråd, som vil gi beskjed videre til SSB.

Tabellene viser eksporten fordelt på samtlige markeder.

- for siste måned ( mengde, verdi og pris, urettede tall). -totalt hittil dette år ( mengde, verdi og pris), foreløpige offisielle tall + urettede tall for siste måned - mengde, verdi og kilopris for januar - siste måned i fjor, endelige tall.

Eksportverdiene er oppgitt FOB, mens importverdiene er oppgitt CIF.
Statistikken er laget bare til bruk for aktører innenfor næringen.

## BRUK AV EKSPORTSTATISTIKKEN

Totaltabellene for fersk og frossen fisk inneholder ikke sild eller filet.
EU er ikke tatt med i totalen. Denne er kun ment som ekstra informasjon.
I den enkelte tabell er land ikke ført opp hvis eksportverdien er mindre enn kr 10.000,-. Hvis vekt er oppgitt til 0 betyr dette at vekt er under 1000 kg . I total mengde og verdi er all eksport inkludert.

Oversikt over eksporten hittil i år inneholder varegrupper som er utelatt fra den detaljerte oversiktstatistikken. Disse grupper er ofte lite spesifiserte varegrupper.

Grupper av annen (eks. annen filet) og ikke spesifisert (eks. fersk fisk - ikke spes.) inneholder varegrupper (restgrupper ol.) som av forskjellige grunner ikke er klassifisert som egne grupper.

Varenummeret refererer til HS-nummeret som benyttes i internasjonal toll- og statistikknomenklatur. Dette nummeret finnes bl.a. i TADs Statistisk varefortegnelse for utenrikshandelen.

Vennlig hilsen Norges sjømatråd:

Mengde i tonn, verdi i 1000 NOK

|  | Desember 2011 Ureviderte tall |  |  | Januar - Desember 2011 Ureviderte tall |  |  | Januar - Desember 2010 Foreløpige tall |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mengde | Verdi | Pris pr.kg | Mengde | Verdi | Pris pr.kg | Mengde | Verdi | Pris pr.kg |
| TOTALT | 192.551 | 4.681 .804 | 24,31 | 2.325 .928 | 52.974 .486 | 22,78 | 2.665 .047 | 53.618.134 | 20,12 |
| EU27 | 98.921 | 2.572 .285 | 26,00 | 1.174.496 | 30.467.389 | 25,94 | 1.268.730 | 30.833.303 | 24,30 |
| Russland | 29.267 | 574.246 | 19,62 | 300.088 | 5.193.061 | 17,31 | 347.538 | 5.222 .964 | 15,03 |
| Frankrike | 17.568 | 499.702 | 28,44 | 148.582 | 5.110.301 | 34,39 | 142.489 | 5.273 .089 | 37,01 |
| Danmark | 10.924 | 262.480 | 24,03 | 176.091 | 3.599.272 | 20,44 | 196.800 | 4.098.816 | 20,83 |
| Polen | 13.387 | 328.546 | 24,54 | 135.577 | 3.525.376 | 26,00 | 153.820 | 4.111 .669 | 26,73 |
| Japan | 9.084 | 243.812 | 26,84 | 140.057 | 2.877.206 | 20,54 | 135.673 | 2.473 .652 | 18,23 |
| Kina | 9.406 | 169.365 | 18,01 | 166.740 | 2.525.259 | 15,14 | 152.071 | 2.341 .106 | 15,39 |
| Storbritannia | 6.962 | 181.167 | 26,02 | 93.219 | 2.399.377 | 25,74 | 127.436 | 2.567.915 | 20,15 |
| Sverige | 6.087 | 204.263 | 33,56 | 71.571 | 2.375.811 | 33,20 | 69.692 | 2.265.240 | 32,50 |
| Tyskland | 10.462 | 236.477 | 22,60 | 104.854 | 2.282.603 | 21,77 | 99.544 | 2.115.907 | 21,26 |
| Portugal | 2.515 | 103.398 | 41,11 | 57.493 | 2.258.208 | 39,28 | 50.103 | 1.857 .578 | 37,08 |
| Spania | 4.985 | 149.066 | 29,91 | 59.572 | 1.916.873 | 32,18 | 49.934 | 1.765 .050 | 35,35 |
| Nederland | 6.426 | 139.409 | 21,69 | 97.258 | 1.881.566 | 19,35 | 94.245 | 1.819 .536 | 19,31 |
| U S A | 4.992 | 178.005 | 35,66 | 39.164 | 1.683 .341 | 42,98 | 56.764 | 2.726.193 | 48,03 |
| Brasil | 4.917 | 198.372 | 40,34 | 34.354 | 1.333.641 | 38,82 | 35.548 | 1.306 .343 | 36,75 |
| Italia | 3.284 | 124.193 | 37,82 | 31.226 | 1.330.784 | 42,62 | 29.340 | 1.431 .715 | 48,80 |
| Ukraina | 10.654 | 135.793 | 12,75 | 109.582 | 1.180.820 | 10,78 | 140.013 | 1.059.113 | 7,56 |
| Finland | 3.627 | 99.646 | 27,48 | 50.210 | 1.134.011 | 22,59 | 45.522 | 1.046.548 | 22,99 |
| Litauen | 5.794 | 96.066 | 16,58 | 72.298 | 1.031.928 | 14,27 | 71.416 | 976.931 | 13,68 |
| Hong Kong | 1.395 | 36.405 | 26,10 | 17.276 | 581.939 | 33,69 | 19.837 | 782.455 | 39,44 |
| Sør-Korea | 1.314 | 32.328 | 24,61 | 23.936 | 567.280 | 23,70 | 20.676 | 517.622 | 25,04 |
| Vietnam | 2.142 | 57.068 | 26,64 | 19.913 | 562.208 | 28,23 | 7.783 | 223.979 | 28,78 |
| Nigeria | 345 | 18.229 | 52,83 | 54.804 | 540.652 | 9,87 | 160.157 | 889.557 | 5,55 |
| Tyrkia | 3.316 | 39.094 | 11,79 | 42.873 | 539.442 | 12,58 | 53.531 | 504.028 | 9,42 |
| Israel | 1.464 | 40.182 | 27,44 | 14.805 | 483.608 | 32,66 | 11.799 | 388.303 | 32,91 |
| Latvia | 2.168 | 41.826 | 19,29 | 20.370 | 407.539 | 20,01 | 14.617 | 257.719 | 17,63 |
| Taiwan | 1.013 | 25.339 | 25,02 | 14.062 | 399.045 | 28,38 | 12.429 | 399.695 | 32,16 |
| Thailand | 1.440 | 42.946 | 29,81 | 14.795 | 386.628 | 26,13 | 13.021 | 417.388 | 32,05 |
| Hellas | 1.929 | 29.088 | 15,08 | 25.684 | 379.641 | 14,78 | 24.741 | 261.309 | 10,56 |
| Belgia | 935 | 30.377 | 32,48 | 10.371 | 365.699 | 35,26 | 11.841 | 417.227 | 35,24 |
| Hviterussland | 3.927 | 68.941 | 17,55 | 17.570 | 312.843 | 17,81 | 29.414 | 354.485 | 12,05 |
| UKJENT | 43 | . | . | 1.527 | 265.613 | 173,96 | 1.634 | 287.668 | 176,08 |
| Canada | 1.029 | 33.034 | 32,09 | 6.479 | 246.414 | 38,03 | 5.353 | 165.320 | 30,88 |
| Sveits | 486 | 26.936 | 55,43 | 4.423 | 241.158 | 54,52 | 4.453 | 232.620 | 52,24 |
| Den Dominikanske Republikk | 451 | 13.143 | 29,16 | 8.306 | 231.408 | 27,86 | 8.932 | 242.672 | 27,17 |
| Singapore | 588 | 18.611 | 31,66 | 6.457 | 222.982 | 34,54 | 5.495 | 220.399 | 40,11 |
| Egypt | 187 | 2.941 | 15,72 | 25.601 | 181.116 | 7,07 | 37.279 | 203.907 | 5,47 |
| Kasakhstan | 1.488 | 20.046 | 13,47 | 14.772 | 174.746 | 11,83 | 19.056 | 177.548 | 9,32 |
| Australia | 198 | 12.584 | 63,47 | 3.053 | 149.537 | 48,97 | 2.944 | 132.391 | 44,96 |
| Kongo | 585 | 16.827 | 28,76 | 5.466 | 145.525 | 26,63 | 4.819 | 119.588 | 24,82 |
| Tsjekkia | 682 | 18.872 | 27,66 | 4.959 | 135.232 | 27,27 | 2.126 | 71.931 | 33,84 |
| Kongo, Brazzaville | 266 | 7.539 | 28,30 | 5.218 | 132.742 | 25,44 | 3.603 | 88.143 | 24,46 |


[^0]:    ${ }^{1}$ For an explanation of the BDM, see Appendix I, section A.2.

[^1]:    ${ }^{2}$ See Appendix I, section A.6.1.2 for a detailed distribution of the education levels

[^2]:    ${ }^{3}$ Some interaction terms also includes the fish types. The parameter estimates for these interaction terms are not random.
    ${ }^{4}$ The log-normal distribution was not chosen, since we assumed that we may obtain "negative utility" from choosing some of the fish types.

[^3]:    ${ }^{5}$ In Appendix I, section A.6.6.1, there is an explanation of why Income was chosen as a dummy variable, and not as a continuous variable.

[^4]:    ${ }^{6}$ The prices for salmon ranged from 7 to 25 EUR/KG. The prices for cod ranged from 9 to 23 EUR/KG, but the prices for loin fillets similar to the ones used in the experiment ranged from 29 to 36 EUR/KG. The prices for monk ranged from 28 to 40 EUR/KG. Fresh fillets of pangasius were not found in grocery stores, but the price for frozen pangasius was $10 \mathrm{EUR} / \mathrm{KG}$. The prices referred to as the "prices found in grocery stores" are the average of the minimum and the maximum value.

[^5]:    Hypothesis 1
    Salmon * IncGr2
    $=$ Farmed Cod $*$ IncGr 2
    $=$ Wild Cod $*$ IncGr 2
    $=$ Monk $*$ incGr 2
    $=$ Pangasius $*$ IncGr 2
    $=0$

[^6]:    Hypothesis 1
    Salmon * Income
    $=$ Farmed Cod $*$ Income
    $=$ Wild Cod $*$ Income
    $=$ Monk $*$ income
    $=$ Pangasius $*$ Income
    $=0$

[^7]:    test sa_age fc_age wc_age mo_age pa_age
    *************************** fish inc3000 age SLUTT

[^8]:    ${ }^{1}$ http://www.seafood.no/Om-oss/Organisasjon/Ansatte (accessed: 08.05.2012)

